

When you see a title like that, you need to worry: “Uh-oh, sounds philosophical.”

Well, I just wanted to tell you two concrete stories about cases when my colleagues and I managed to do something useful by virtue of knowing something about inference. The ideas we needed were things I didn't know a few years ago, so I thought you might be interested too.

Inference in biological physics

Phil Nelson
University of Pennsylvania

For these slides see:
www.physics.upenn.edu/~pcn



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There's more, of course, but t

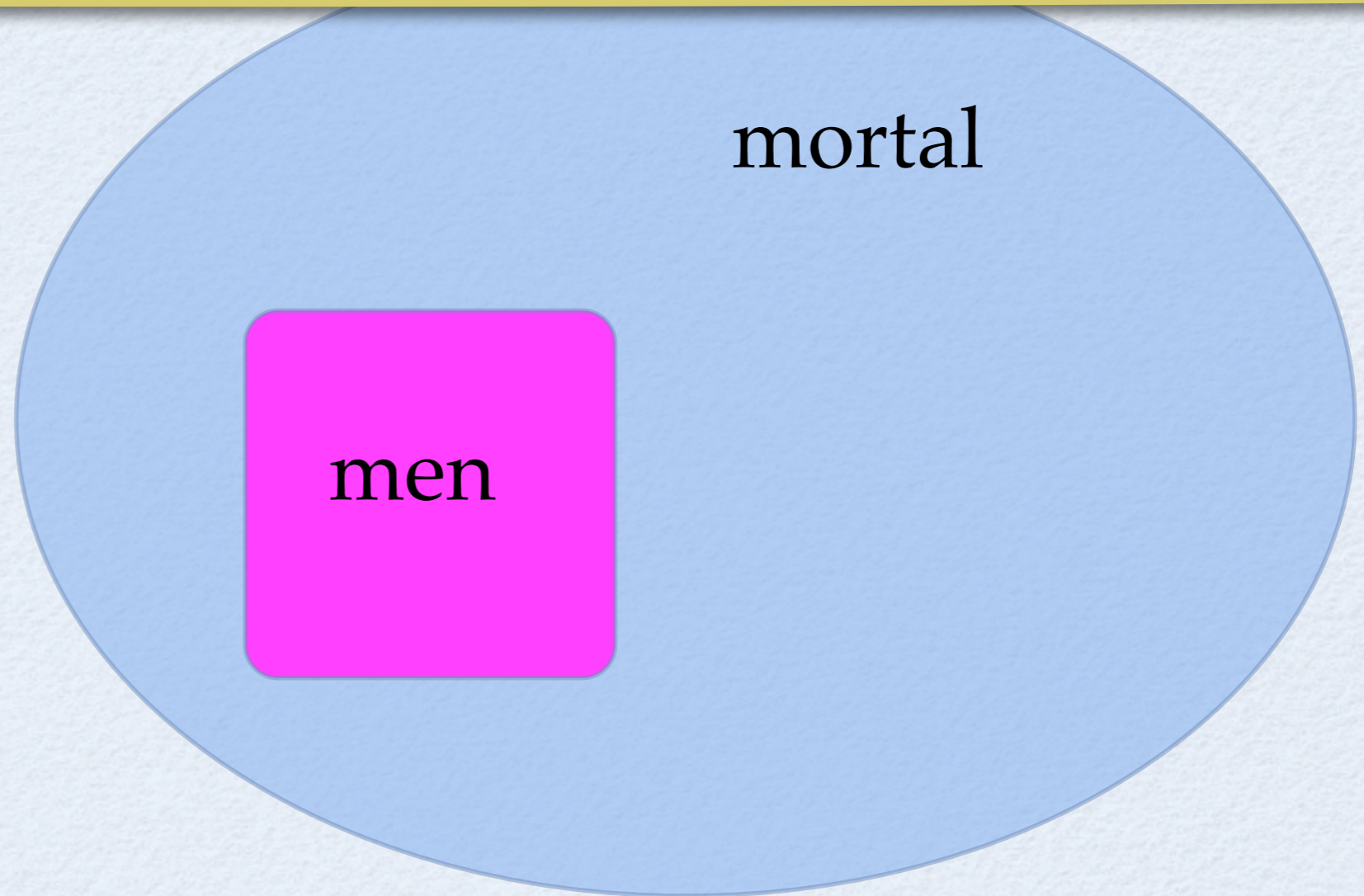
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Part I:

start with a topic that may not be obviously biophysical in character.
Suppose I stood here and said "all men are mortal; Socrates is mortal; therefore Socrates is a man"

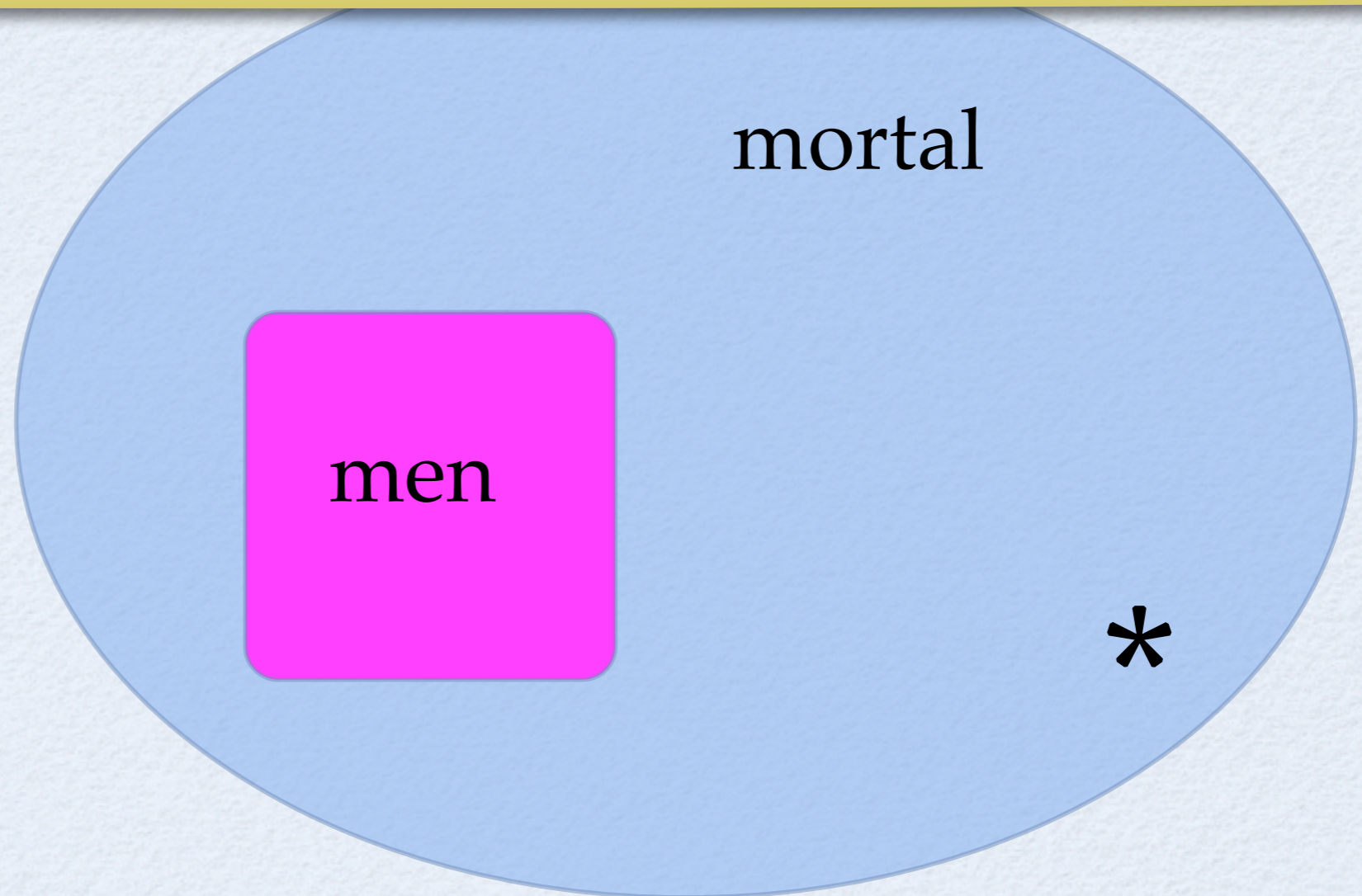
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In classical logic it's fairly easy to spot errors of inference.

An everyday question in clinical practice

To diagnose colorectal cancer, the hemoccult test—among others—is conducted to detect occult blood in the stool. This test is used from a particular age on, but also in routine screening for early detection of colorectal cancer. Imagine you conduct a screening using the hemoccult test in a certain region. For symptom-free people over 50 years old who participate in screening using the hemoccult test, the following information is available for this region:

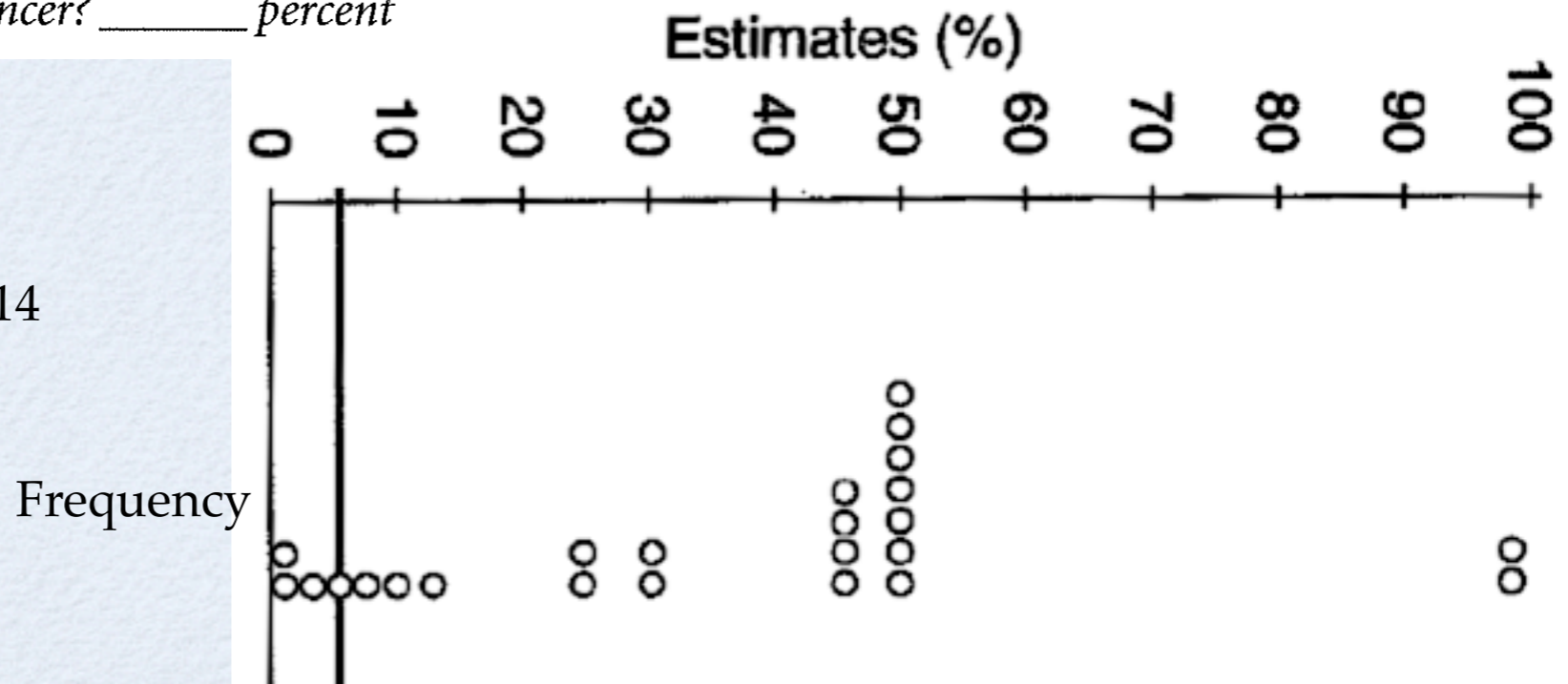
The probability that one of these people has colorectal cancer is 0.3 percent. If a person has colorectal cancer, the probability is 50 percent that he will have a positive hemoccult test. If a person does not have colorectal cancer, the probability is 3 percent that he will still have a positive hemoccult test. Imagine a person (over age 50, no symptoms) who has a positive hemoccult test in your screening. What is the probability that this person actually has colorectal cancer? _____ percent

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Here are the replies of 24 practicing physicians, who had an average of 14 years of professional experience:



G. Gigerenzer, *Calculated risks*

Work it out

We are asked for $P(\text{sick} | +) = B / (B+D)$.

A=Sick, -

C=Healthy, -

B=Sick, +

D=Healthy, +

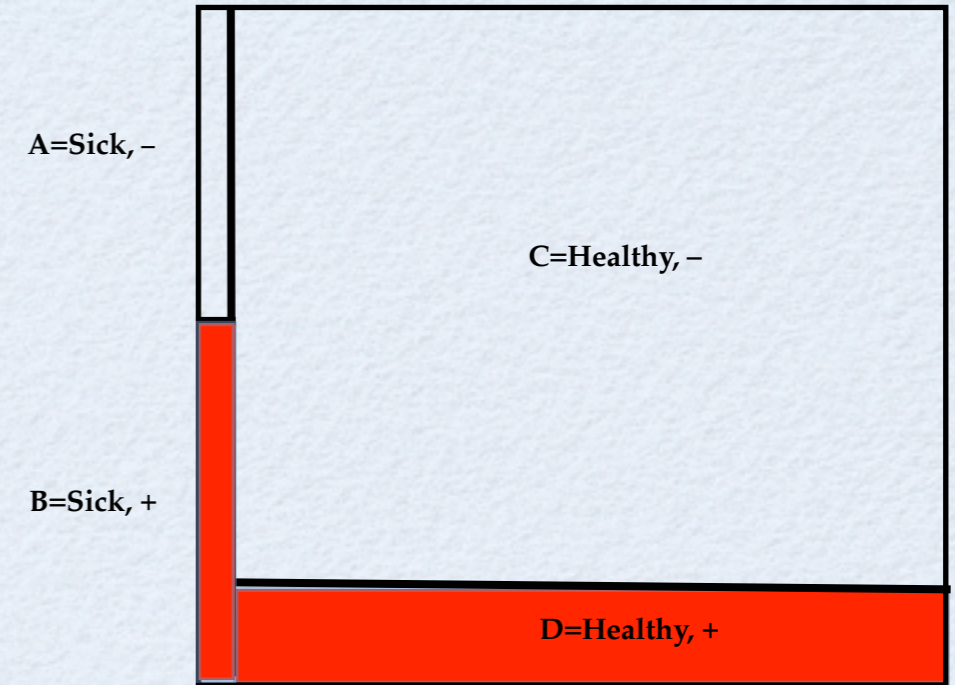


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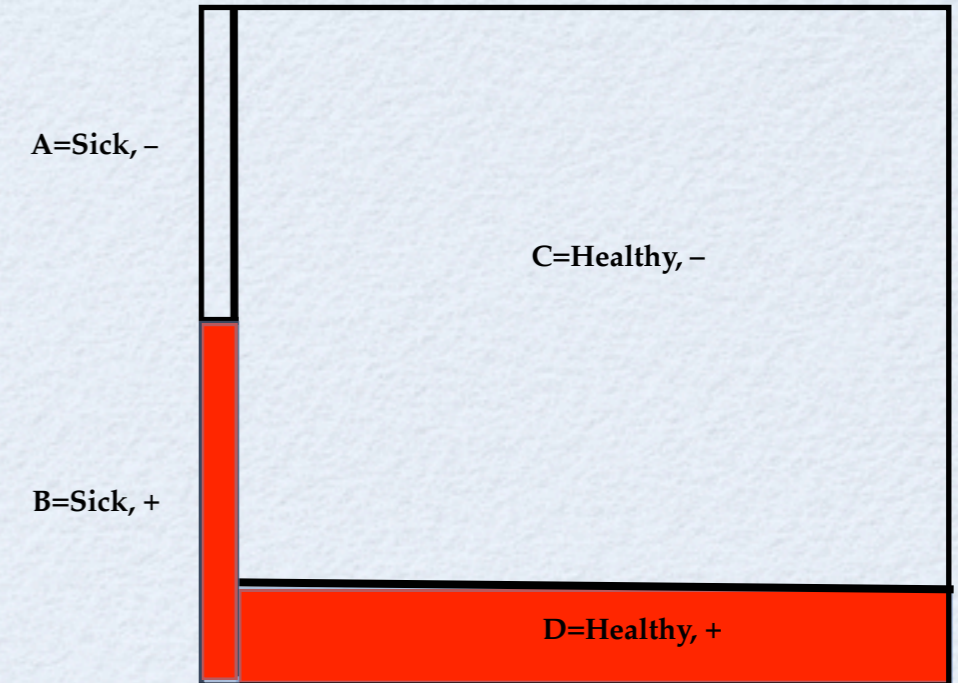
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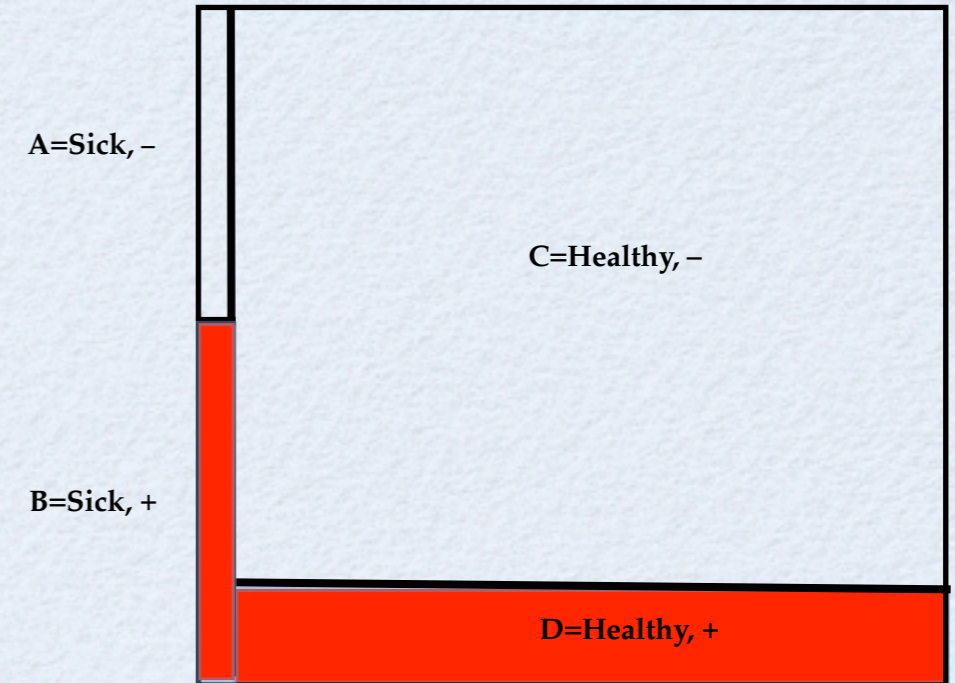
But what we were given was $P(+ | \text{sick}) = B / (A+B)$.



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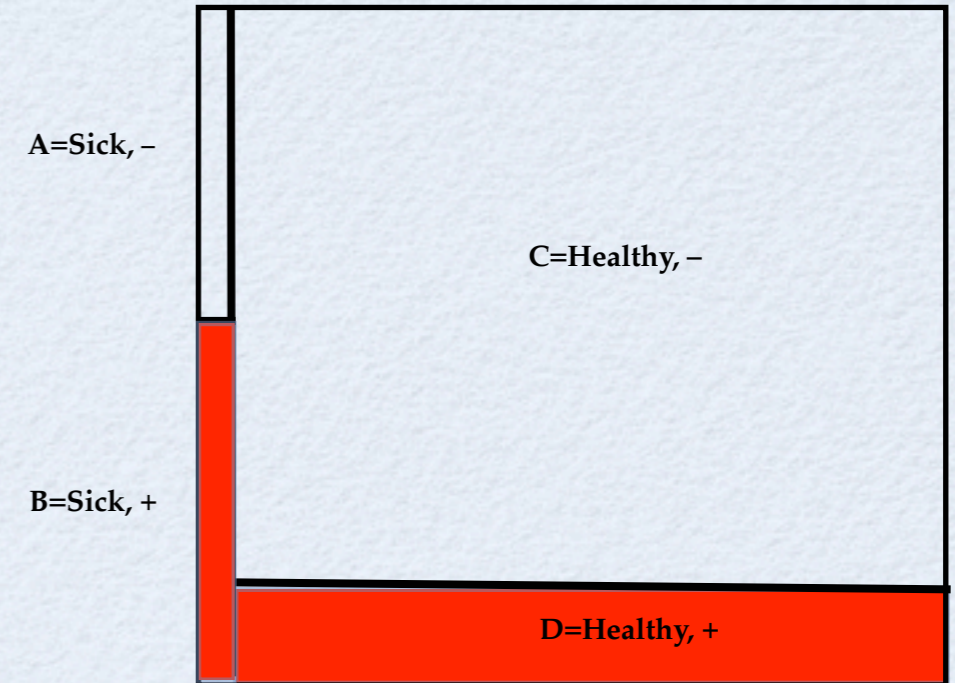
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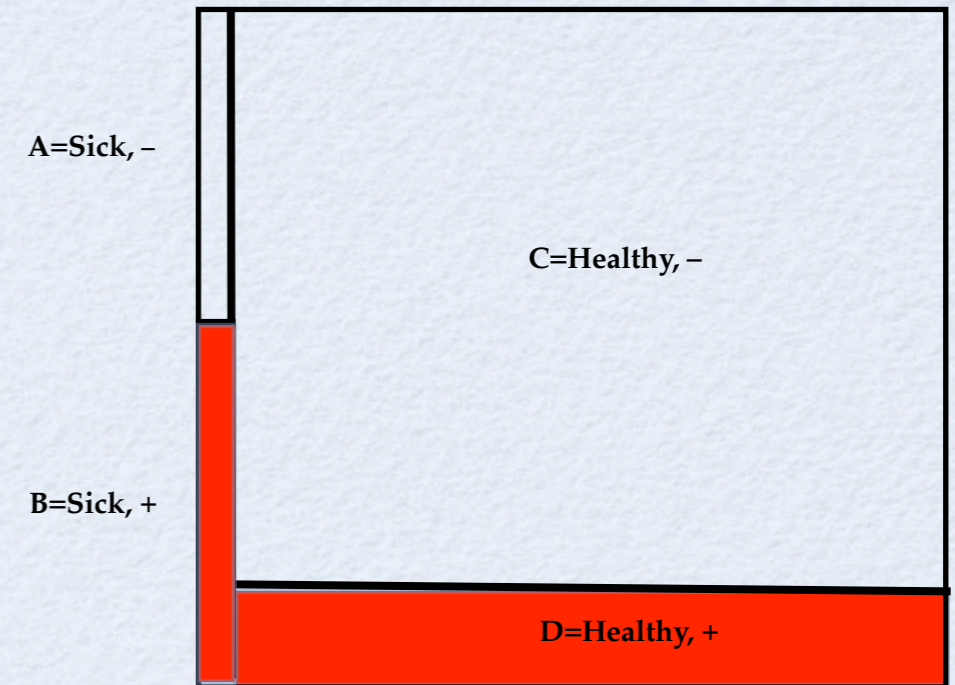
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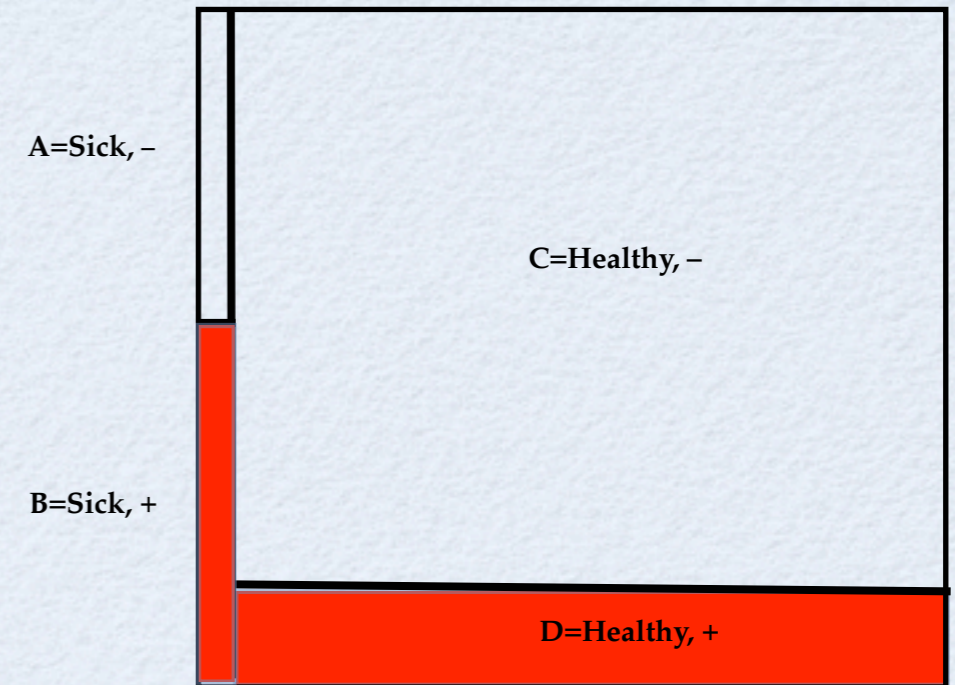


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These are not the same thing: they have different denominators. To get one from the other we need some more information:



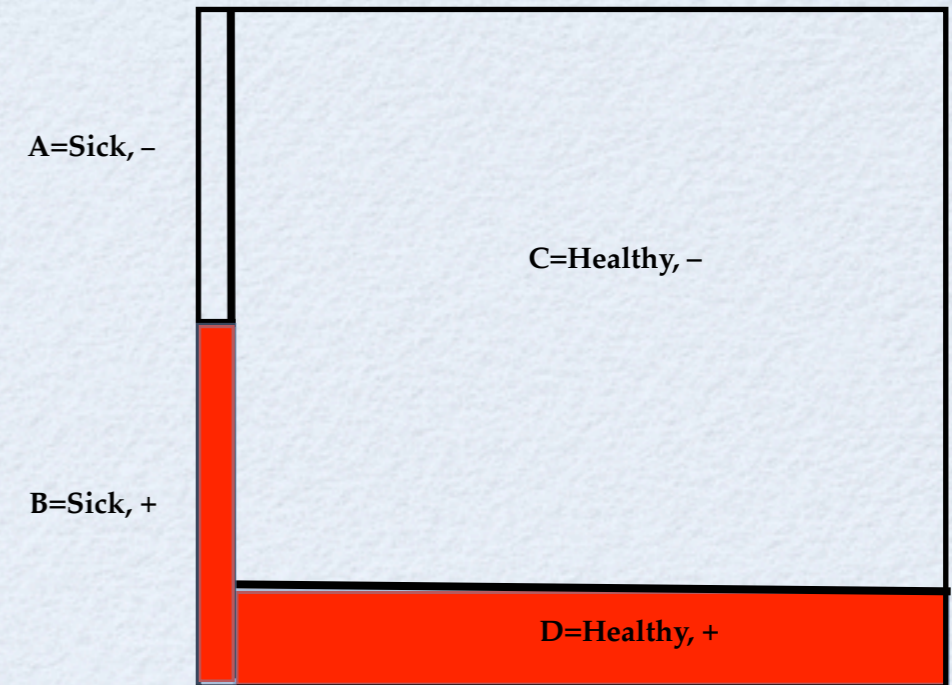
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$$\frac{B}{B+D} = \frac{B}{A+B} \times \frac{A+B}{B+D}$$



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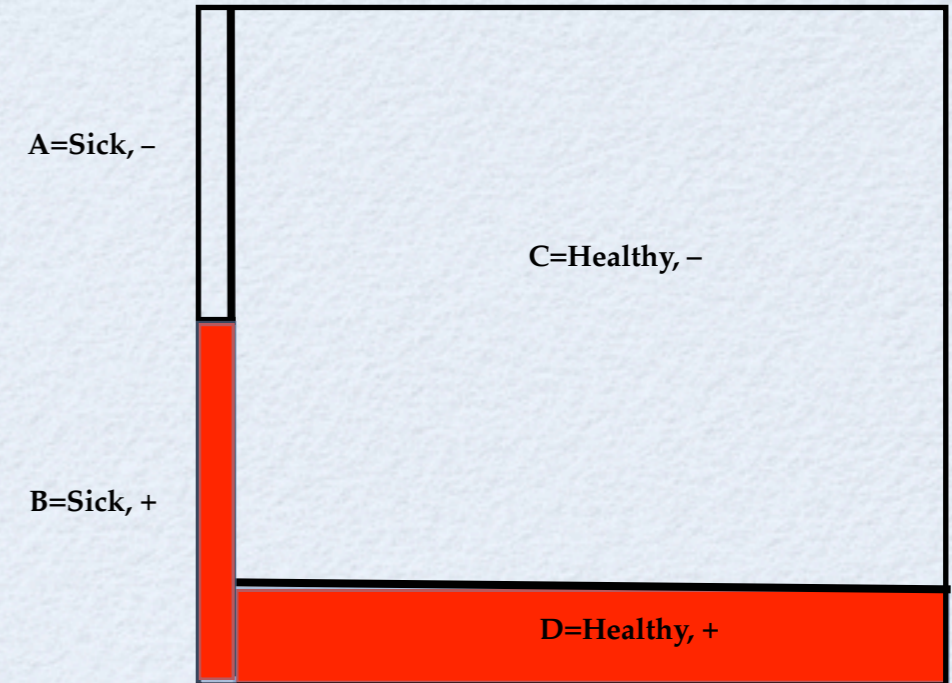
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Posterior
estimate
(desired)

Likelihood
(given)

Prior
estimate
(given)

Still need this



Finish working it out

$$P(\text{sick}|+) = P(+|\text{sick}) \times \frac{P(\text{sick})}{P(+)}$$

Is that last factor a big deal?
P(sick) was given, but we need:

A=Sick, -

B=Sick, +

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Finish working it out

$$P(\text{sick}|+) = P(+|\text{sick}) \times \frac{P(\text{sick})}{P(+)} \quad \begin{array}{l} \mathbf{A=Sick, -} \\ \mathbf{B=Sick, +} \end{array}$$

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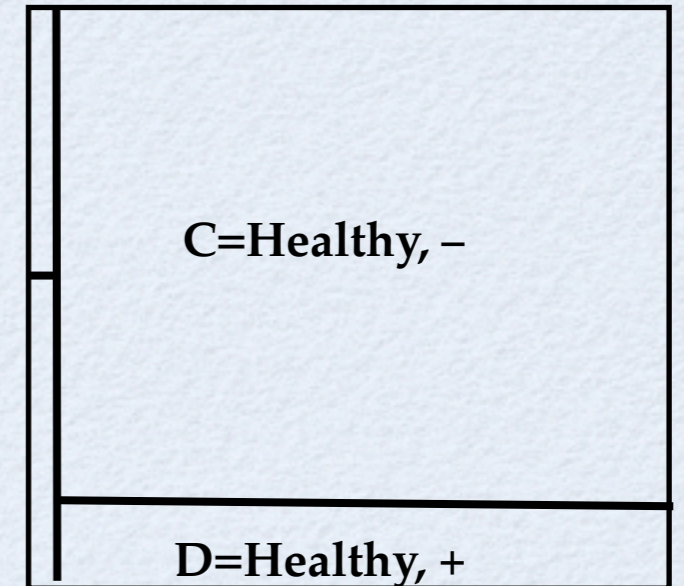
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$$P(+)=B+D$$

$$= \frac{B}{A+B}(A+B) + \frac{D}{C+D}(C+D)$$

$$= P(+|\text{sick})P(\text{sick}) + P(+|\text{healthy})P(\text{healthy})$$

$$= (0.5)(0.003) + (0.03)(0.997) \approx 0.03$$



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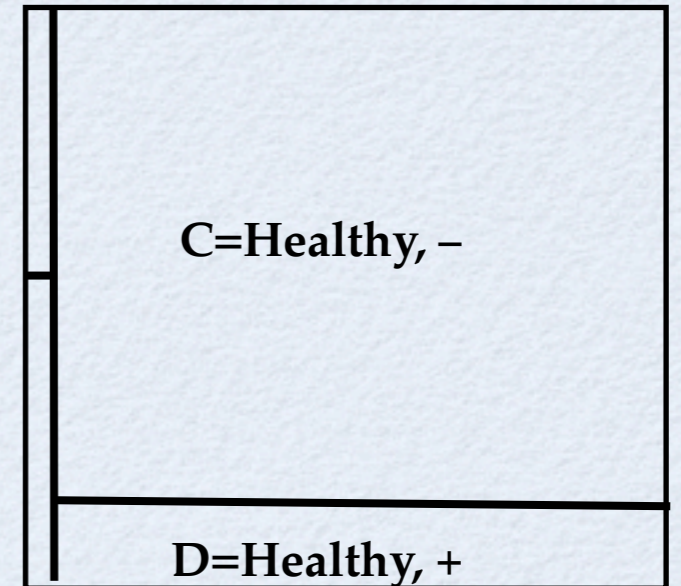
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$$\frac{P(\text{sick})}{P(+)} \approx \frac{0.003}{0.03} \approx 0.1$$

It's huge: a positive test result means only a 5% chance you're sick. Not 97%.



Part II: Change point analysis in single-molecule TIRF

JF Beausang, Yale Goldman, PN

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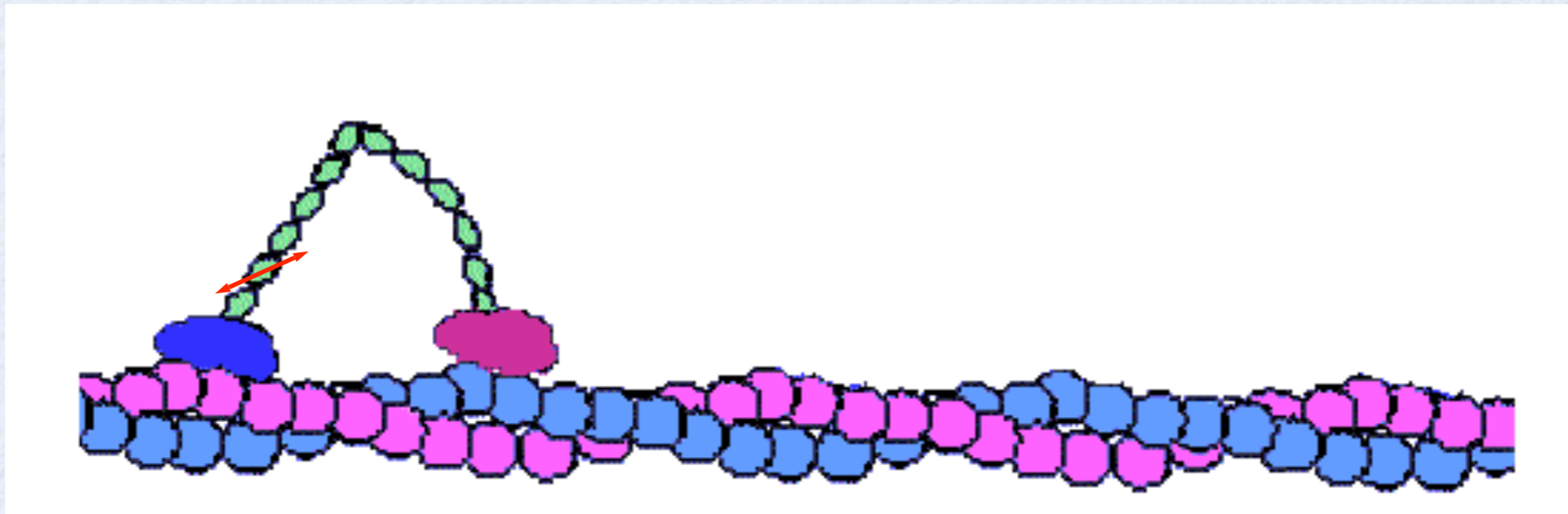


Many thanks to Haw Yang. See also Lucas P. Watkins and Haw Yang *J. Phys. Chem. B* **2005**

Myosin V Processivity

We'd like to know things like: How does it walk? What are the steps in the kinetic pathway? What is the geometry of each state?

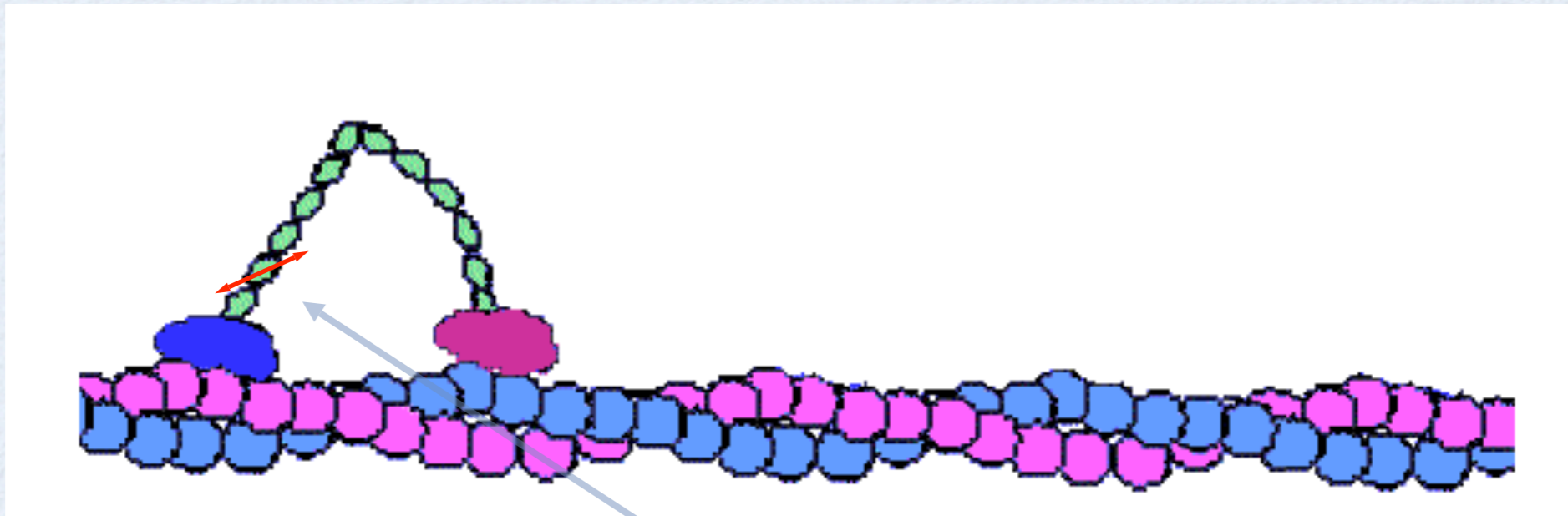
One classic approach is to monitor the position in space of a marker (e.g. a bead) attached to the motor. But this does not address the geometry of each state.



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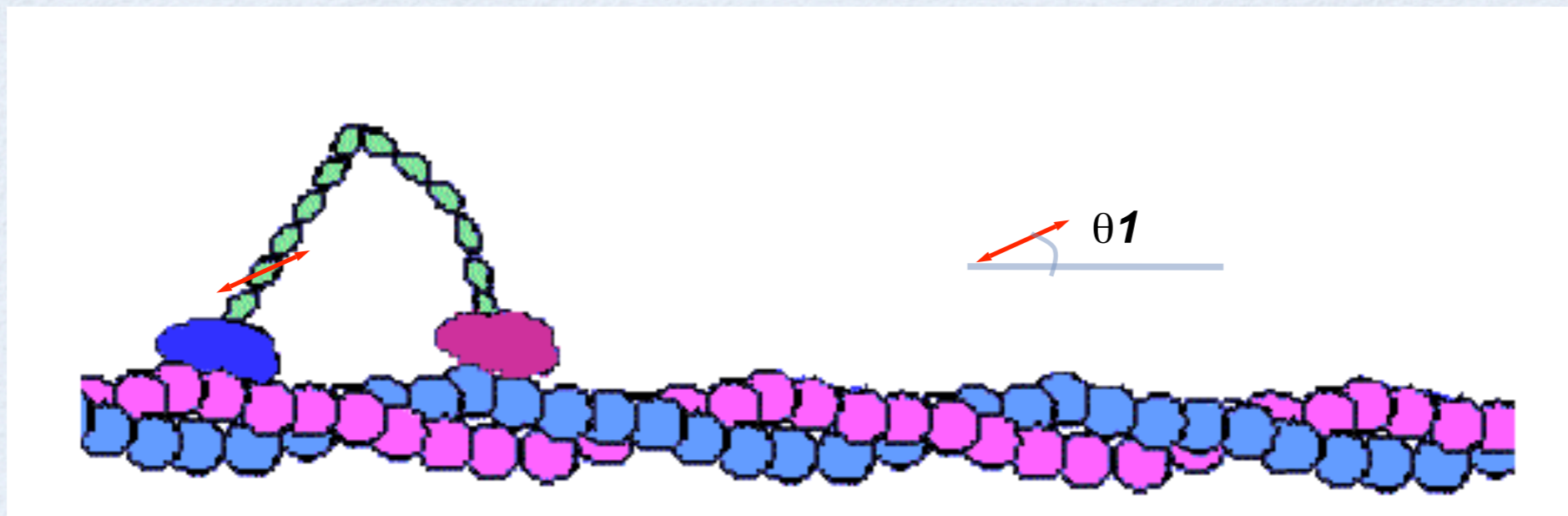
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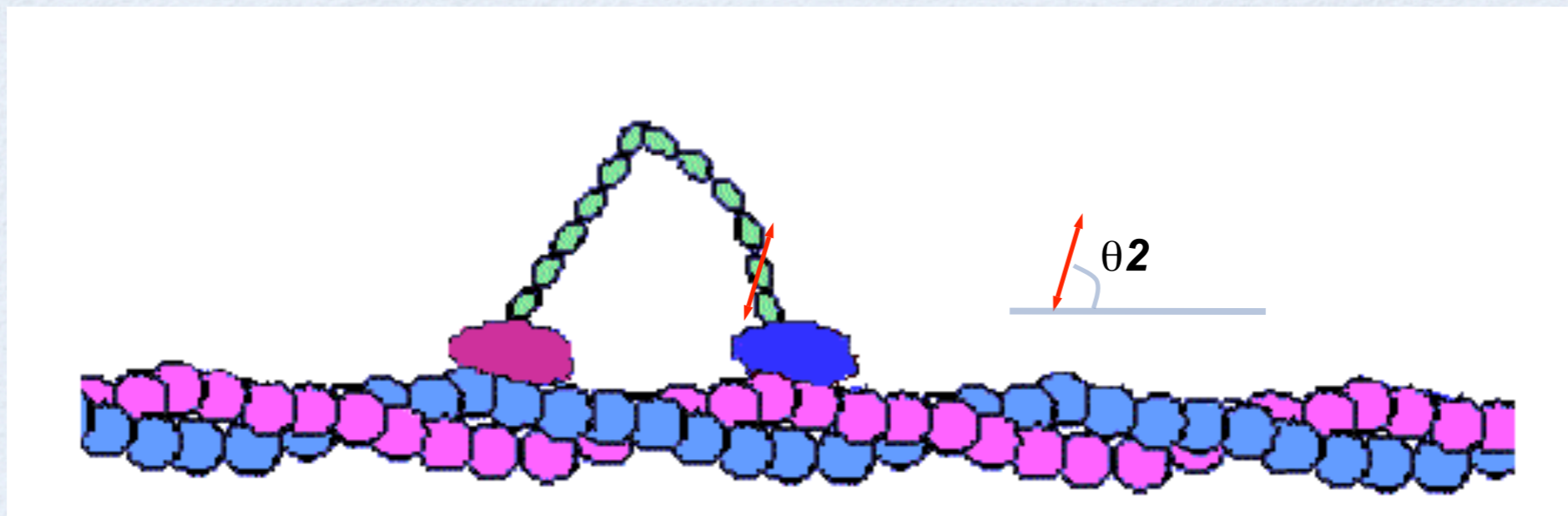


The approach I'll discuss involves attaching a bifunctional fluorescent label to one lever arm. The label has a dipole moment whose *orientation* in space reflects that of the arm.

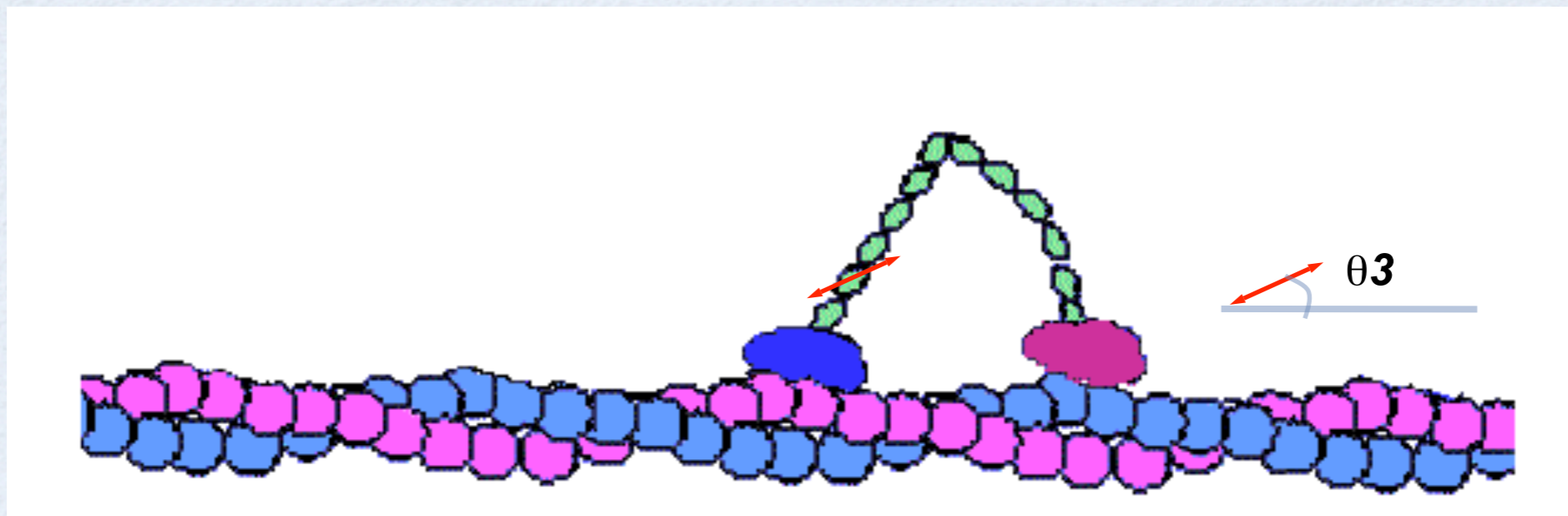
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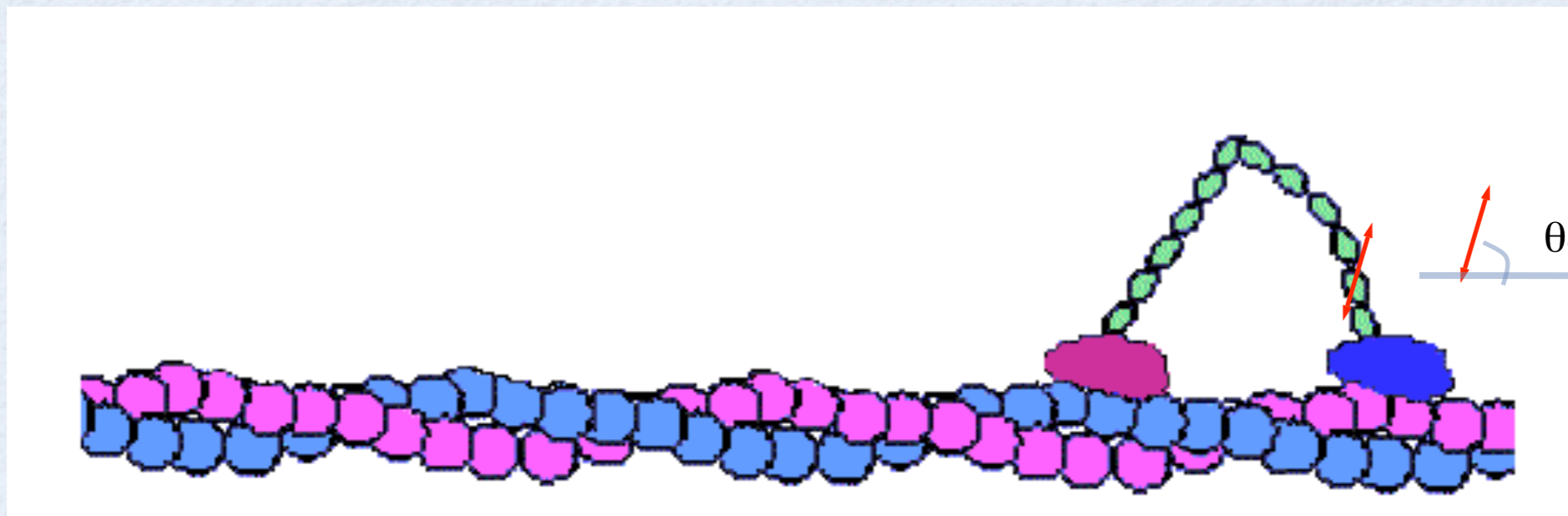
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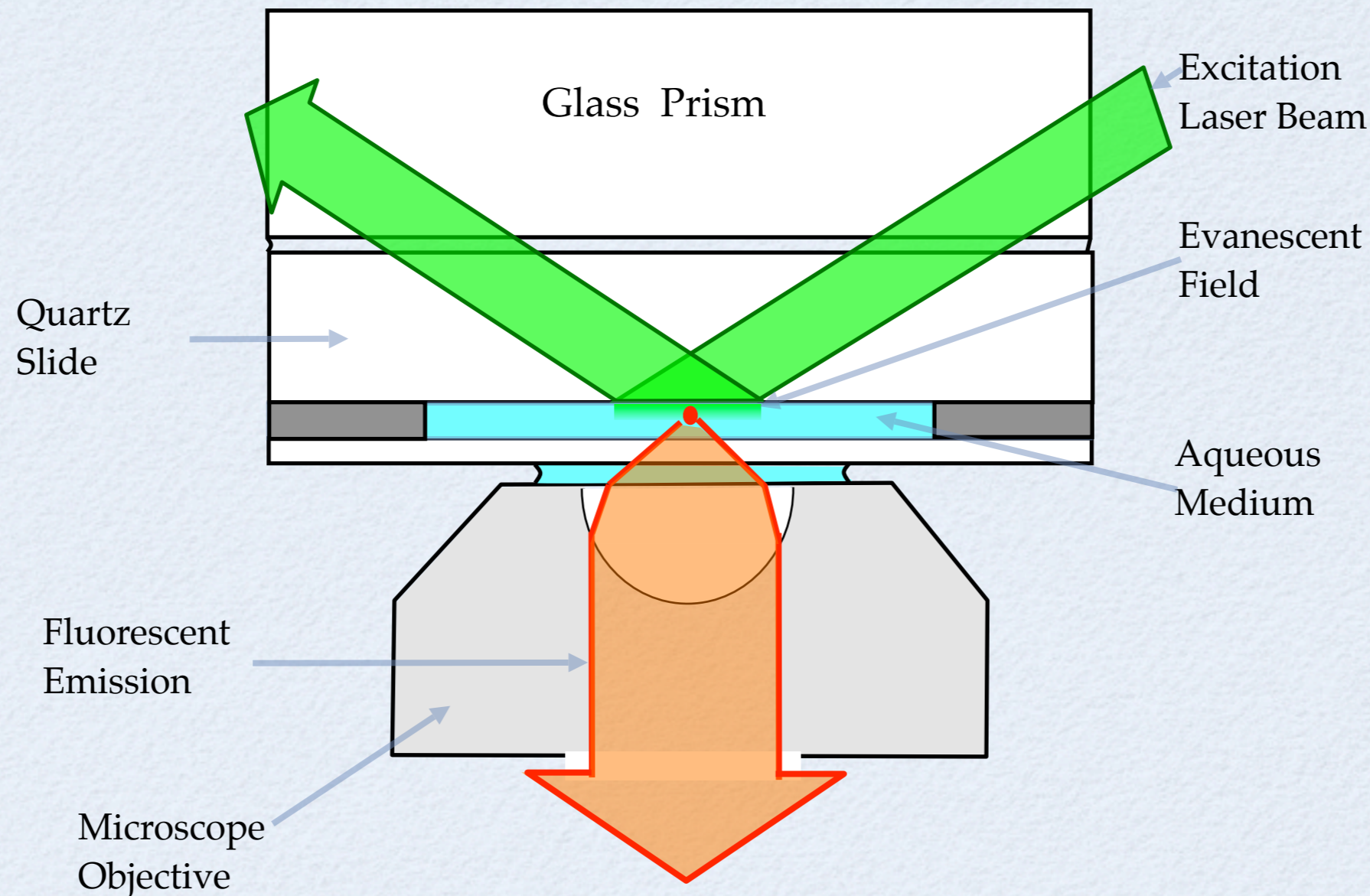
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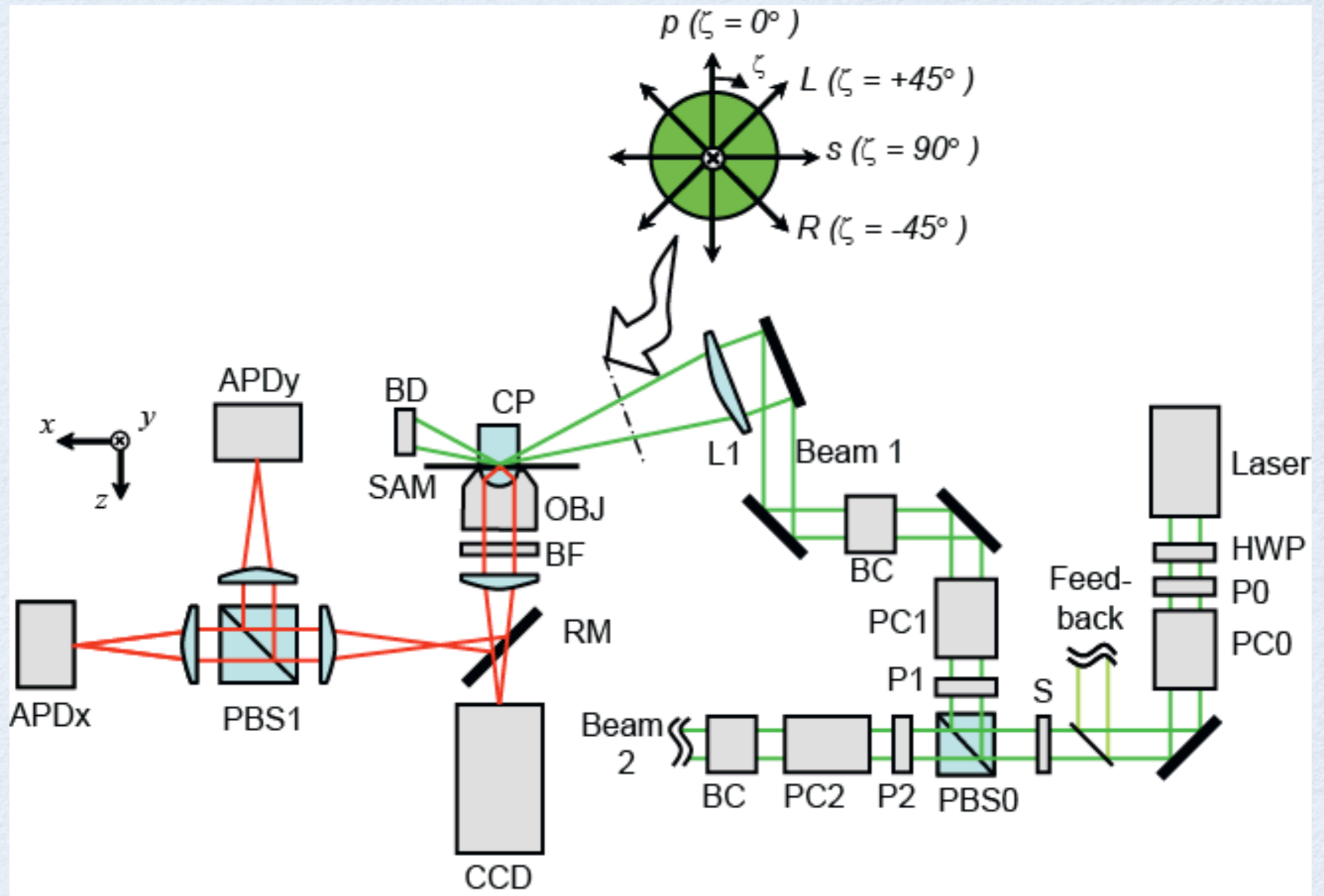
Polarized total internal reflection fluorescence microscopy



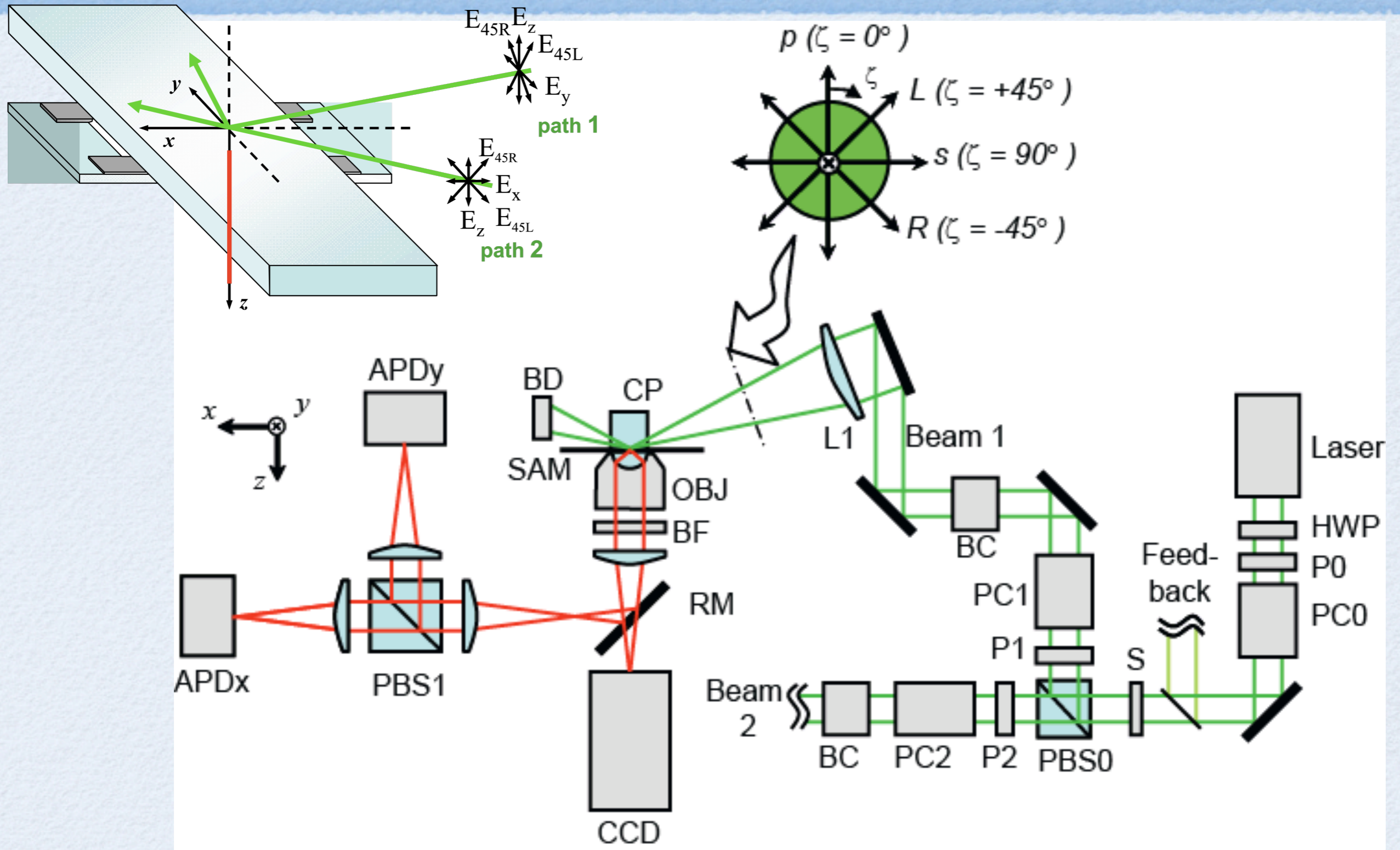
Fluorescence illumination by the evanescent wave eliminates a lot of noise, and importantly, maintains the polarization of the incident light.

To tickle the fluorophore with every possible polarization, we need the incoming light to have at least two different beam directions.

pol-TIRF setup



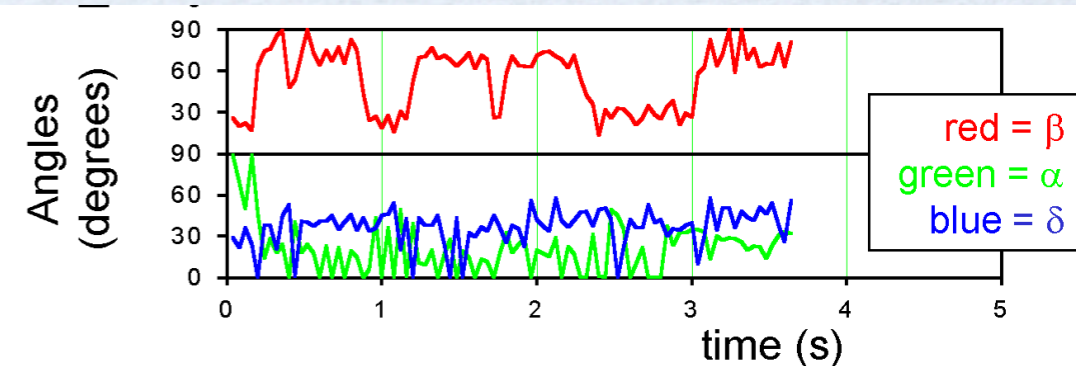
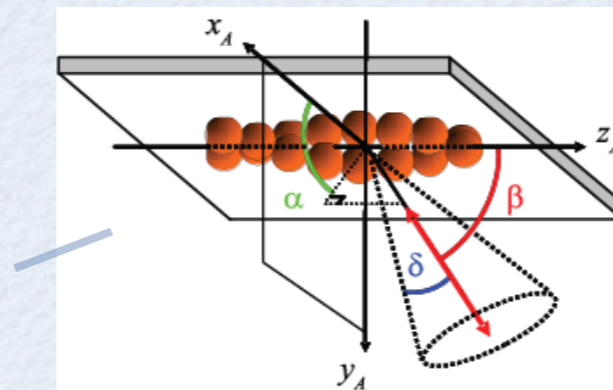
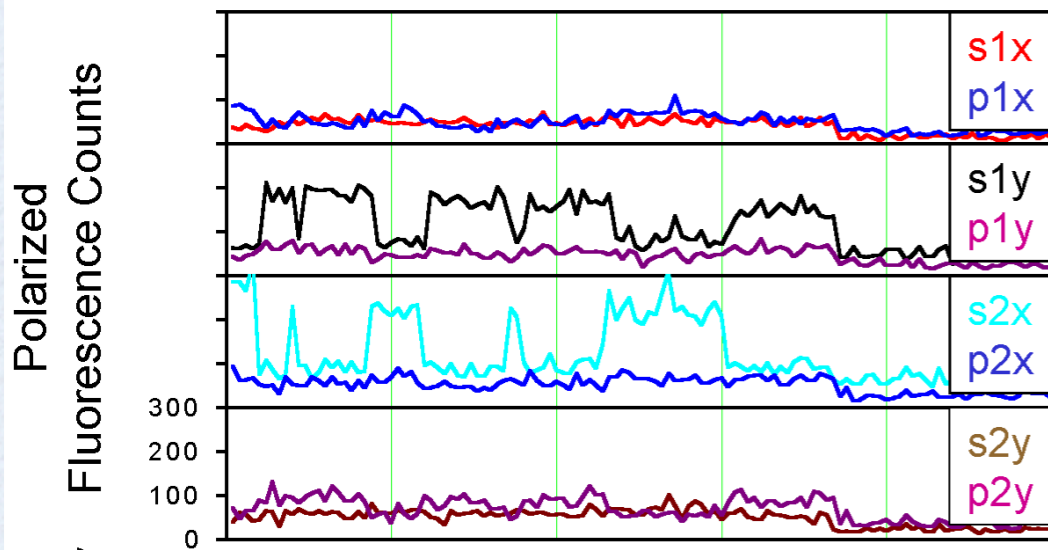
pol-TIRF setup



8 polarized illuminations x 2 detectors = 16 fluorescent intensities per cycle

Current state of the art

Myosin V - 5 μ M ATP



It's a bit more meaningful to convert from lab-frame angles θ, ϕ to actin-frame angles α, β . Even then, however, state of the art calculations give pretty noisy determinations, with pretty poor time resolution.

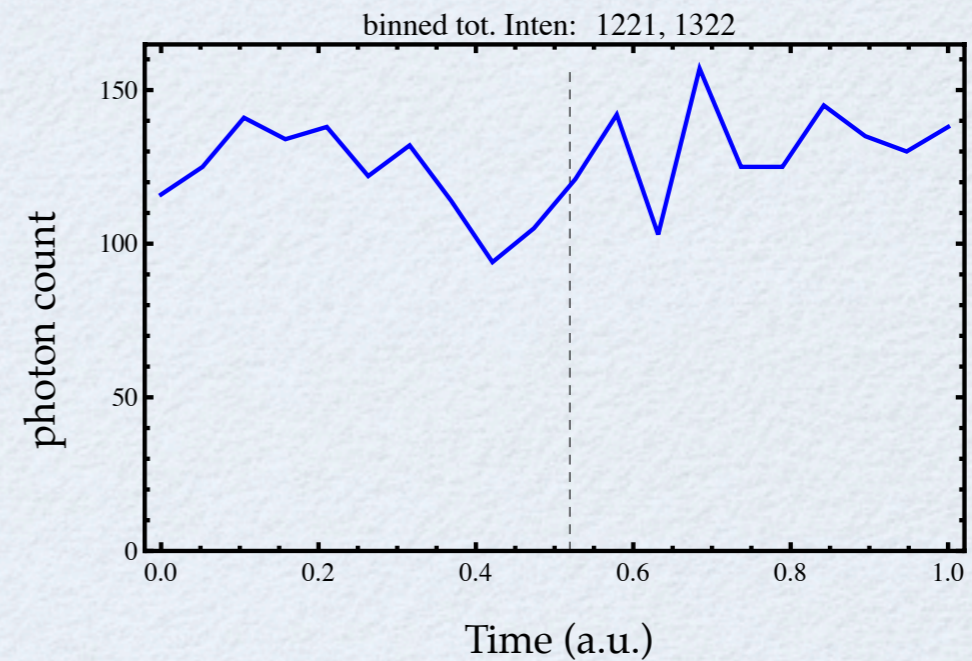
You could easily miss a short-lived state -- e.g. the elusive diffusive-search step (if it exists).

Can't we do better?

JN Forkey et al. Nature 2003

Unfortunately, the total photon counts from a fluorescent probe may not be very informative. Here we divided a time period of interest into 20 bins. There is some Poisson noise in the photon counts, of course.
([ATP]=10uM)

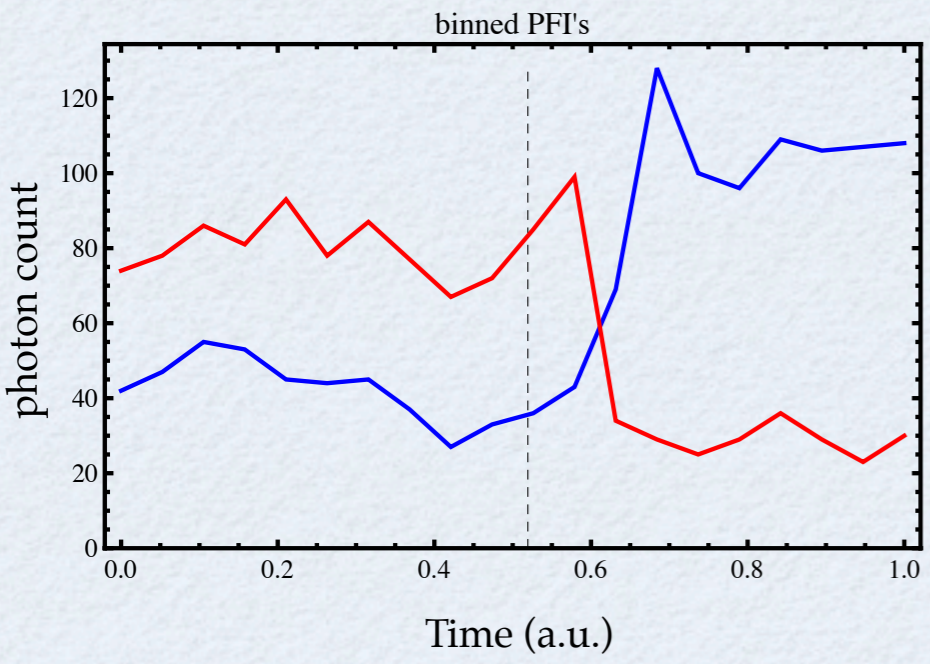
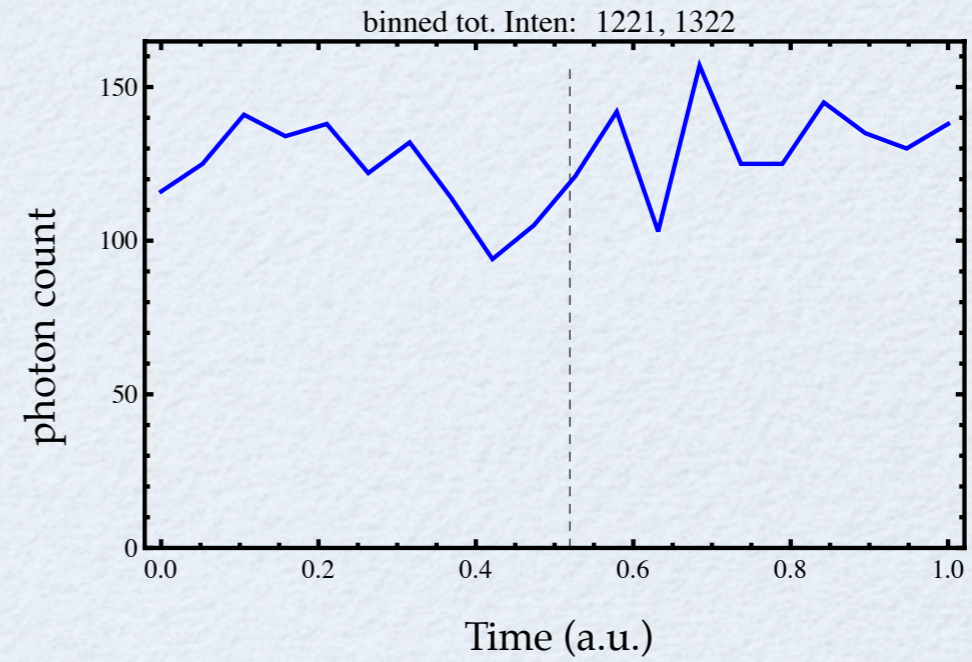
Horizontal axis is time. Vertical axis is binned photon count, PFI = polarized fluorescence intensity



JF Beausang, YE Goldman, and PCN, Meth. Enzymol. 487:431 (2011).

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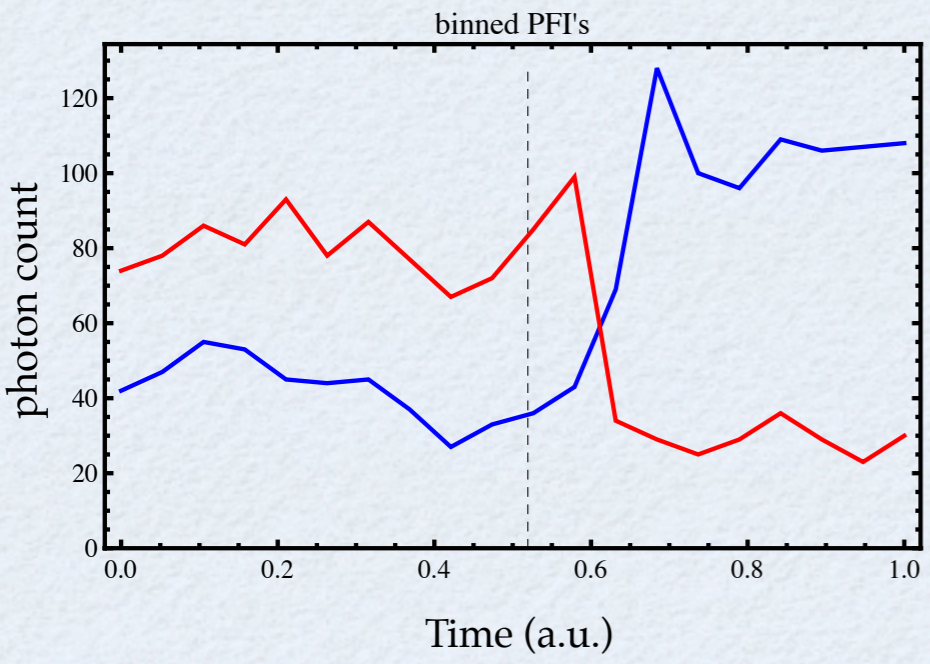
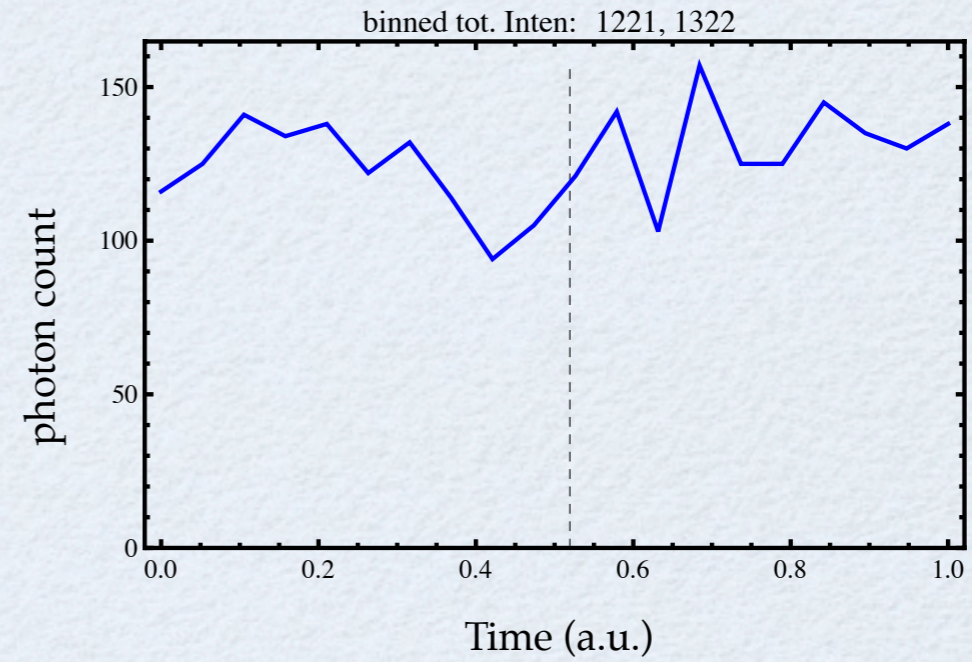


If we classify the photons by polarization and bin them separately, that reveals a definite changepoint. But when exactly did it occur? Probably not at the dashed line shown, but how can we be more precise?

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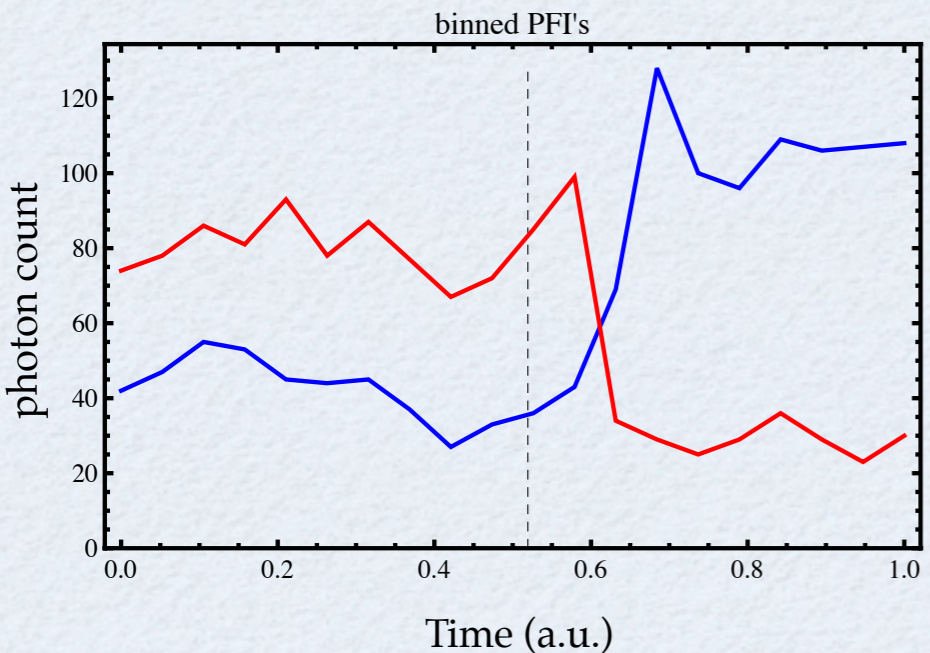
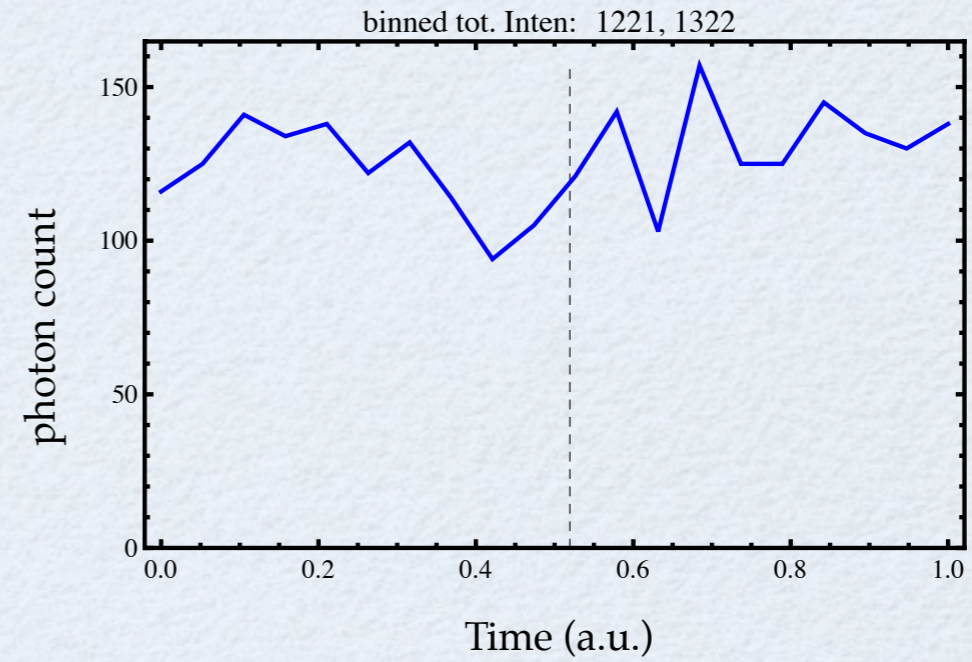
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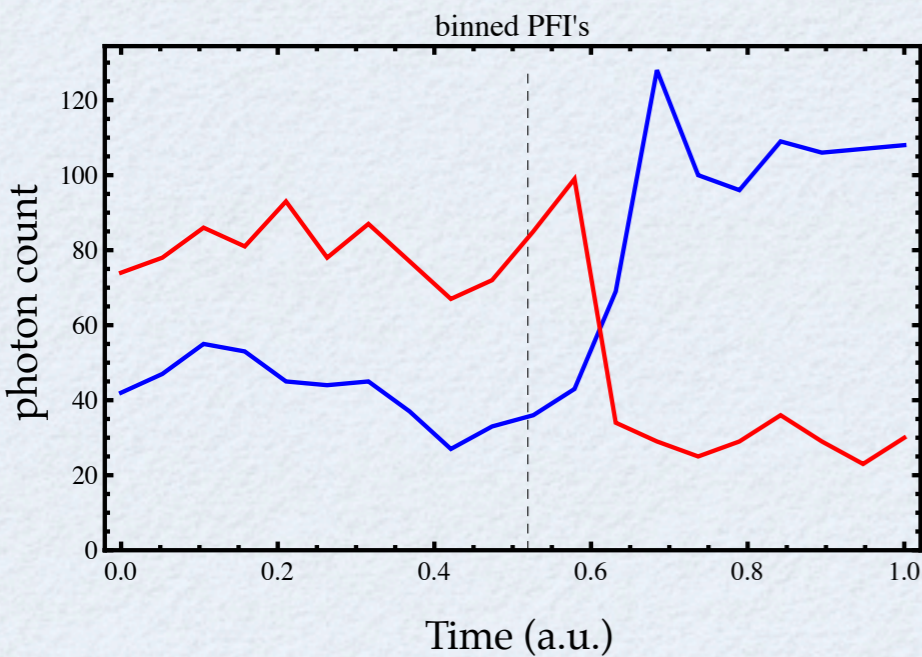
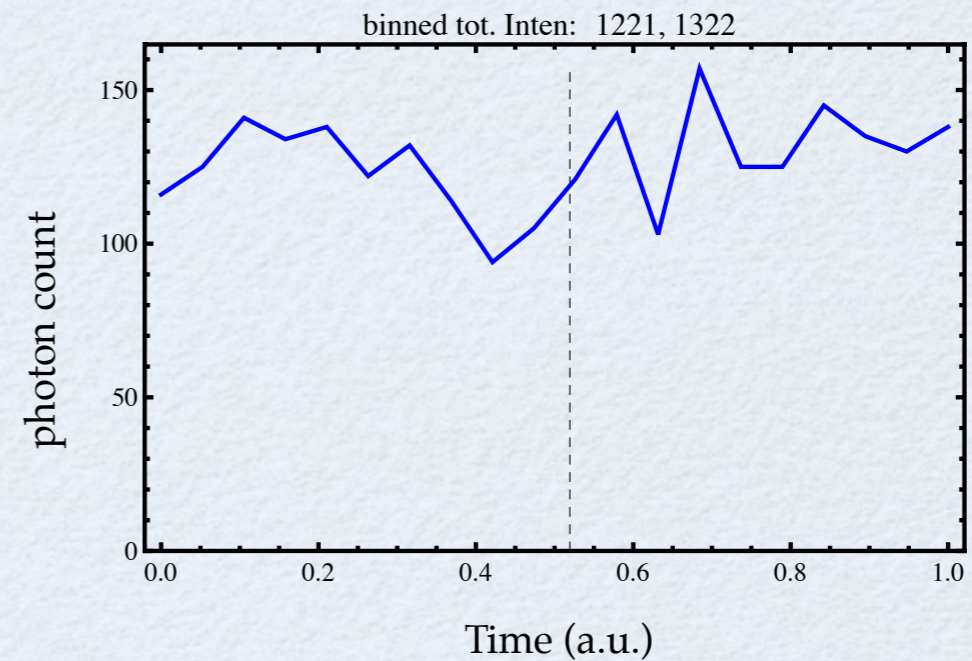
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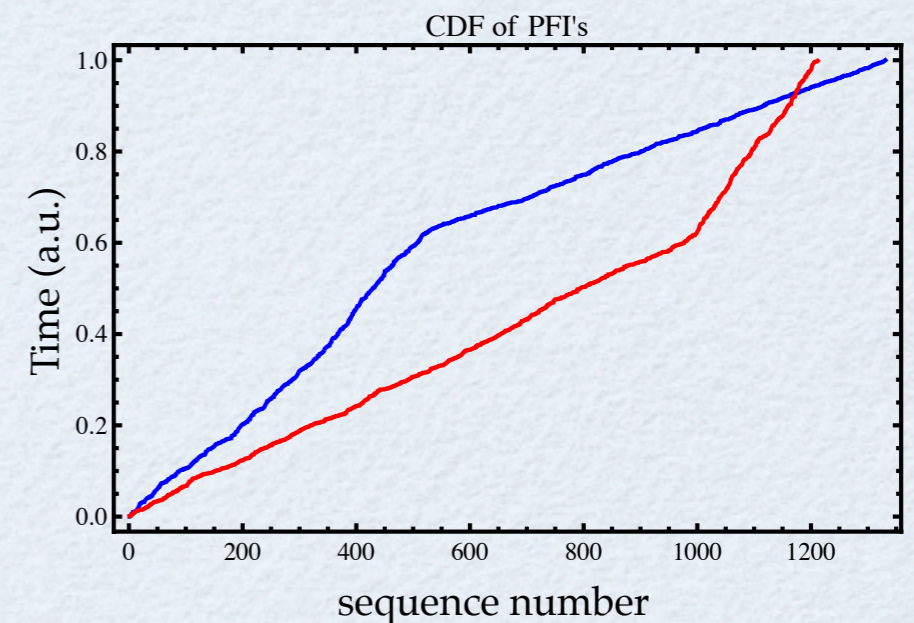
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Can we evade the cruel logic of photon statistics?

It turns out that *binning the data destroyed some information*. Something magical happens if instead of binning, we just we plot photon arrival time versus photon sequence number. Despite some ripples from Poisson statistics, it's obvious that each trace has a sharp changepoint, and moreover that the two changepoints found independently in this way are simultaneous.

(A similar approach in the context of FRET was pioneered by Haw Yang.)



JF Beausang, YE Goldman, and PCN, Meth. Enzymol. 487:431 (2011).

Now that I have your attention

- *Why did that trick work?* How did we get such great time resolution from such cruddy data?
- *How well does it work?* If we have even fewer photons, for example because a state is short-lived, how can we quantify our confidence that any changepoint occurred at all?
- *Could we generalize and automate this trick?* Ultimately we'll want to handle data with multiple polarizations, and find lots of changepoints.

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The appropriate tool is **maximum-likelihood analysis**:

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The appropriate tool is **maximum-likelihood analysis**:

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$$R = n/t_* , \quad R' = n'/(T - t_*)$$

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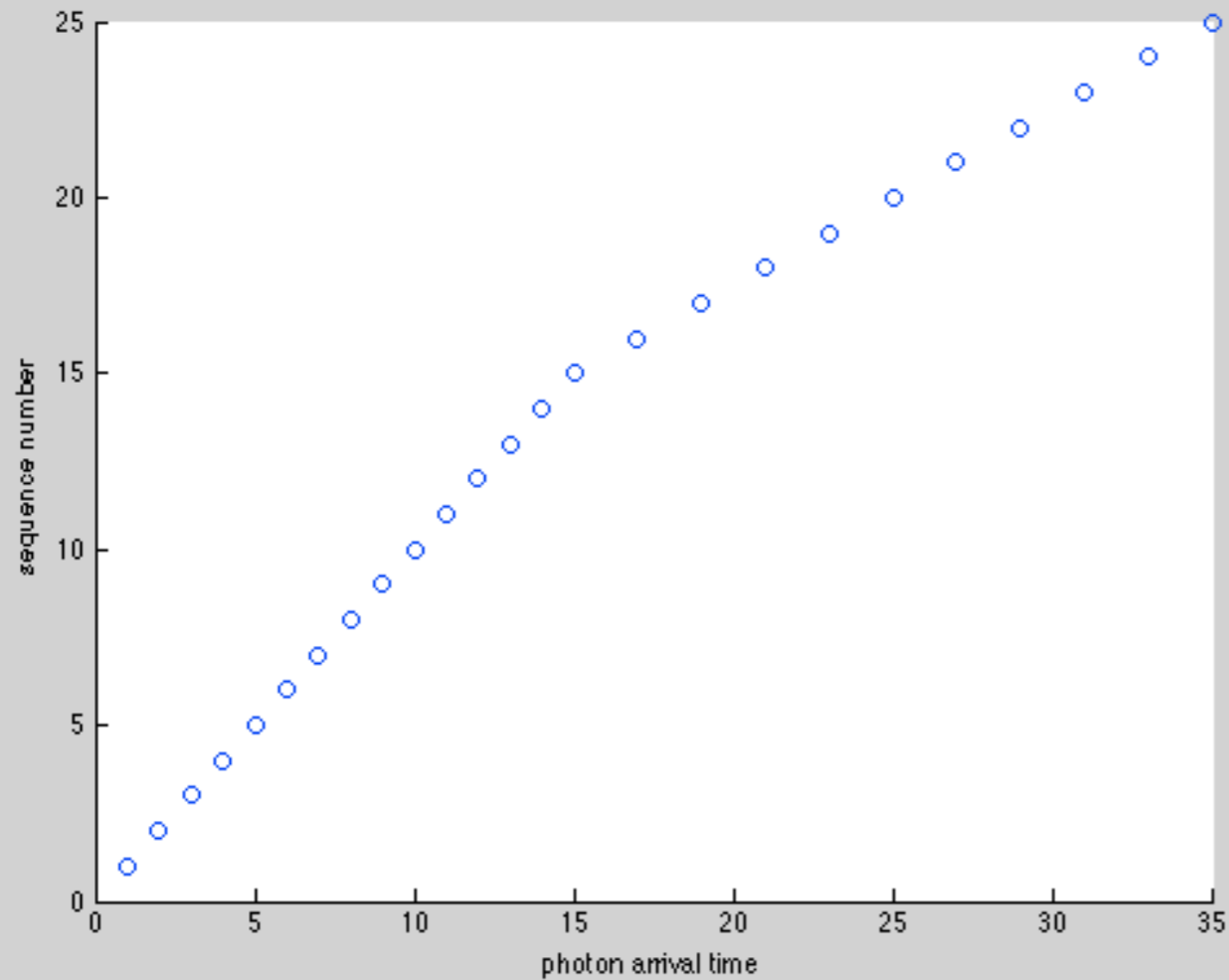
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More interestingly, we can substitute these optimal rates into the formula for P to find the likelihood as a function of putative changepoint:

Application

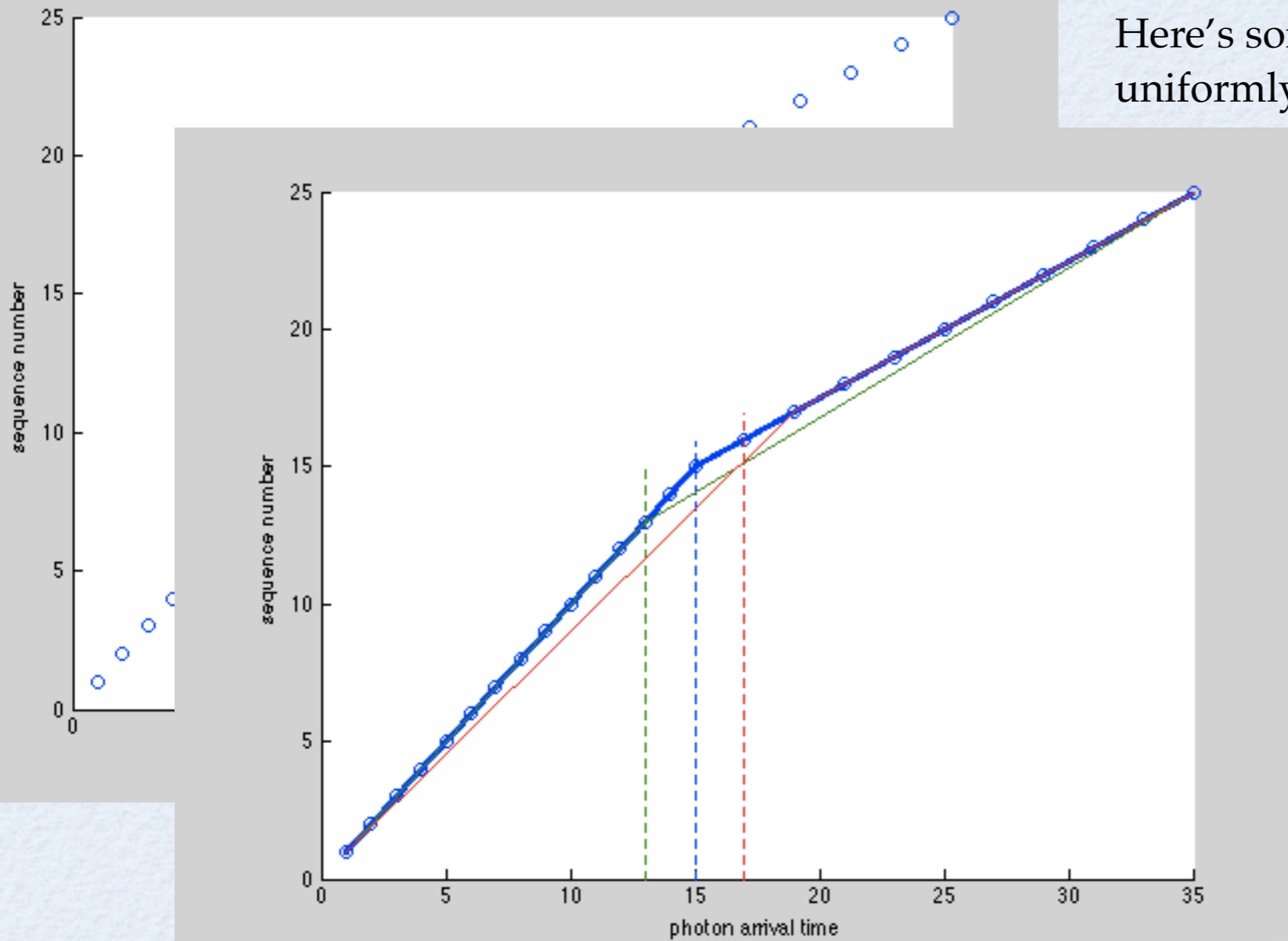
Application



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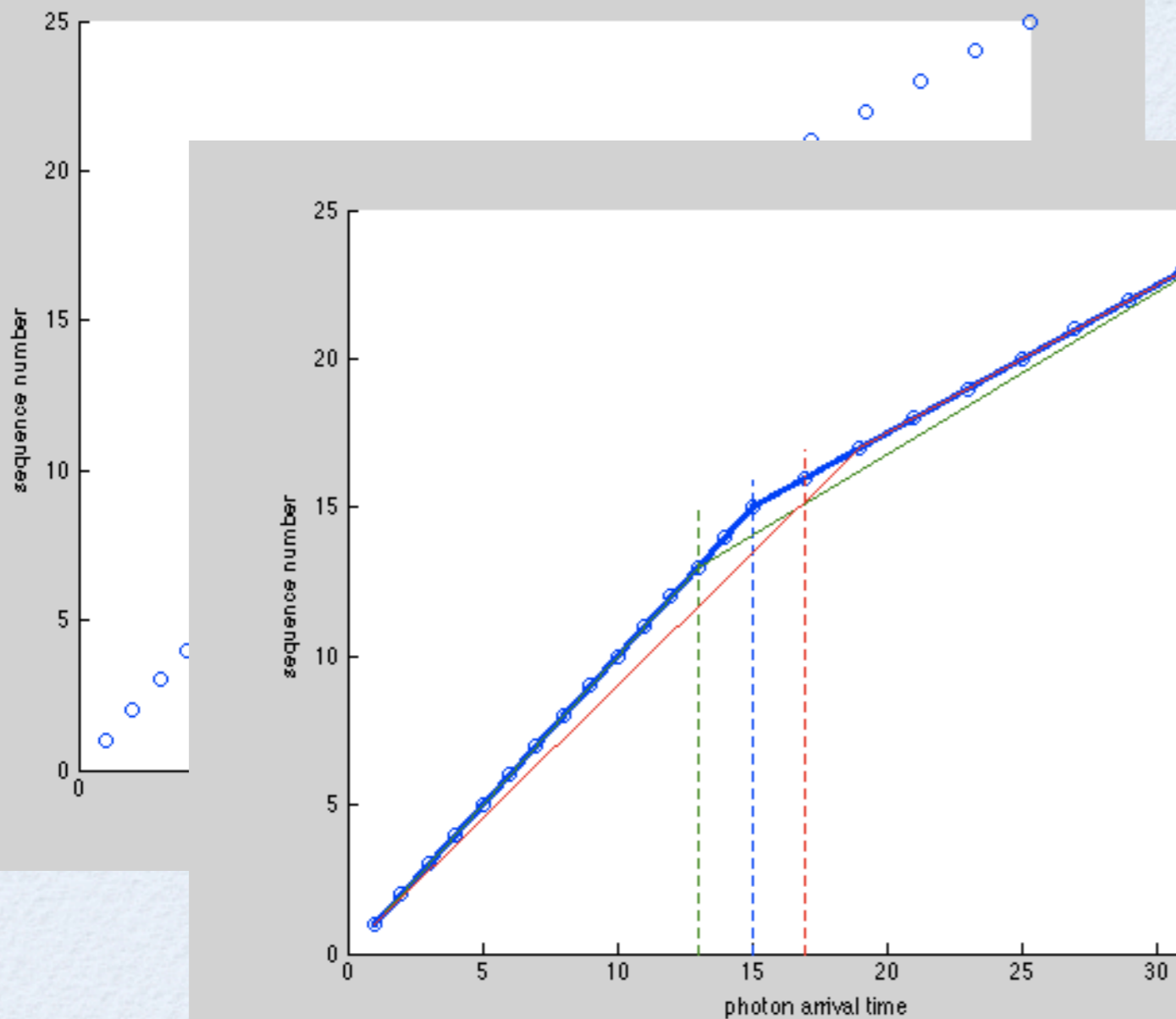


Here are two lines corresponding to non-optimal choices of the changepoint. We'd like to see the likelihood function and how it selects the "right" changepoint, which for fake data is known.

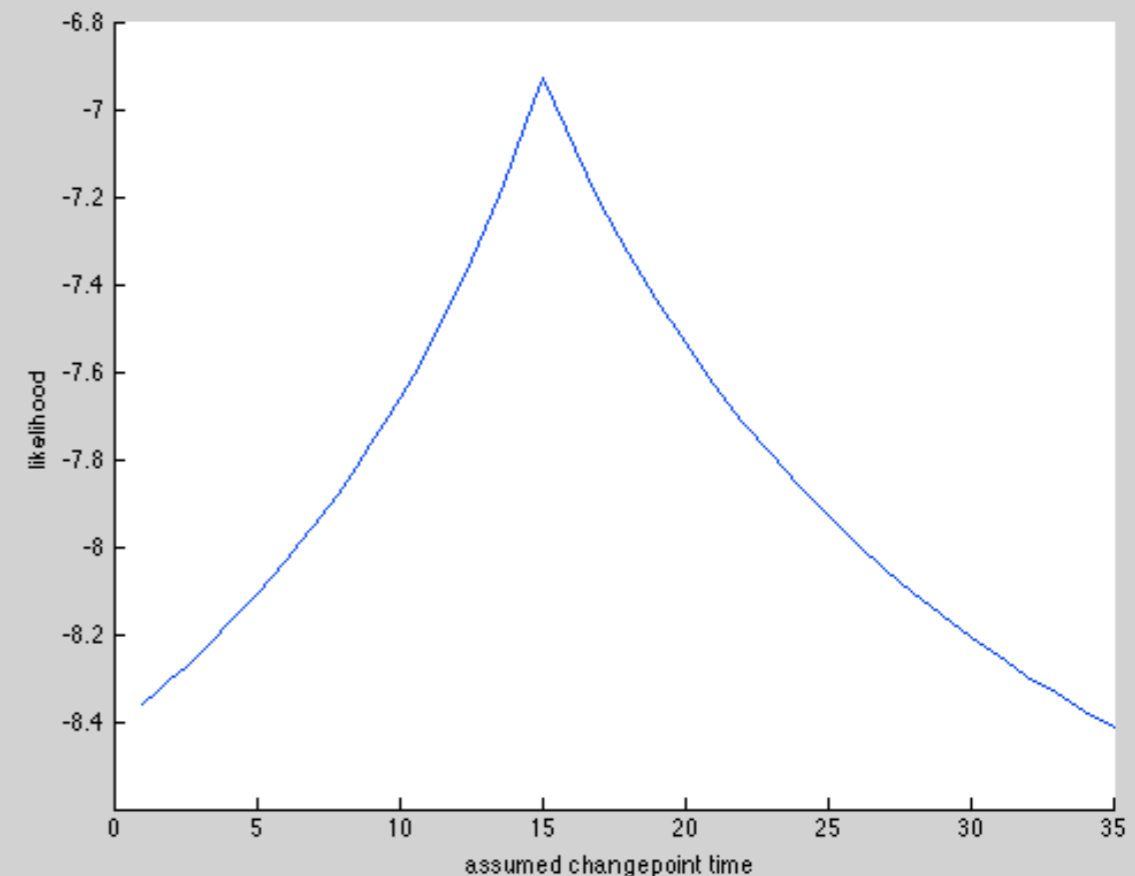
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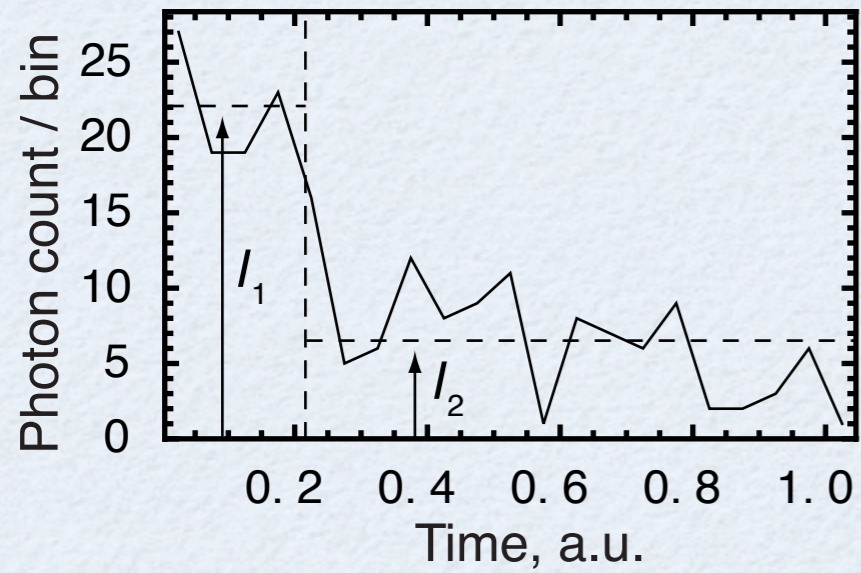
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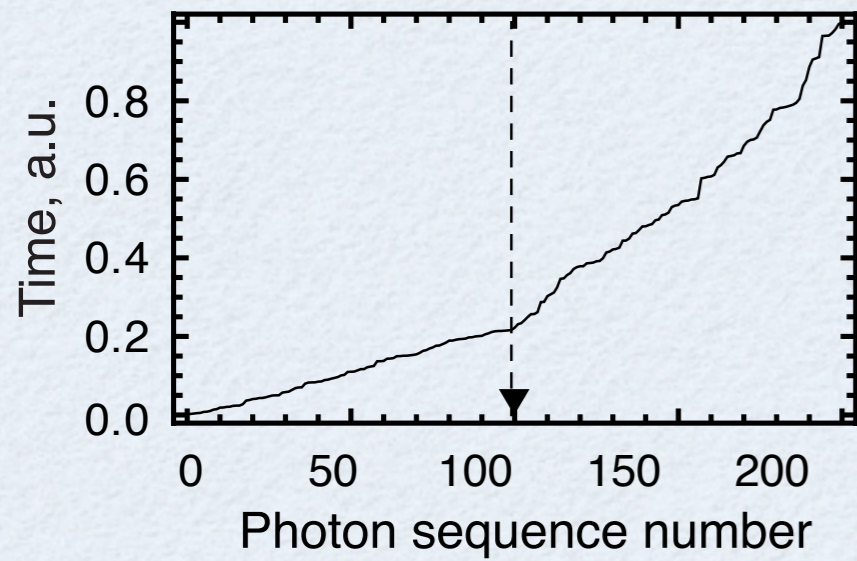


Here is our log-likelihood function as a function of putative changepoint time.

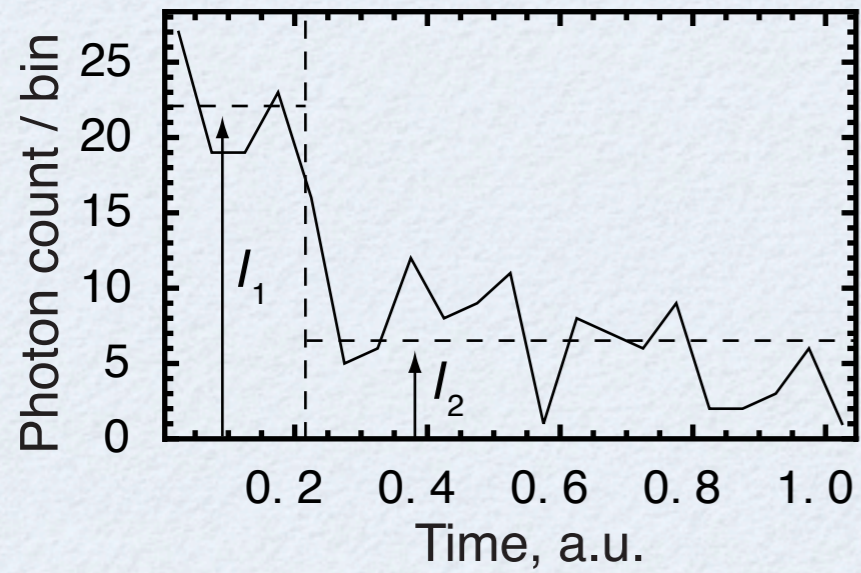




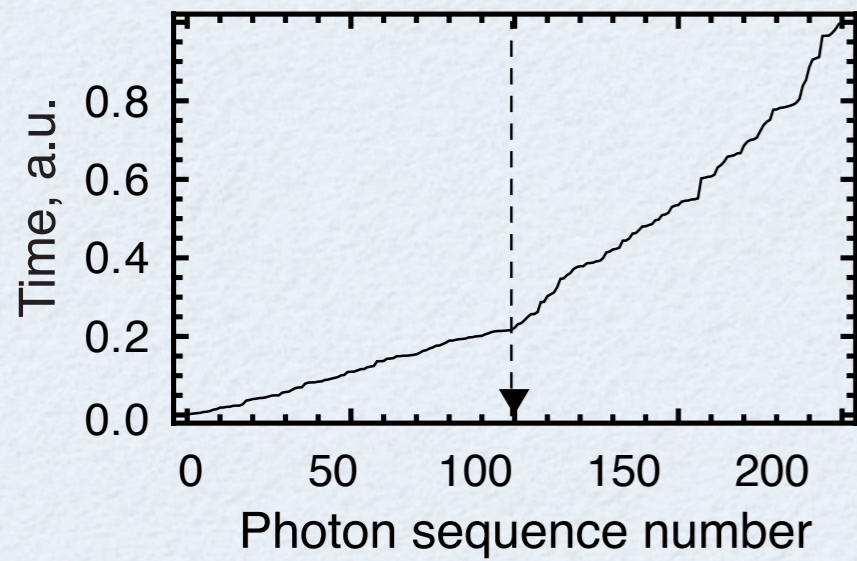
Left: Some more realistic (Poisson-arrival) simulated data, shown in traditional binned form and in the improved version.



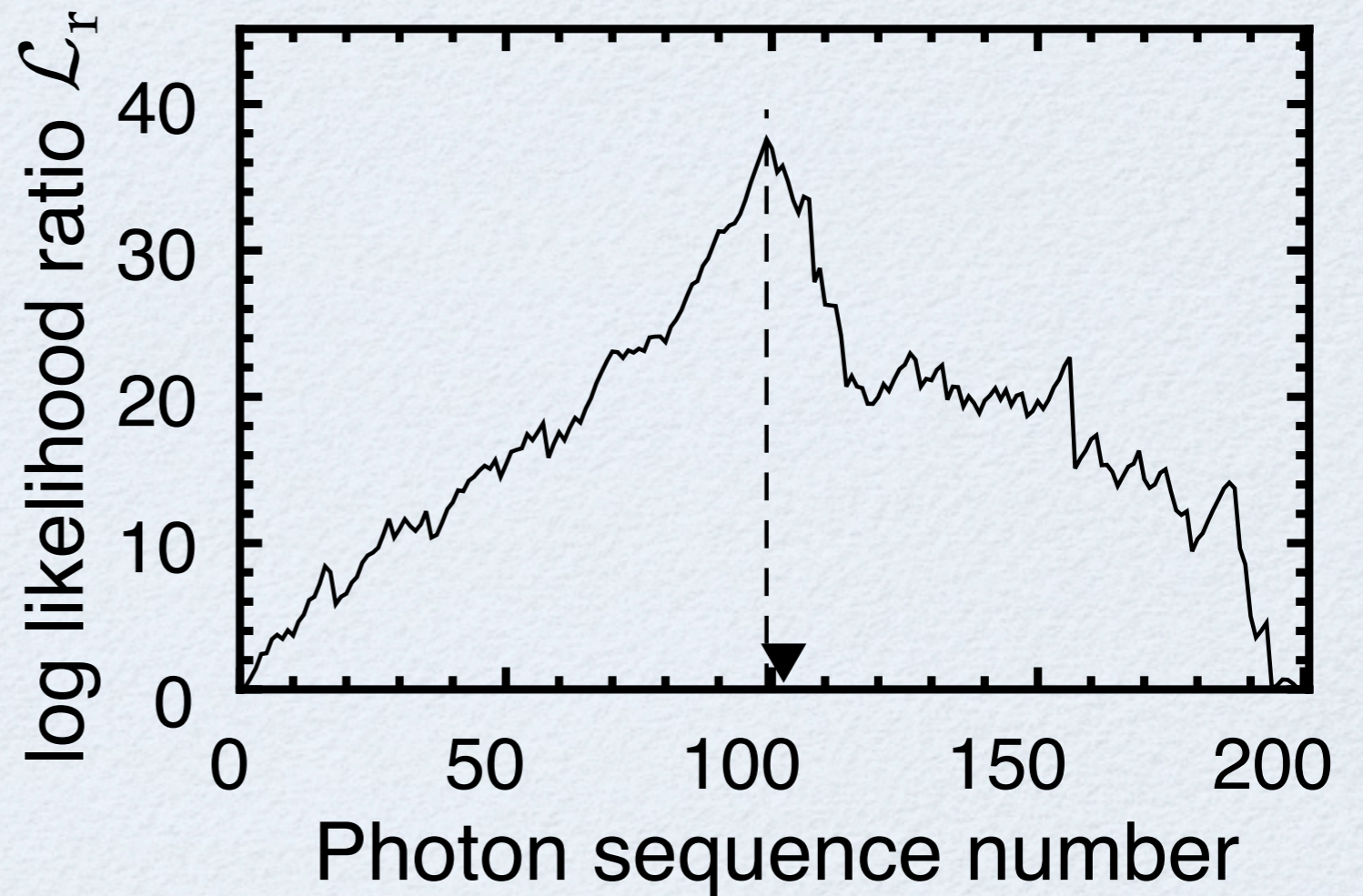
Right: Likelihood function for placement of the changepoint. Dashed line, maximum-likelihood point. Black triangle: Actual changepoint used in the simulation.



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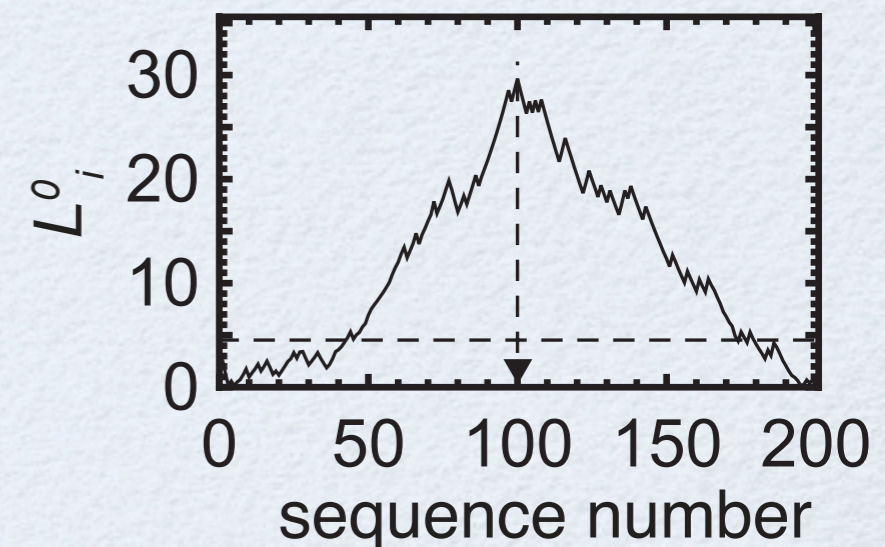
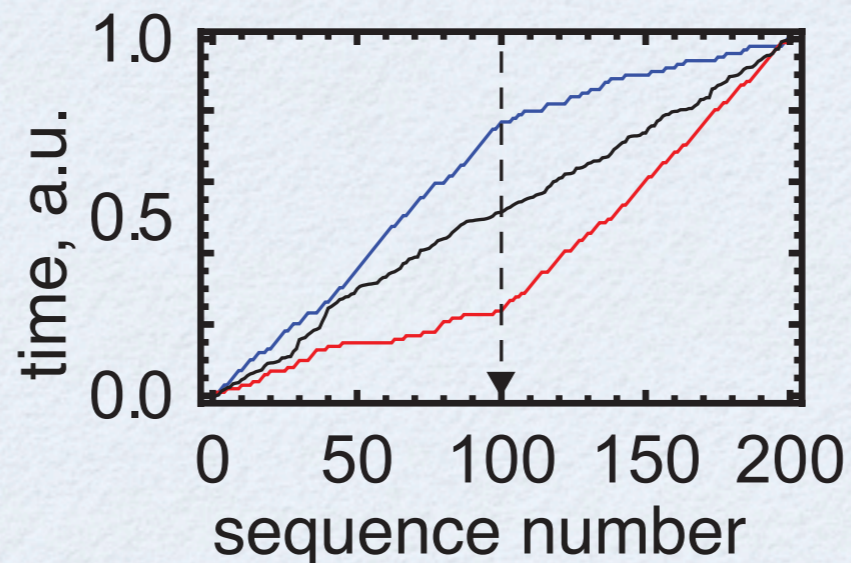
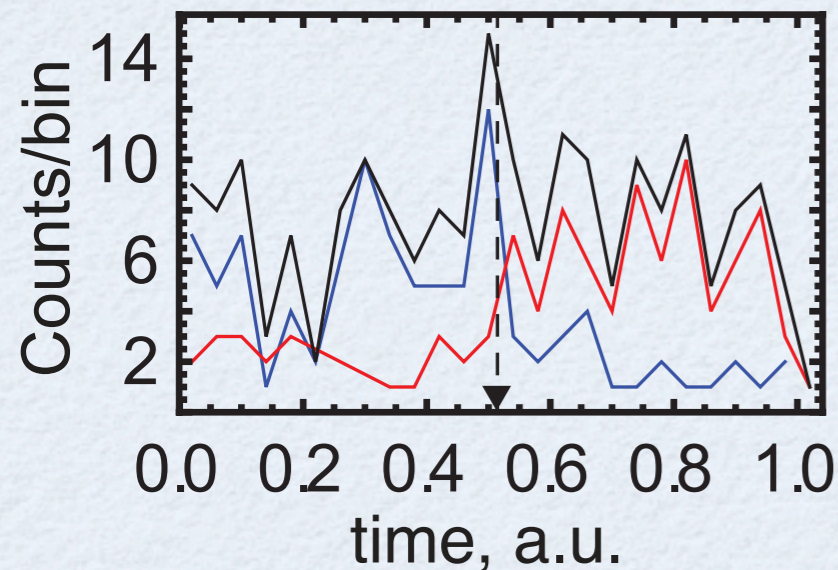


Right: Likelihood function for placement of the changepoint. Dashed line, maximum-likelihood point. Black triangle: Actual changepoint used in the simulation.



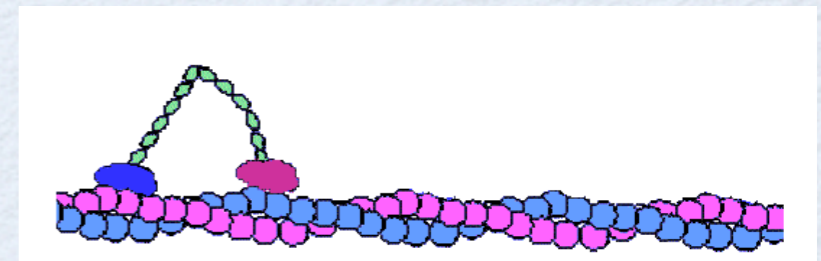
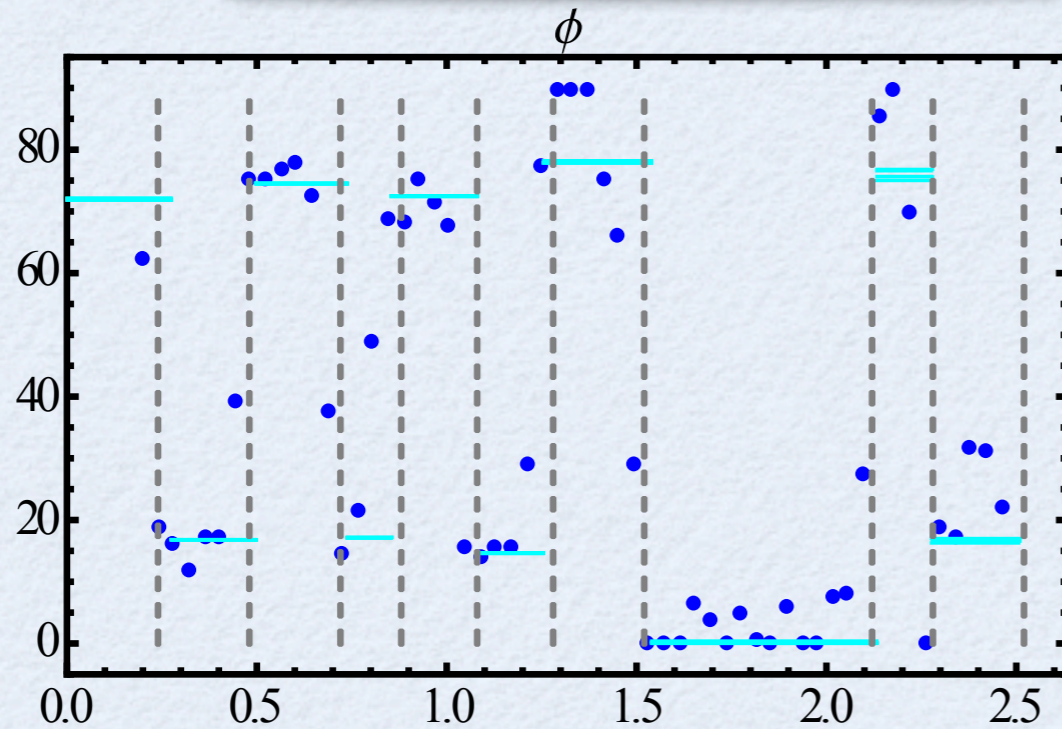
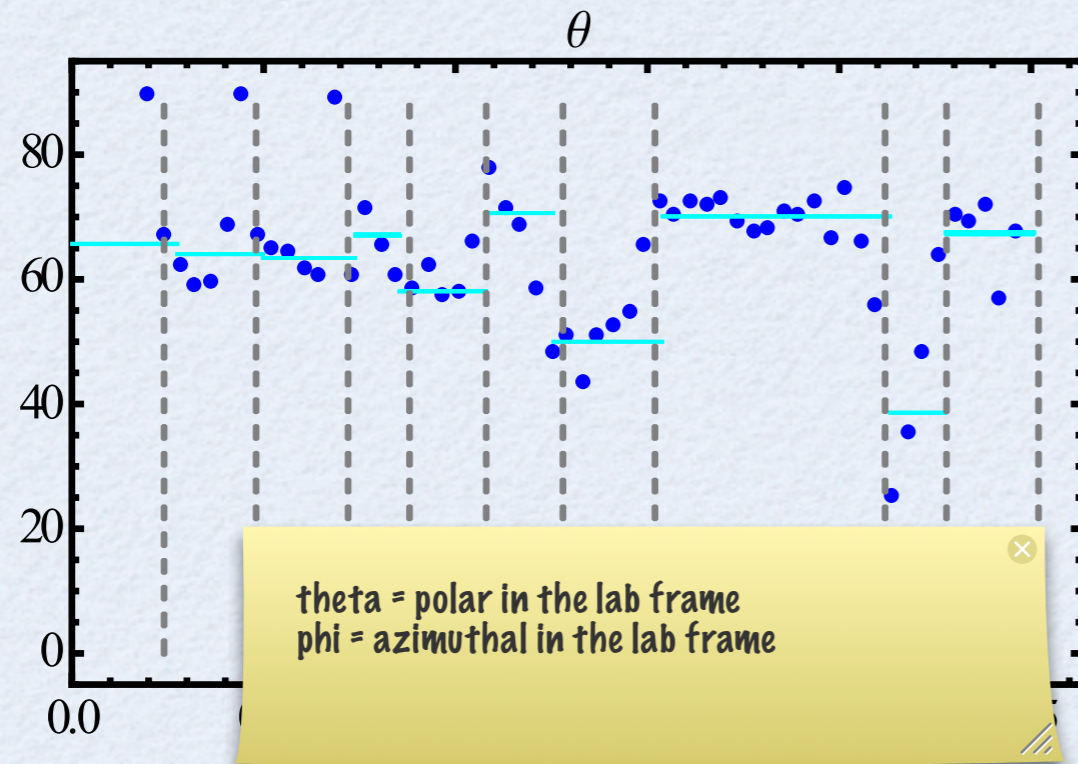
Oh, yes -- the method also works on multiple-channel data. *Left*: one channel (red) starts with rare photons, then jumps to higher intensity. Another channel (blue) does the opposite. The sum of the intensities (black) doesn't change much at all.

Middle: "kink" representations of the same data. *Right*: both channels contribute to a likelihood function with a robust peak, **even though there were only a total of just 200 photons in the entire dataset.**



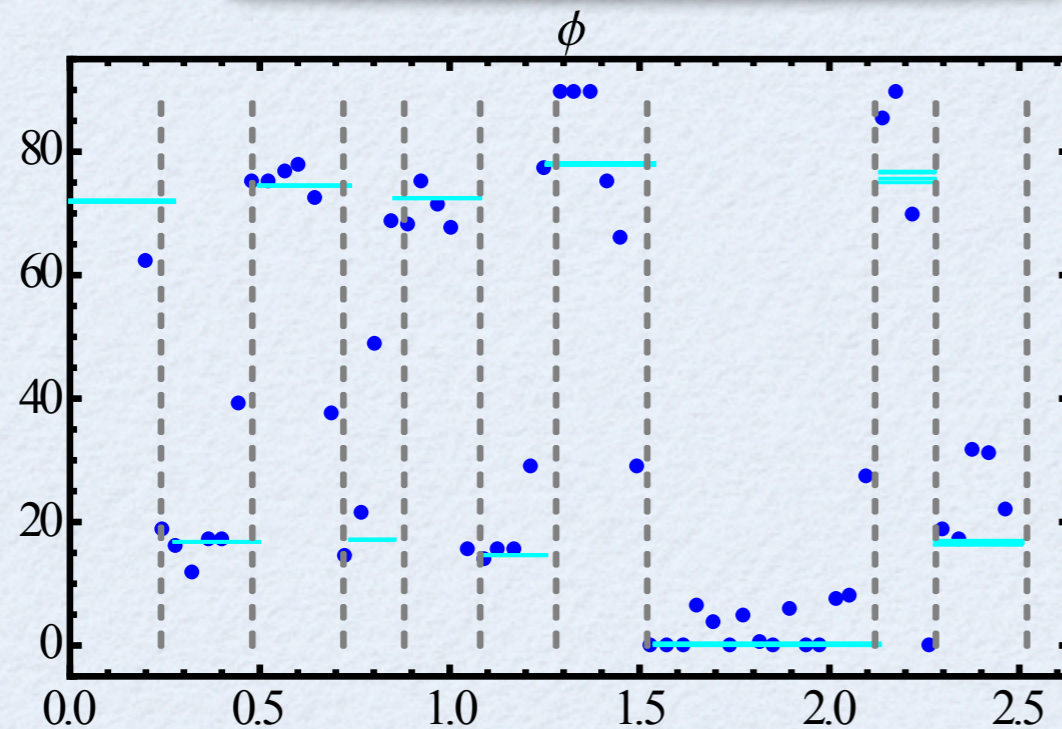
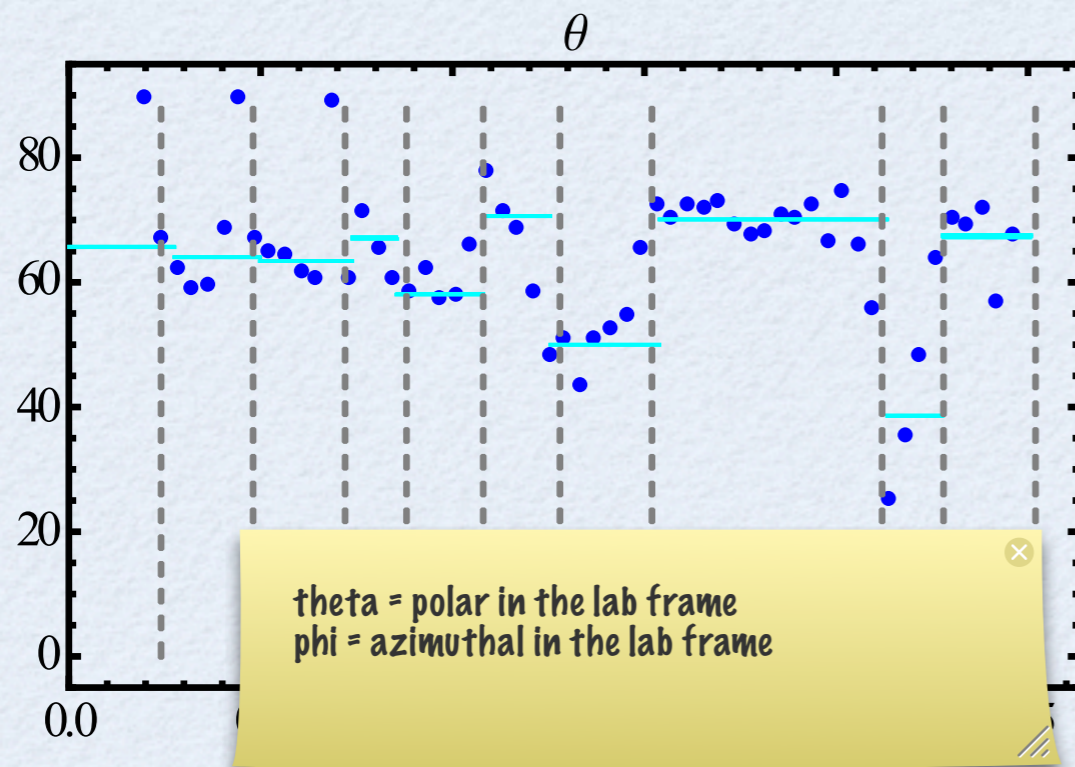
JF Beausang, YE Goldman, and PCN, Meth. Enzymol. 487:431 (2011).

Payoff



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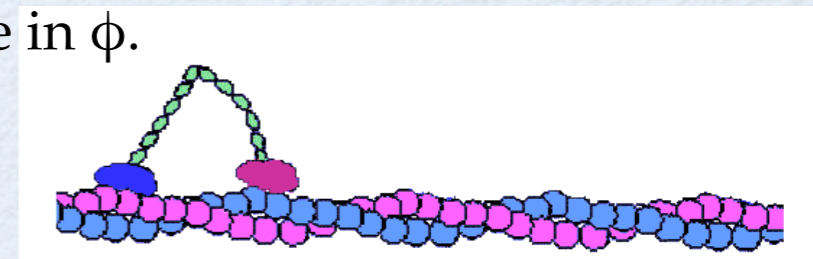


Oh, yes--it also works on real experimental data.

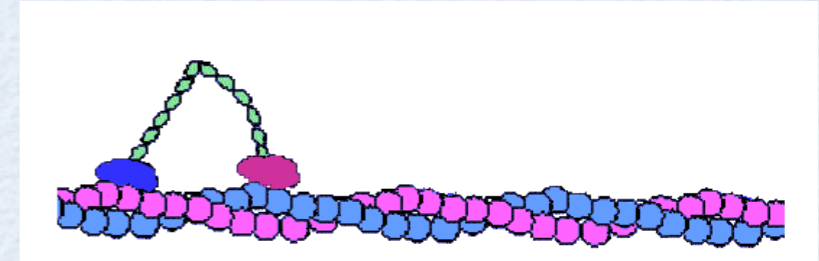
Now we can get back to the original motivation. Previously, people would take data from multiple polarizations, bin it, and pipe the inferred intensities into a maximum-likelihood estimator of the orientation of the fluorophore. That procedure leads to the rather noisy dots shown here.

One problem is that if a transition happens in the middle of a time bin, then the inferred orientation in that time bin can be crazy.

Here the solid lines are the inferred orientations of the probe molecule during successive states defined by changepoint analysis. We see a nice alternating stride in ϕ .



Summary Part II




- *When you only get a million photons, you'd better make every photon count.
- *A simple maximum-likelihood analysis accomplishes this.
- *In the context of TIRF it can dramatically improve the tradeoff between time resolution and accuracy.

Part III: Parallel recordings from dozens of individual neurons

- * Sometimes suggests a new kind of measurement that tests a model more stringently, or distinguishes two different models more completely, than previous measurements.
- * Sometimes our model is not obviously connected with what we can actually measure experimentally, and we need to make a connection.
- * **Sometimes the model that interests us involves the behavior of actors that we can only see indirectly in our data; theory may be needed to separate them out from each other, and from noise.**

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Sources of energy

Experiments done in the lab of Vijay Balasubramanian (Penn).

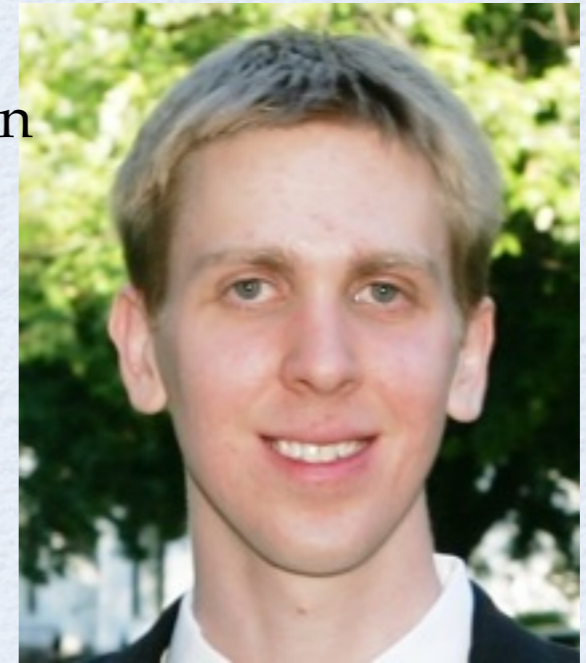
Sources of energy



Experiments done in the lab of Vijay Balasubramanian (Penn).

Kristy Simmons, Penn Neuroscience

Jason Prentice, Penn
Physics



(plus Gasper Tkacik.)

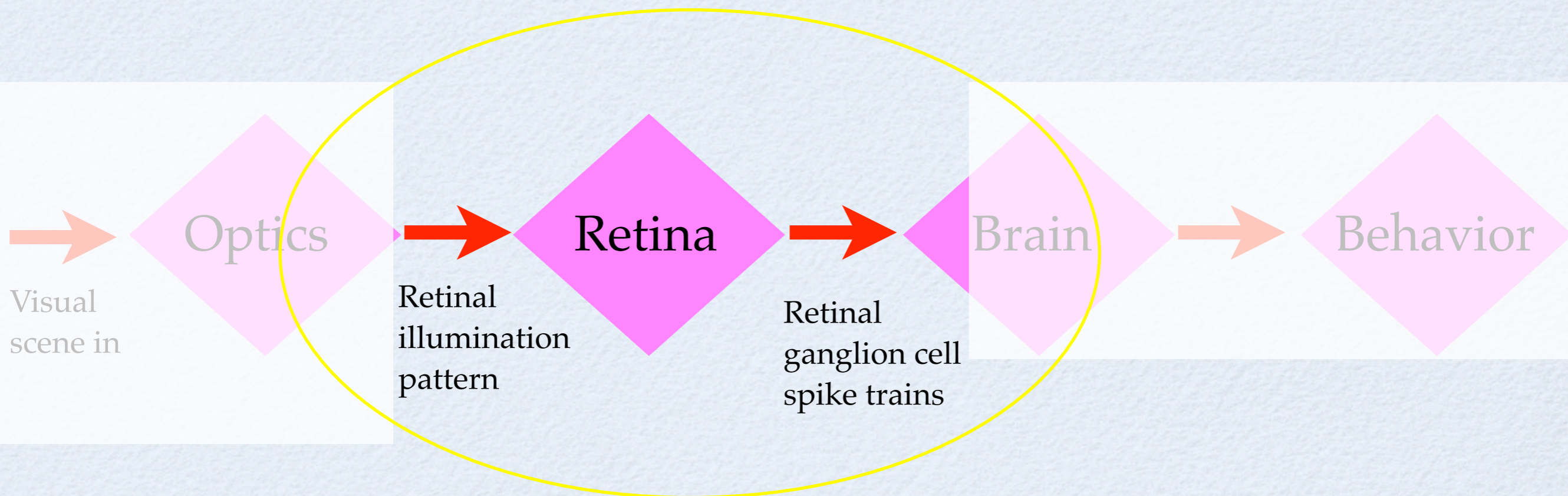
(Many thanks to Michael Berry and Olivier Marre, Princeton; Bart Borghuis, Janelia Farms; Michael Freed and others at Penn Retina Lab; Joerg Sander, U Alberta; Ronen Segev, BGU, Chris Wiggins, Columbia.)

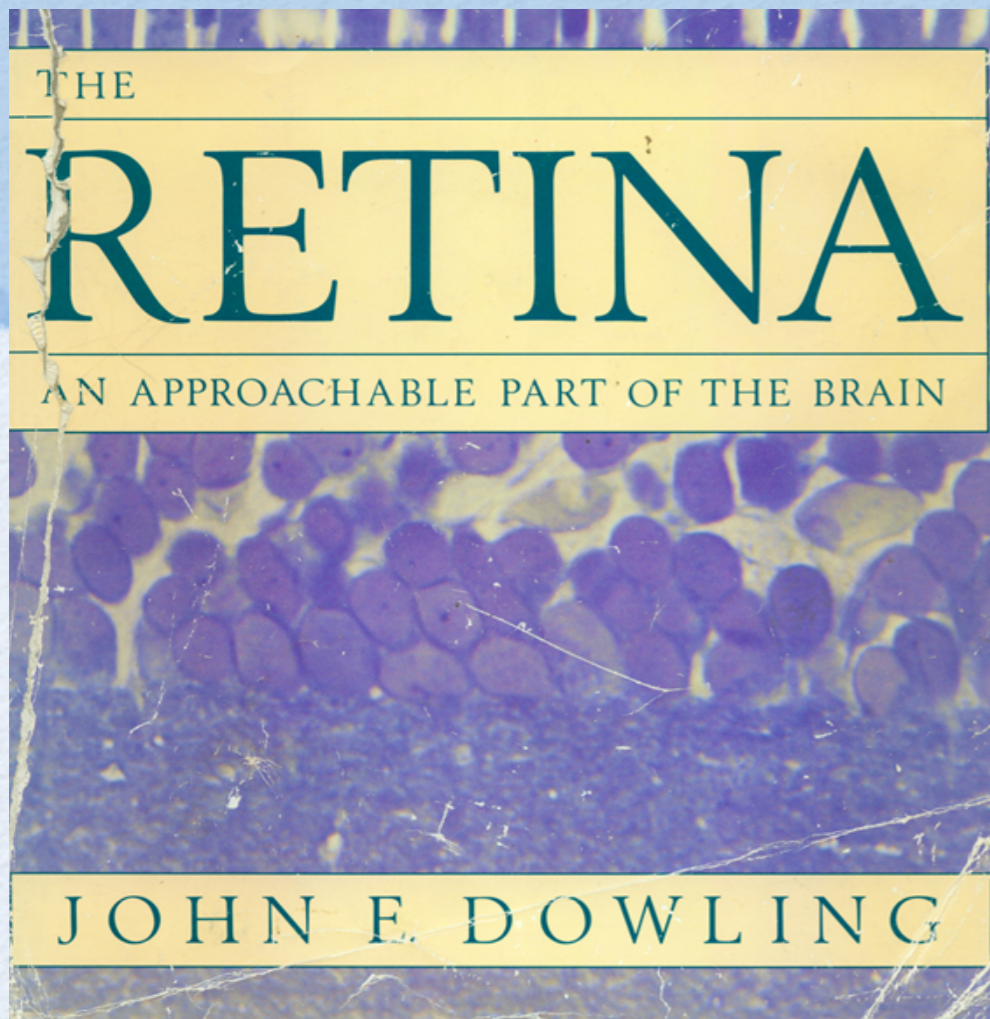


Jan Homann, Penn Physics

Really big picture

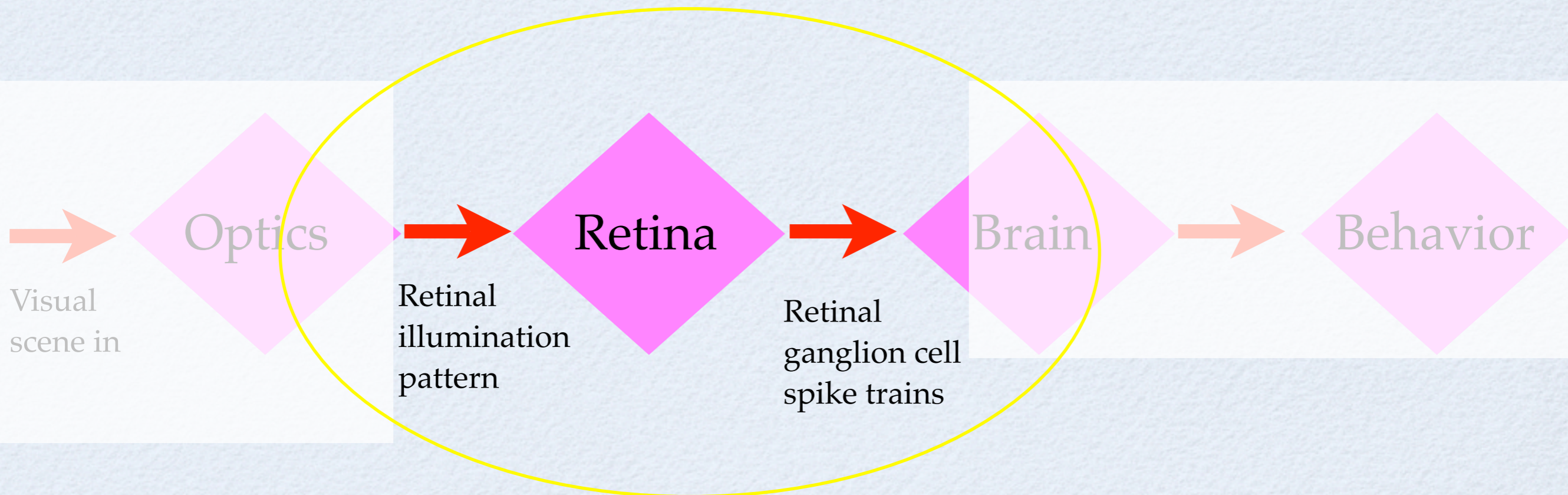
Retina is also an approachable, yet still complex, part of the brain. It's a 2D carpet consisting of "only" three layers of neurons.





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It matters

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(100%) Thu 4:21

Artificial retina more capable of restoring normal vision; animal study shows including retina's neural 'code' improved prosthetic

http://www.sciencedaily.com/releases/2010/11/101116102647.htm

shiela nirenberg

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ScienceDaily (Nov. 16, 2010) — Researchers have developed an artificial retina that has the capacity to reproduce normal vision in mice. While other prosthetic strategies mainly increase the number of electrodes in an eye to capture more information, this study concentrated on incorporating the eye's neural "code" that converts pictures into signals the brain can understand.

The research was presented at

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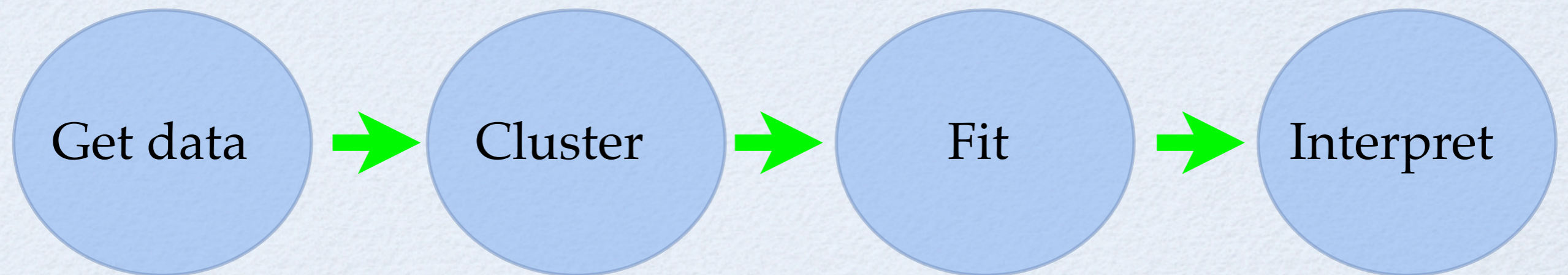
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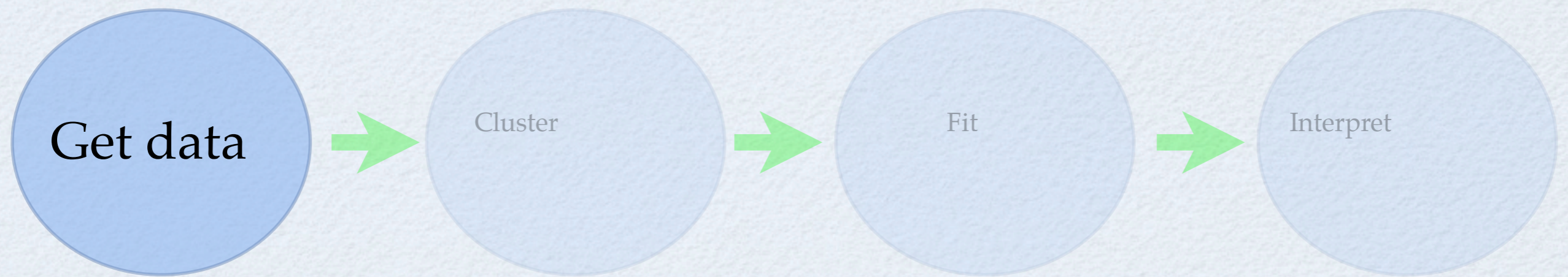
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The research was presented at

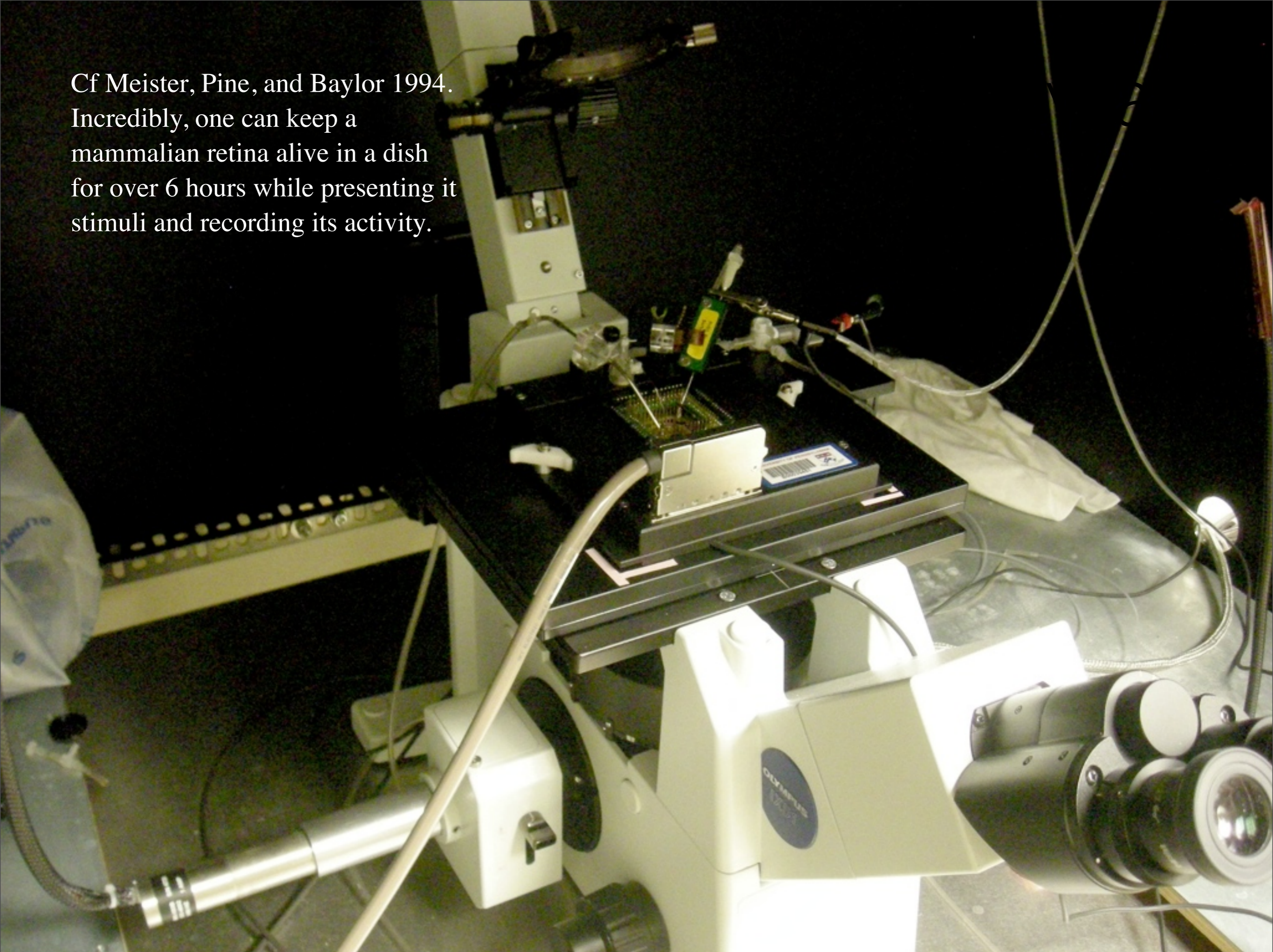
Summary, Part III



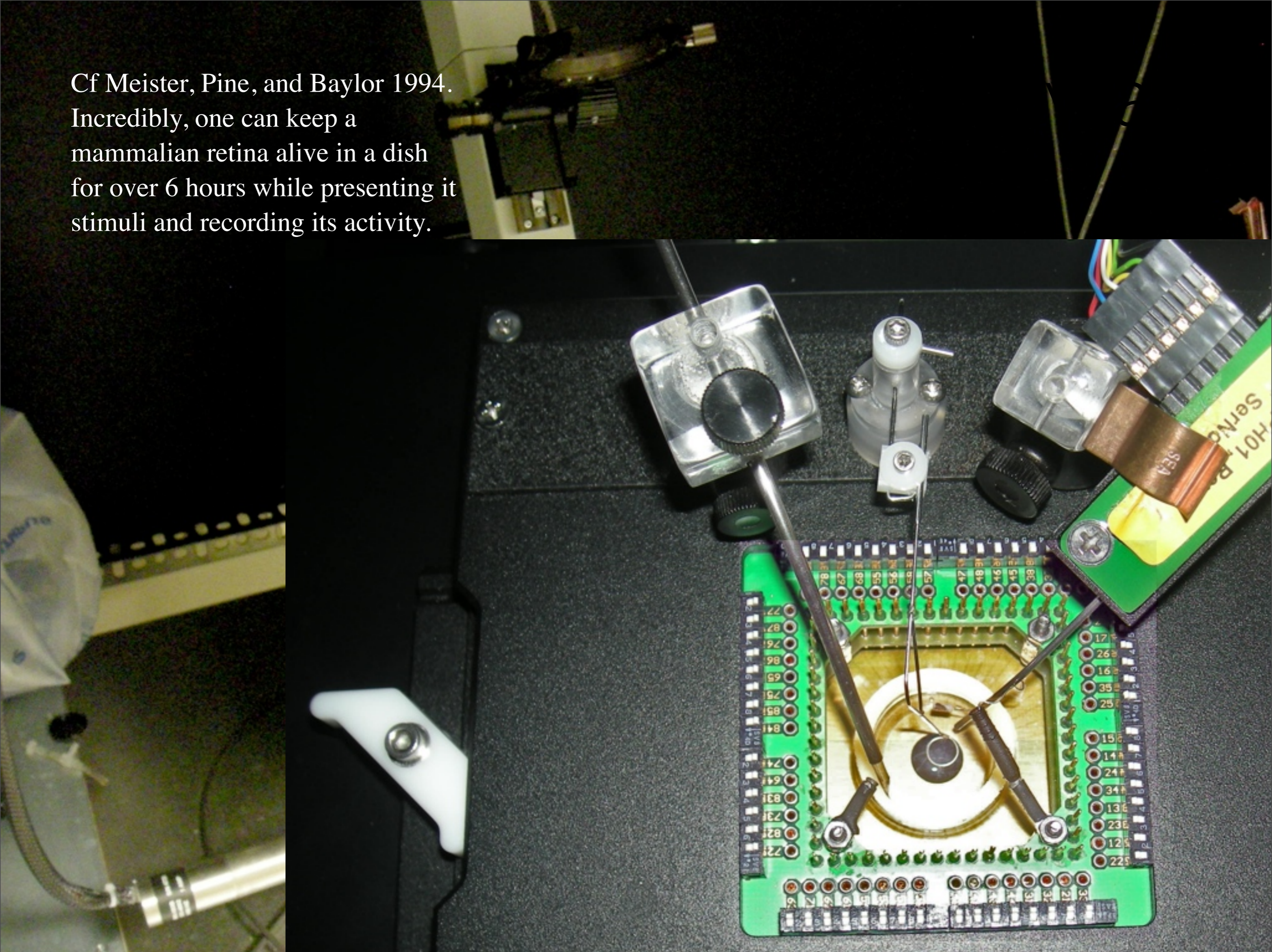
1. Experiment
2. Clustering
3. Fitting
4. Performance



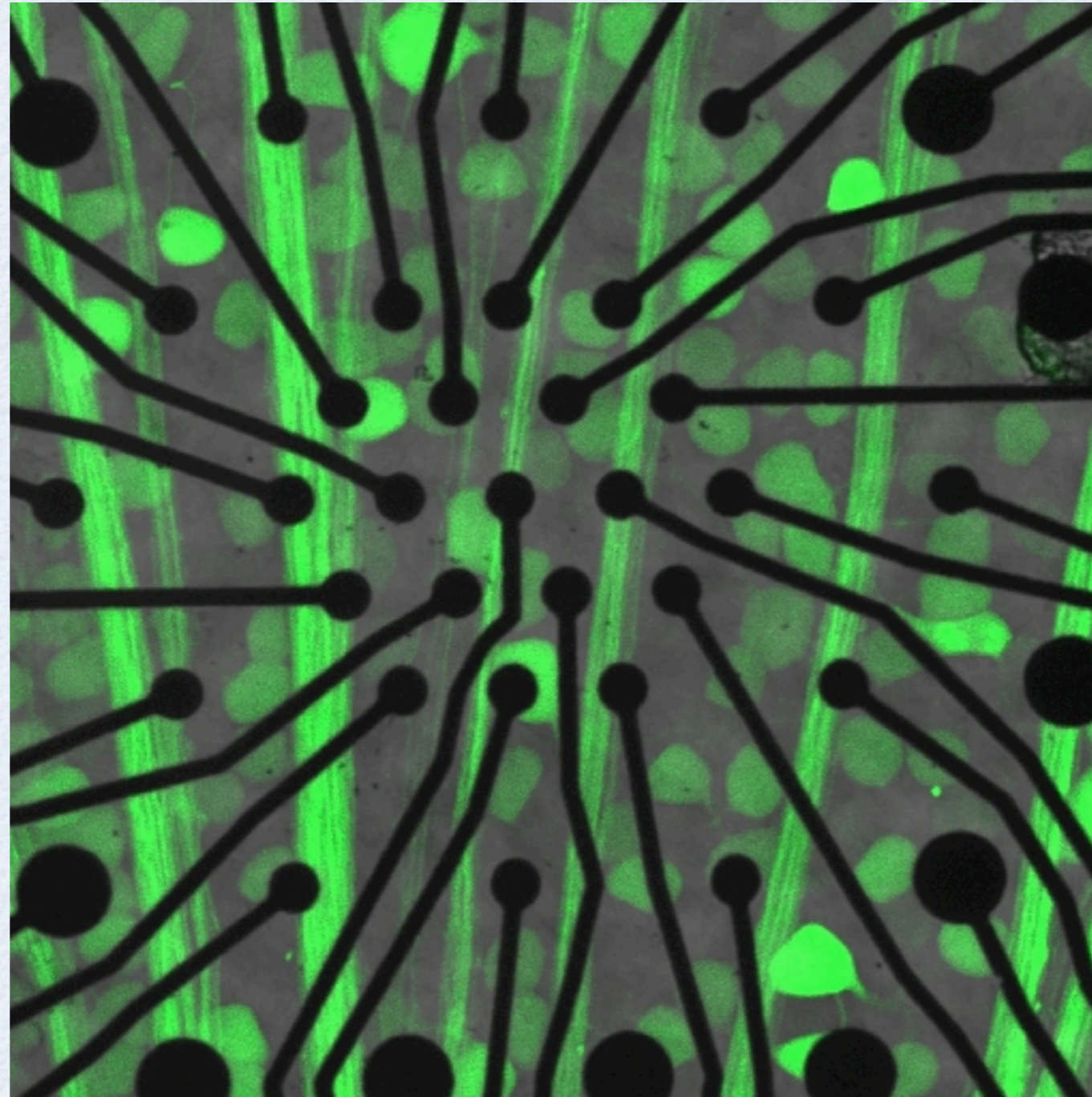
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Incredibly, one can keep a
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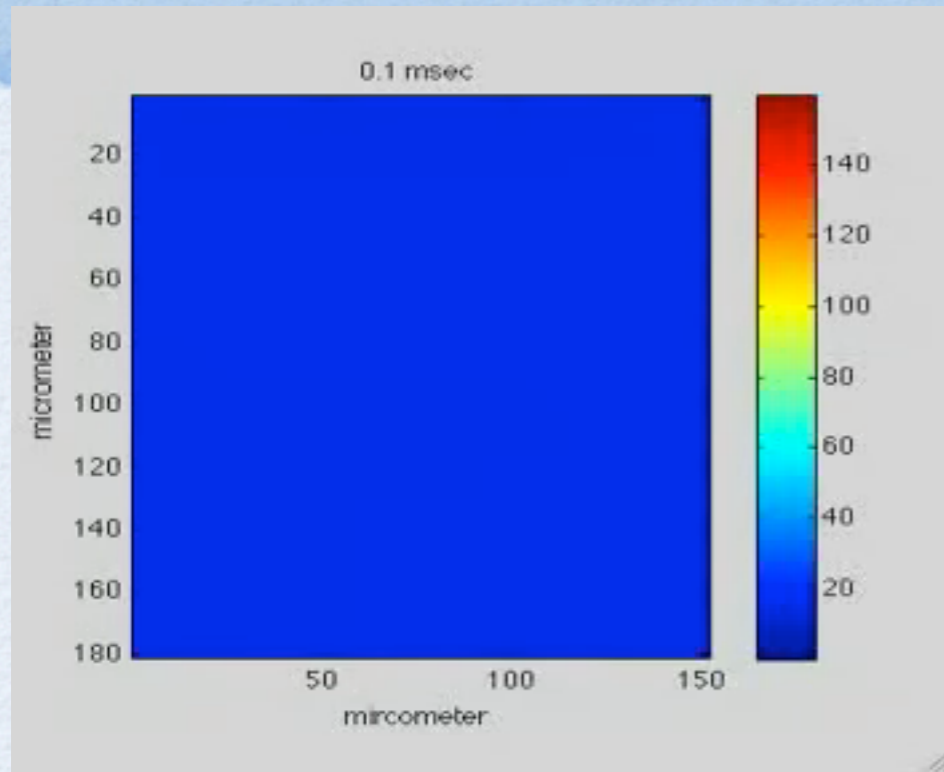


What's in the dish



Michael Berry, Princeton

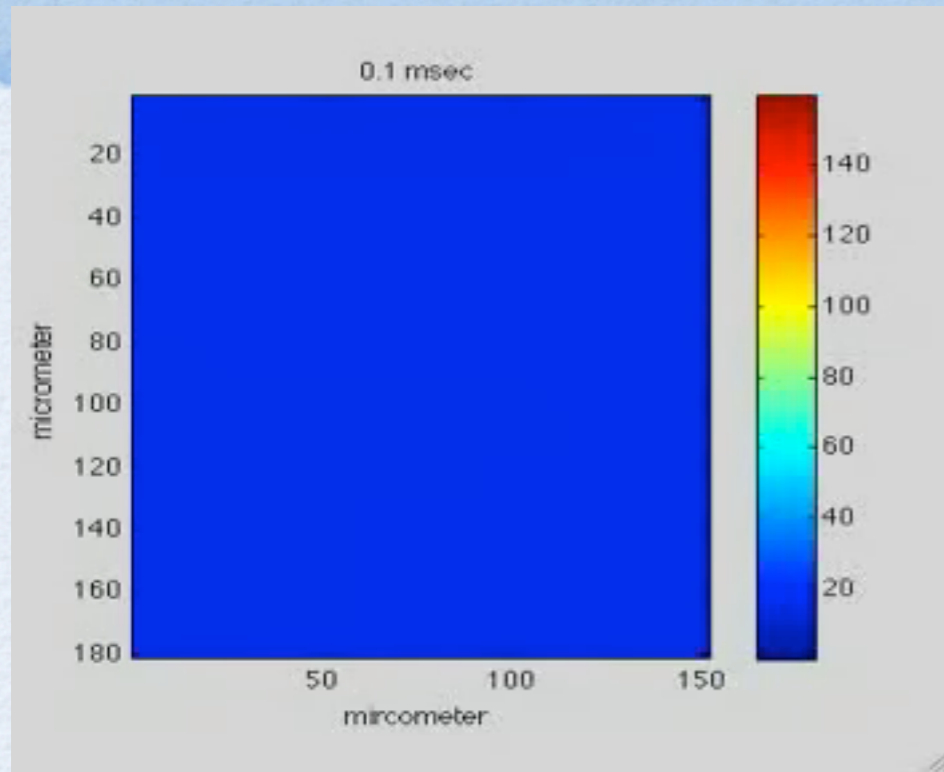
Simple events



67 ms of data,
viewed as a movie.
[data have been smoothed]

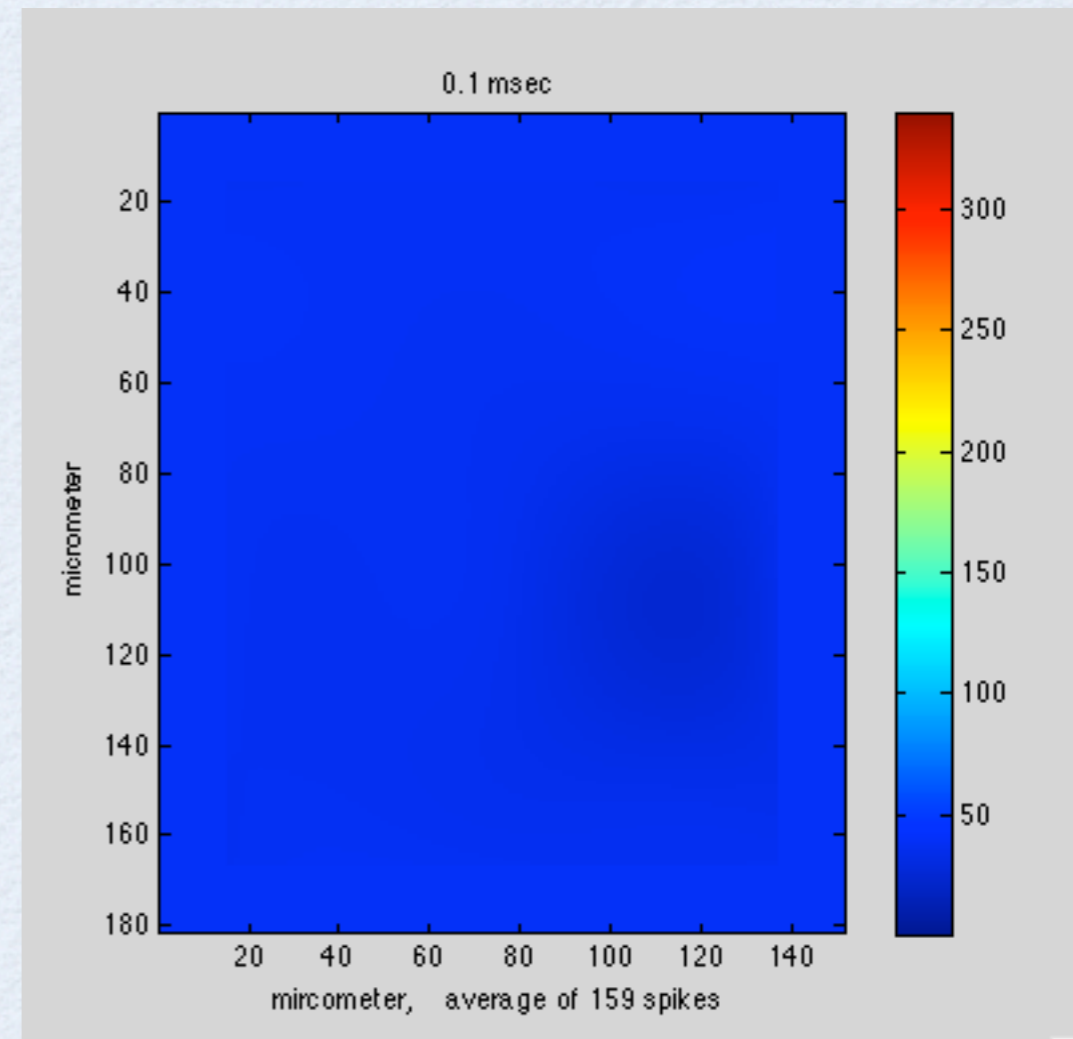
Classic: Gerstein+Clark 1964; Abeles+Goldstein 1977; Schmidt 1984.

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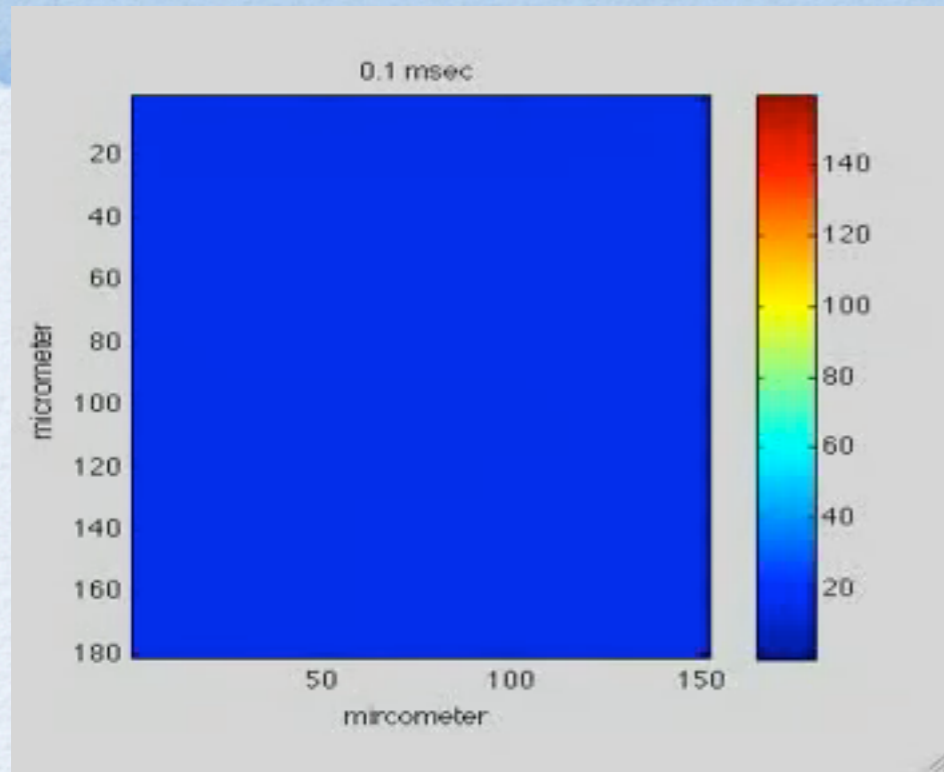
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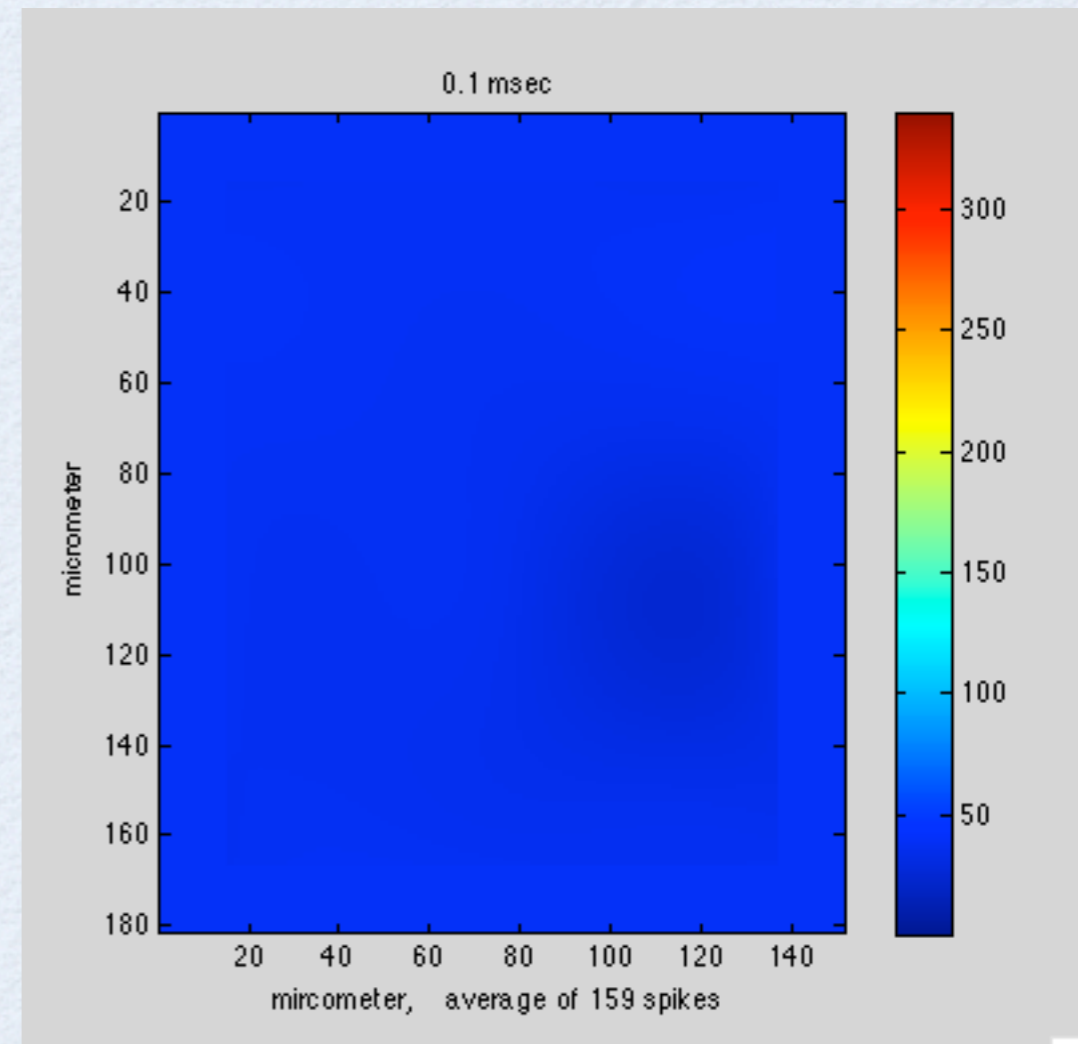
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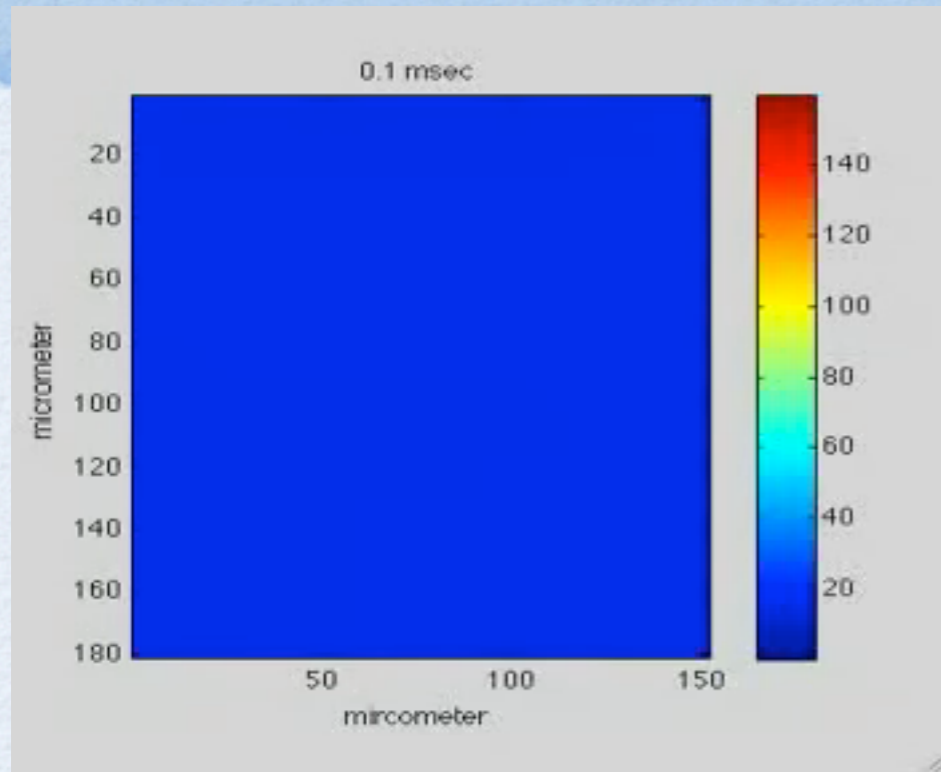
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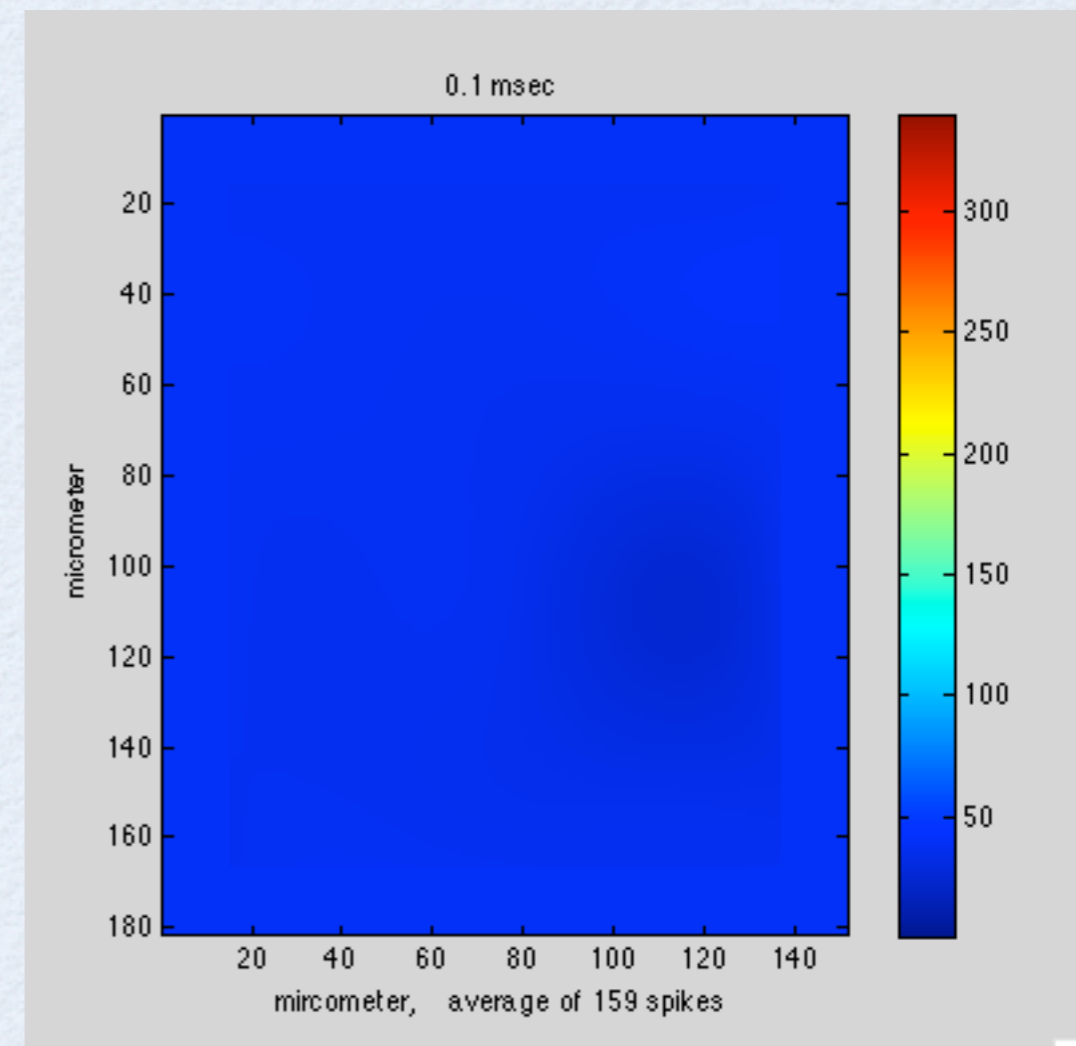
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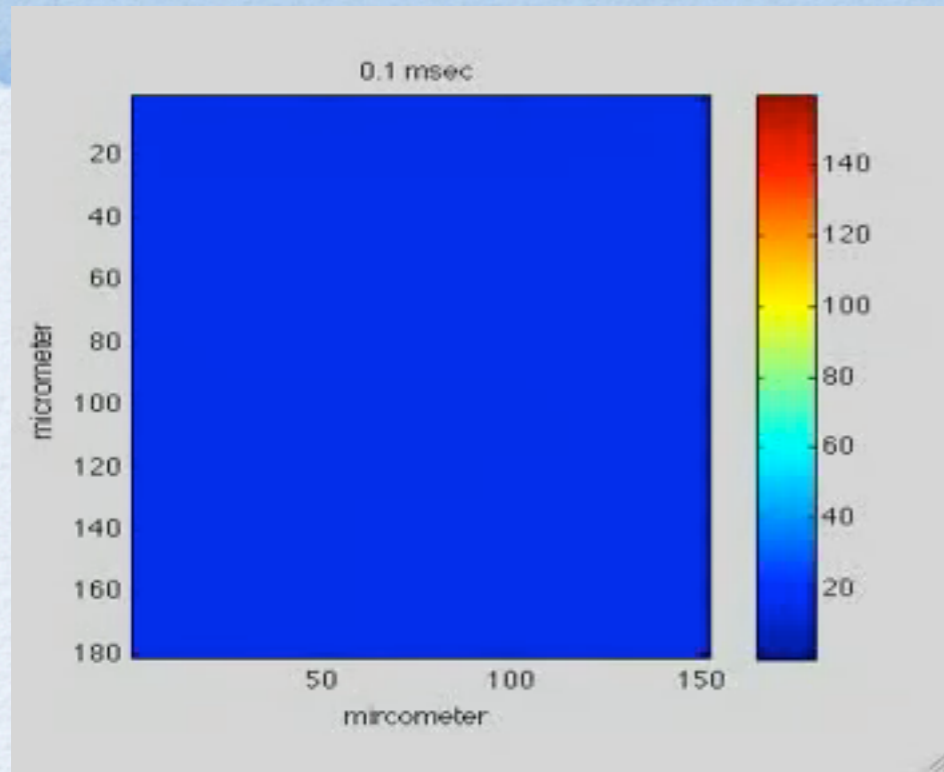
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Mostly we are hearing retinal ganglion cells, as desired, because they're the ones that spike.

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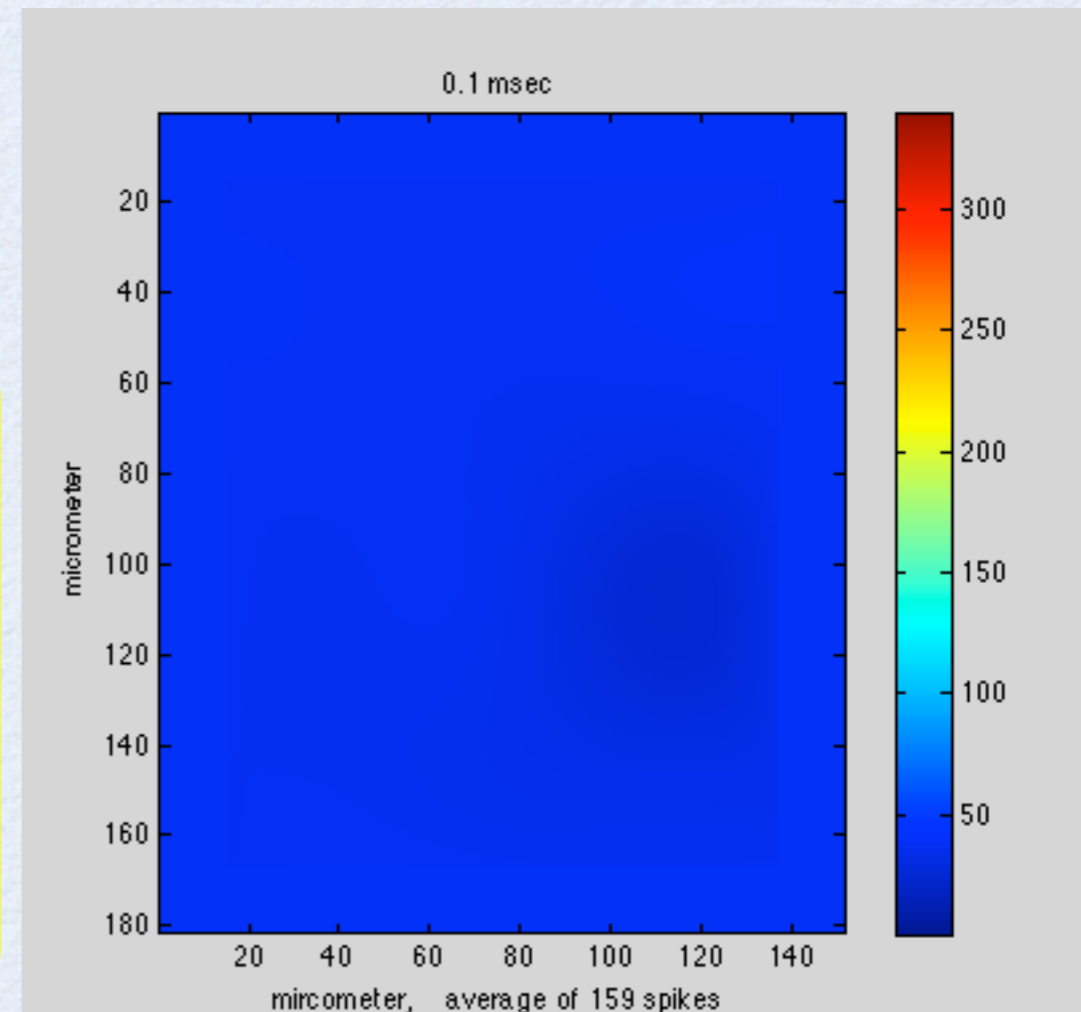


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The **spike-sorting problem** is: Given raw data like these, convert to a list of discrete events (which cells fired at what times).



Classic: Gerstein+Clark 1964; Abeles+Goldstein 1977; Schmidt 1984.

Not-so-simple events

Unfortunately many events are complex, with multiple overlapping spikes in many locations. And of course these may be the **most interesting ones!**

It really matters because “Failure in identification of overlapping spikes from multiple neuron activity causes artificial correlations” [Bar-Gad ‘01]. Moreover, when we graduate to bigger arrays, nearly all events will involve overlaps in time!!

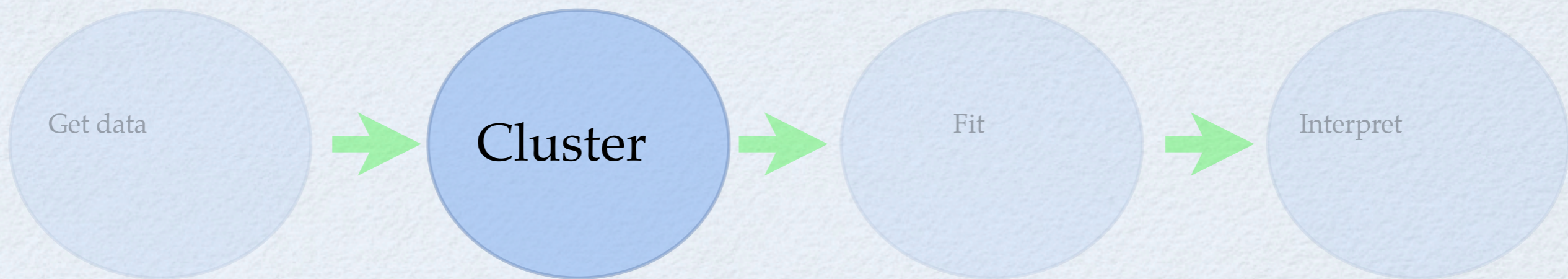
Many authors say **bursts** are a big problem, but here is a nice fit that we obtained with no special effort. See later.

We even handle **overlapping spikes**, which some algorithms do not attempt. See later.

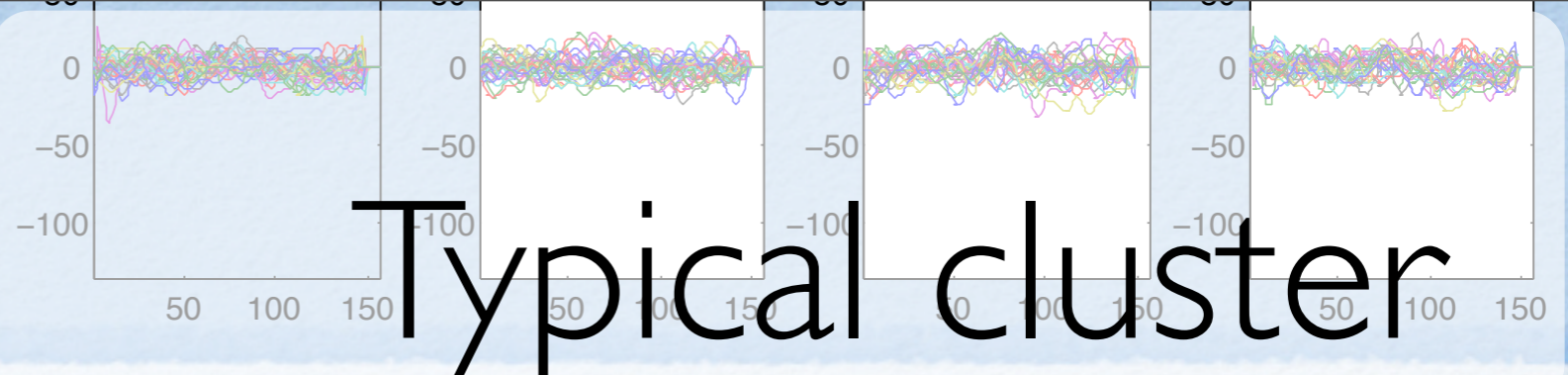


JS Prentice, J Homann, KD Simmons, G Tkacik, V Balasubramanian, PCN, PLoS ONE 6(7): e19884 (2011).

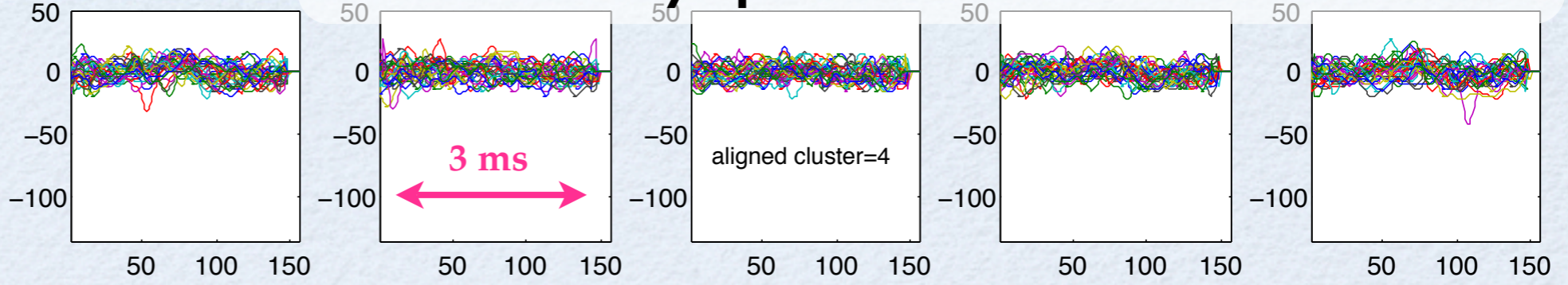
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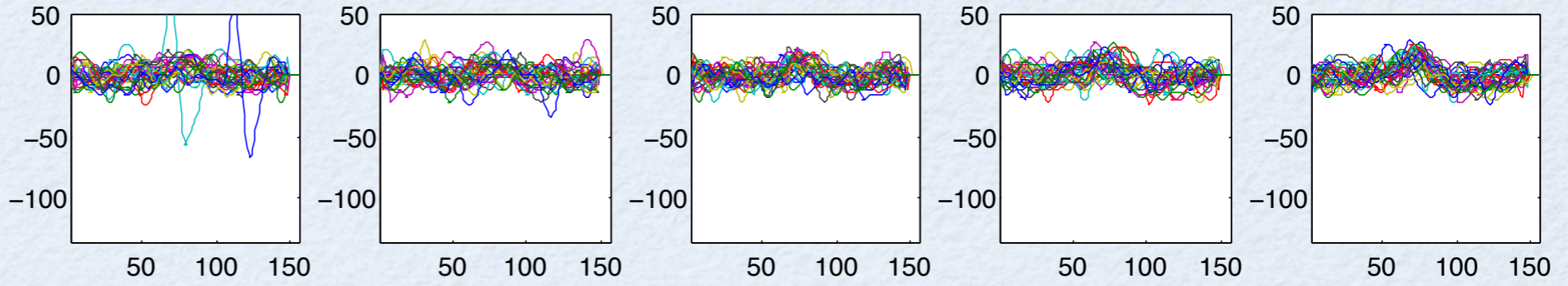
[Sorry, no time to discuss our method for this step.]



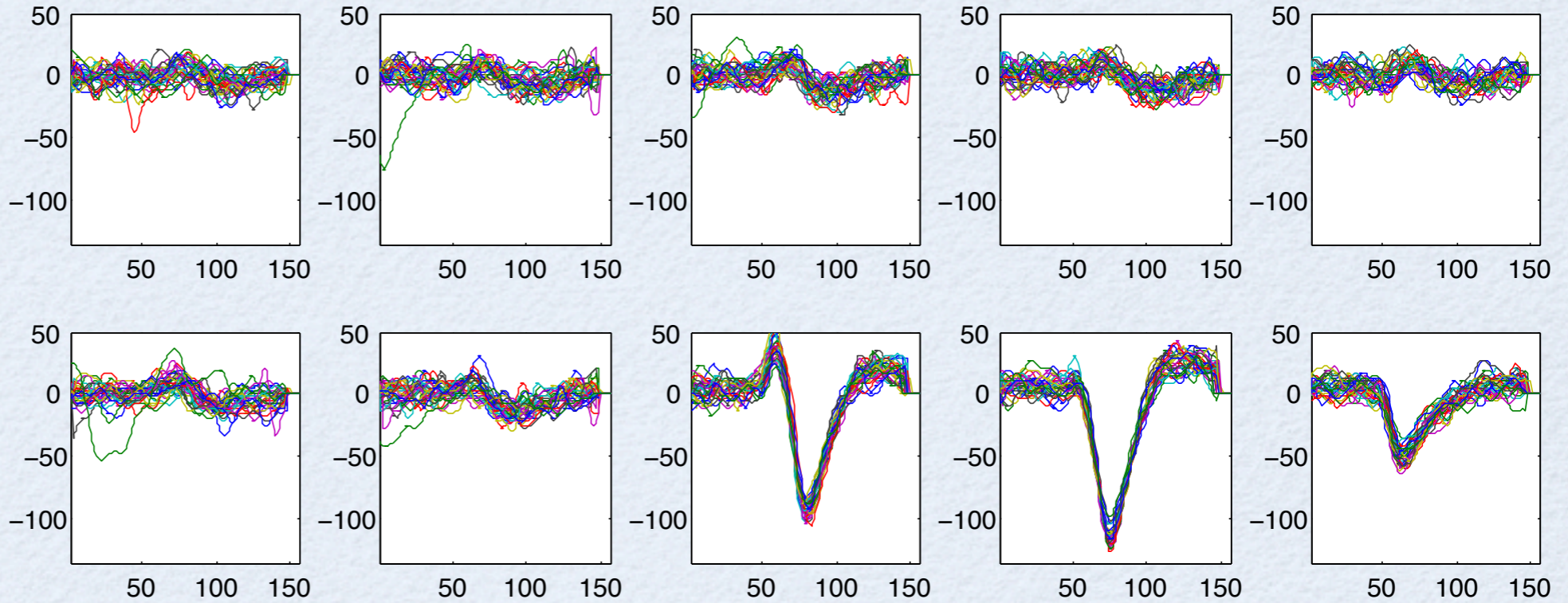
Superposing 50 traces chosen from 284 in this cluster shows that they really do all resemble each other.



Occasional events in which this event collides with another don't affect the "archetype waveform" (template) (next slide).

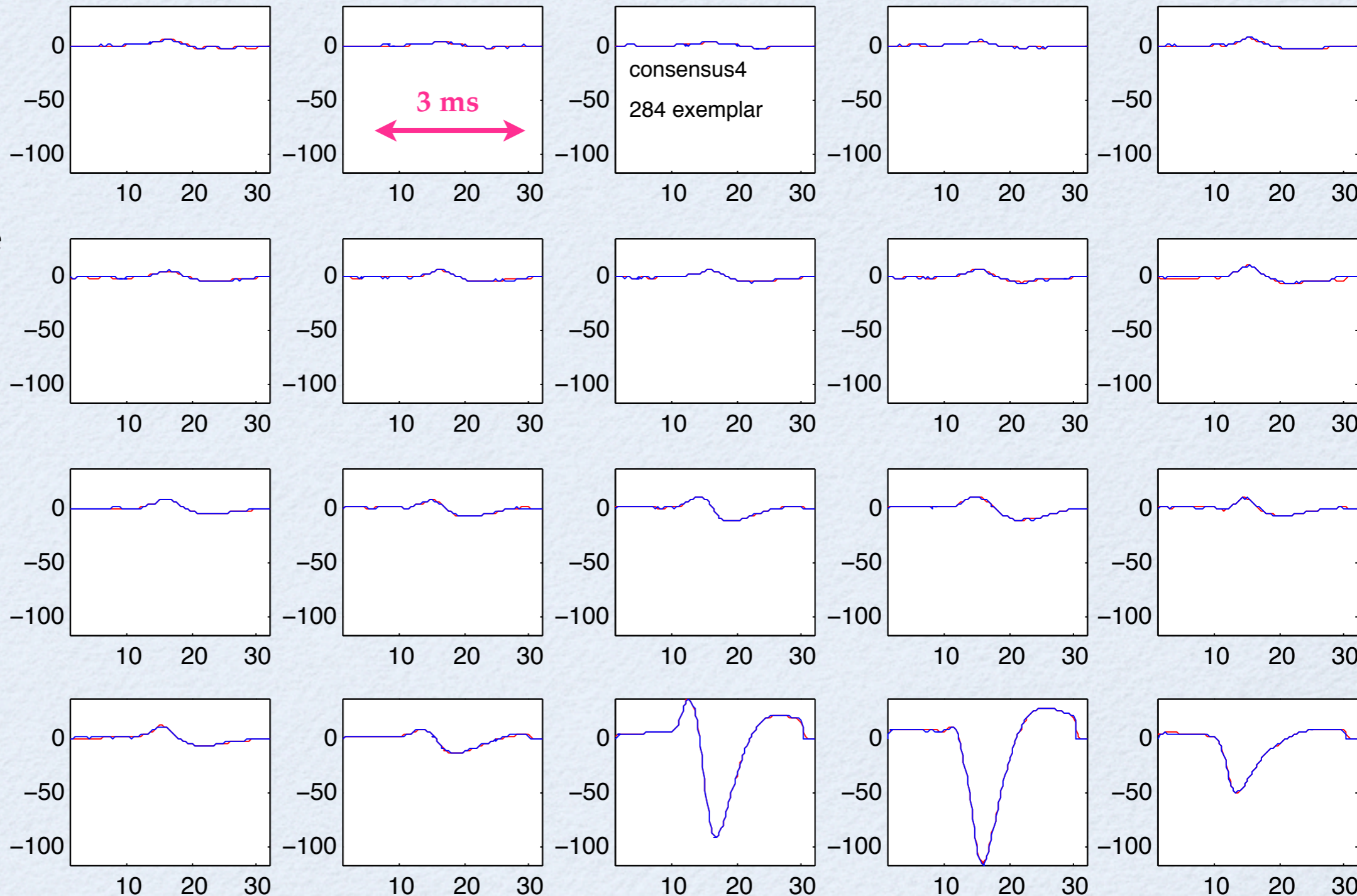
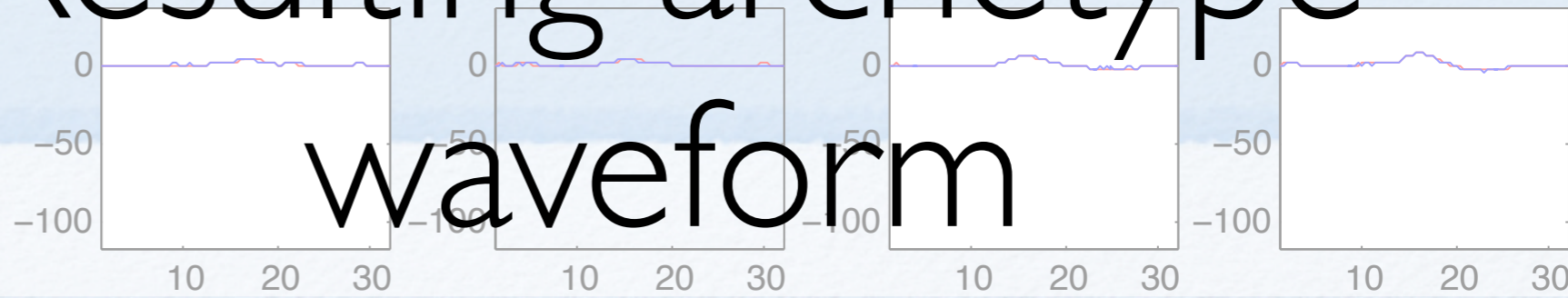


Although the *shape* of each instance of the archetype is quite constant, still its *amplitude* has significant variation.



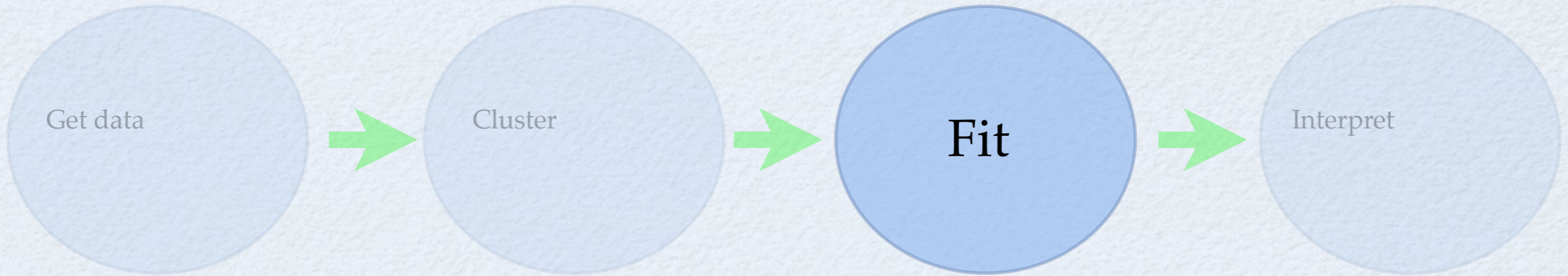
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Resulting archetype waveform



We scaled each instance of each archetype to get best agreement with the others, then took the median at each time point to find our best estimate of the consensus waveform (blue). As a check, the pointwise mean waveform looks the same (red).

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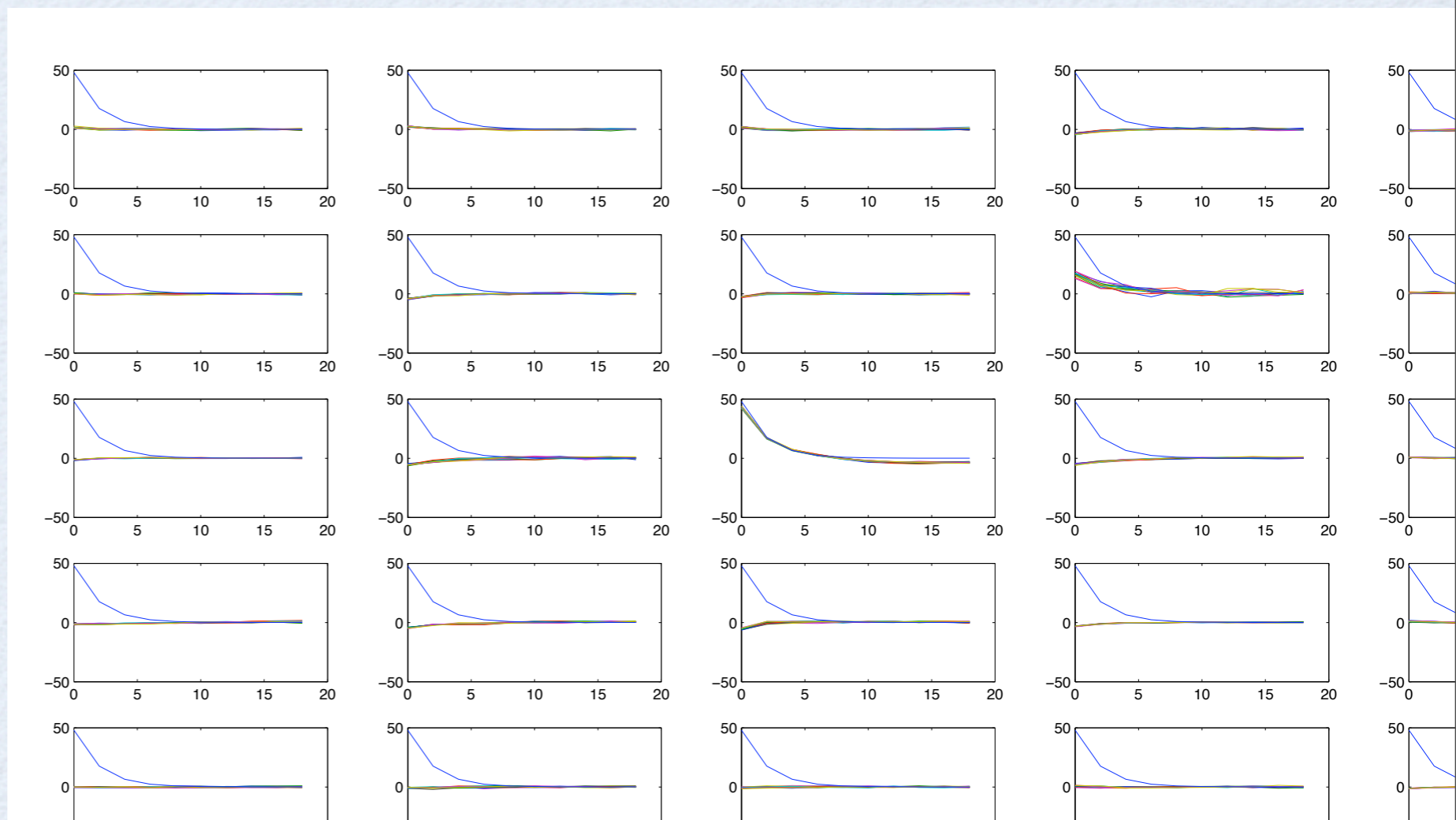
Noise covariance

Vanilla least-squares fitting is not appropriate for time series, because it assumes that every sample is independent of all others--whereas actually, **successive samples are correlated**.

Here is the covariance of channel #13 with all other channels (after an initial spatial filter, also obtained from data). For reference, each channel has a single blue curve showing an exponential function.

We see that #13 is **correlated only with itself**, and it has a **simple covariance matrix** that is easy to invert. The inverse covariance thus obtained defines our correlated Gaussian model of the noise.

[Again: The covariance is **not** a delta function, contrary to what is assumed in naive least-squares fitting.]



On inference

Suppose we measure some experimental data, and wish to make an inference about some situation that we cannot directly observe. That is, we imagine a variety of worlds with different values of X , and ask which is most probable given the observed data.

See M. Denny and S. Gaines, *Chance in Biology*.

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If we know the probability that those data would have arisen in a world with a particular value of X , then Bayes's formula gives us what we actually want:

$$P(X|\text{observed data}) = P(\text{data}|X) \frac{P(X)}{P(\text{data})}$$

We can ignore the denominator, if all we want is to compare two hypotheses (e.g. maximize over X).

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For our application, we'd like $\mathbf{P}(\text{spikes} | \text{data})$, where "data" is an observed waveform and "spikes" refers to a collection of spike archetypes μ_1, \dots occurring at times t_1, \dots with amplitudes A_1, \dots relative to the amplitude of the corresponding archetype (neuron). Bayes's formula gives what we want as

$$\mathbf{K} \times (\text{likelihood}) \times (\text{prior}) = \mathbf{KP}(\text{data} | \text{spikes})\mathbf{P}(\text{spikes})$$

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Bayesian idea

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To get the **prior**, $\mathbf{P}(\text{spikes})$, assume that for a single spike it has the form

$$\underline{P^{\text{cell}}(\mu)} \underline{P^{\text{time}}(t)} \underline{P^{\text{ampl}}(A|\mu)}$$

The three factors are respectively the **popularity of this neuron**, **uniform in time**, and a Gaussian reflecting its **typical amplitude and amplitude variability**. We get these priors from the data subset used in clustering.

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To get the **likelihood function** $\mathbf{P}(\text{data} \mid \text{spikes})$, suppose that the data consist of one archetype, plus noise. And suppose that the noise is some **Gaussian**, independent of which spikes fired. We know all about this Gaussian from our measurement of noise covariance.

Then the likelihood is that distribution, evaluated at the difference between the actual waveform and the idealized one. [Pouzat et. al. 2002]

Priors con
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Bayesian idea, II

We start with an experimental trace (“data”).

We find its peak (absolute minimum), and start looking for a spike there.

We ask for the **likelihood ratio** between the hypotheses of no spike versus one spike of given type, at given time, with given amplitude.

- ★ To compute the likelihood of no spike, evaluate the noise distribution on the trace.
- ★ To compute the probability of one spike, choose a spike archetype and a value of t , the spike time. Holding the “data” fixed, **the probability is now a Gaussian function in the remaining parameter A** , so it’s fast and easy to marginalize over A .

[Nuts and Bolts]

Let $V_\alpha(t)$ be measured voltage, electrode α and $F_{\mu\alpha}(t)$ be archetype waveform of type μ . Define the deviation $[\delta\mathbf{V}]_{\alpha t} = V_\alpha(t) - AF_{\mu\alpha}(t - t_1)$

Then the probability that one spike, of type μ , is present is

$$P(\text{spikes} \mid \text{data}) = K_\mu \exp \left[-\frac{(A - \gamma_\mu)^2}{2\sigma_\mu^2} - \frac{1}{2} (\delta\mathbf{V})^t \mathbf{C}^{-1} (\delta\mathbf{V}) \right]$$

The noise covariance

which is a Gaussian in A . So it's easy to marginalize over A : **just complete the square!**
[Here $K_\mu = P^{\text{cell}}(\mu)P^{\text{time}}(t_1) (2\pi\sigma_\mu^2)^{-1/2}$ doesn't depend on A .]

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If the winner's likelihood ratio is good enough (bigger than about 1), we say there's a spike here. **That's the absolute criterion I promised earlier.**

Test our assumptions

Can we really assume that the spikes from a particular cell differ only in overall amplitude? We took many events that contained a single spike of each type. Point by point in time, we subtracted the scaled shifted archetype and found the residual (on each channel).

Green: the archetype itself.

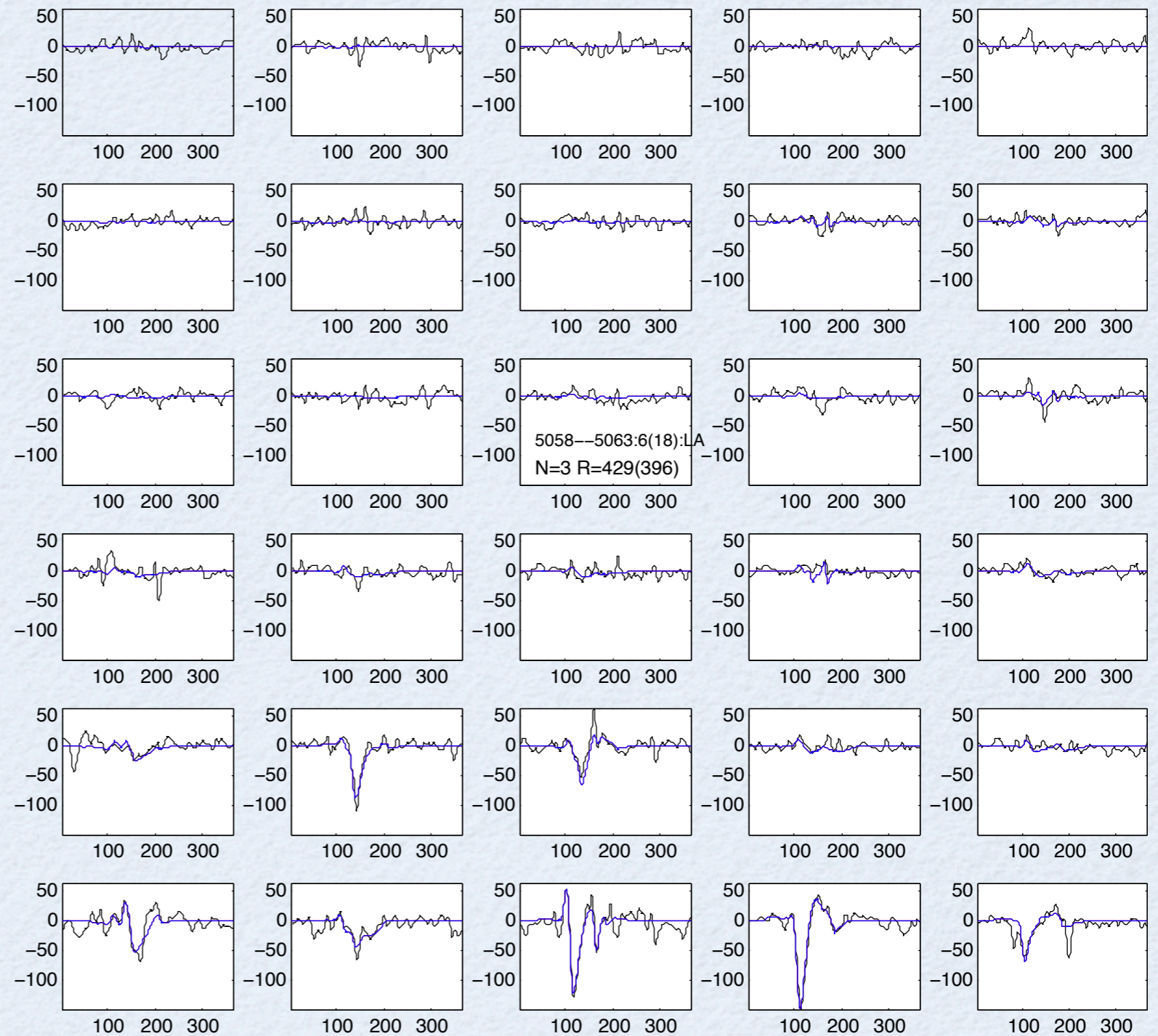
Red: mean deviation from archetype.

Blue: std deviation from archetype.

We really do subtract spikes pretty completely.



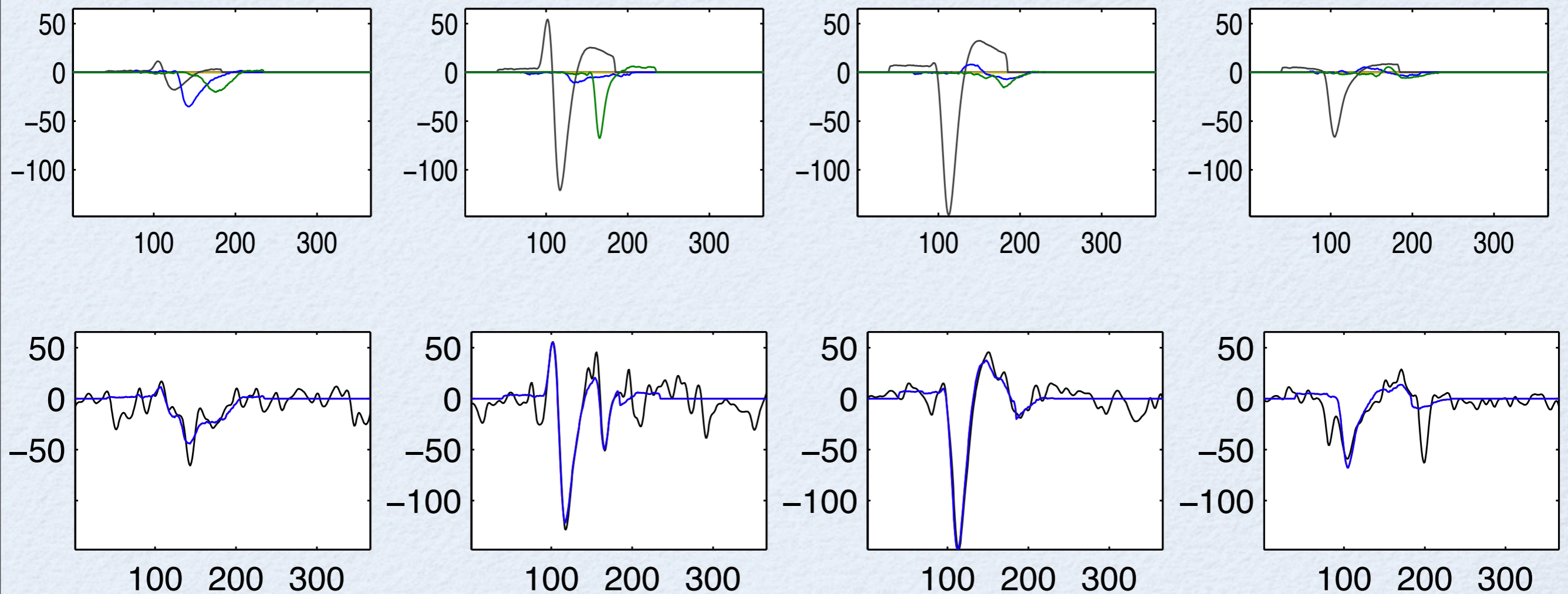
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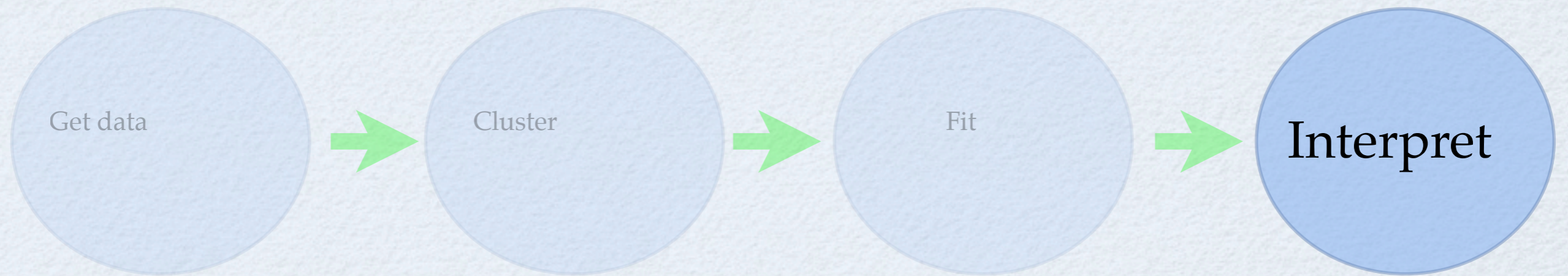
Successfully fit overlaps

Top: Closeup of four channels, showing three fit archetypes found by the algorithm.
Bottom: sum of those fits (color) versus actual data (black).



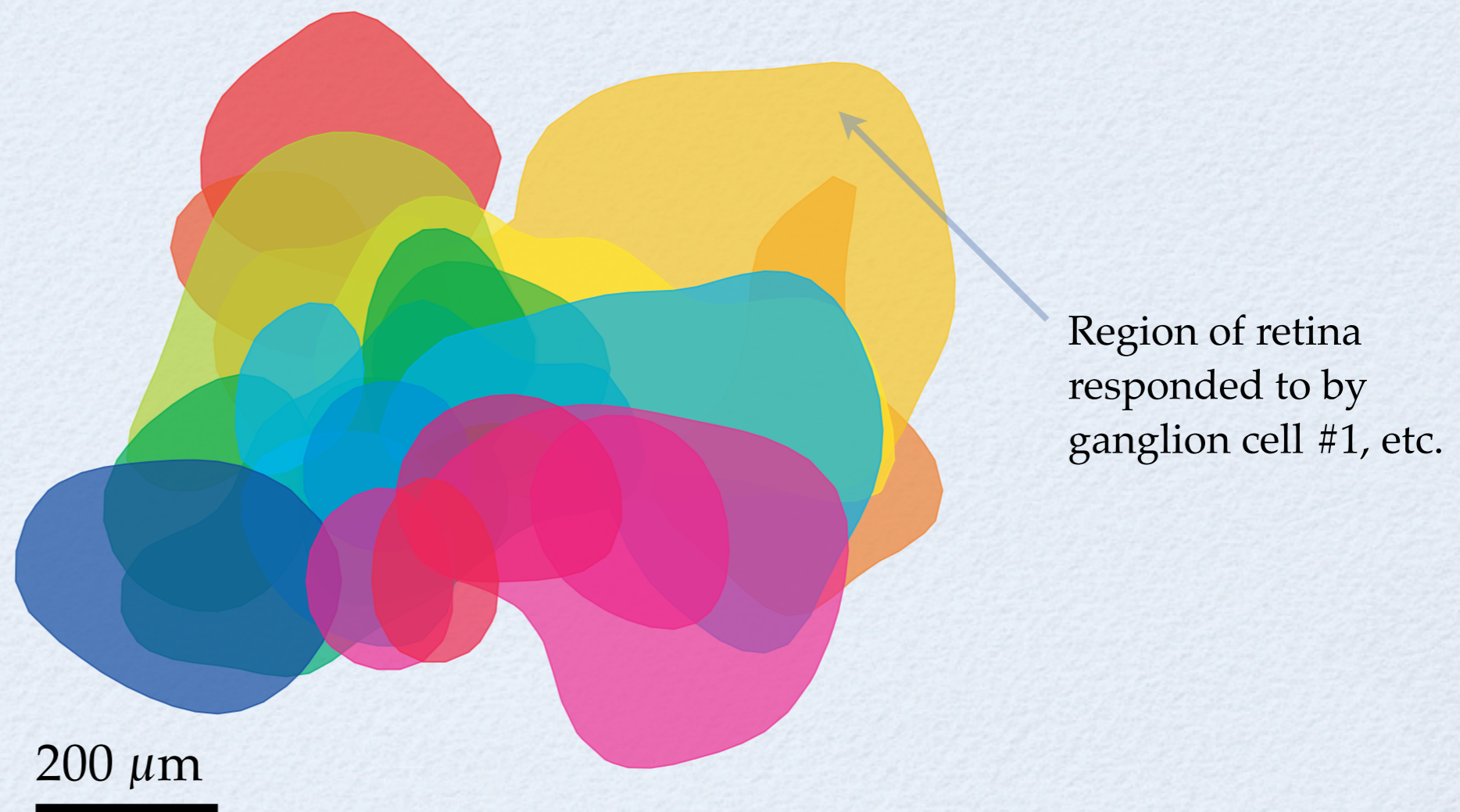
Successfully fit bursts





Each cell has a receptive field...

... and they tile the whole visual field. MEA recording is **high throughput**: We got dozens of cells all at once. Here are cells from just one functional group, “on cells.” Each putative receptive field is a single connected region of image space.



KD Simmons, JS Prentice, G Tkacik, J Homann, PCN, V Balasubramanian, submitted.

Receptive fields

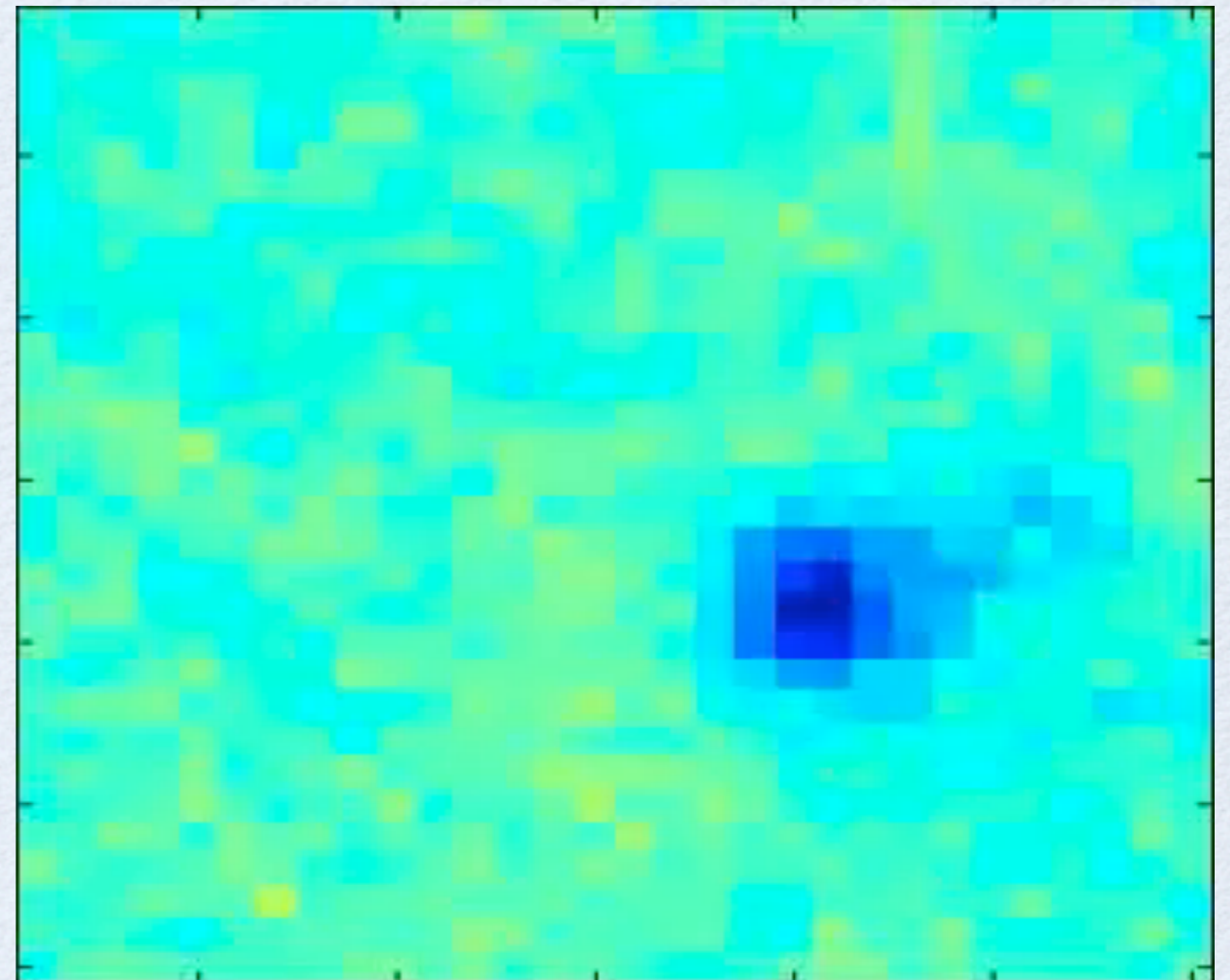
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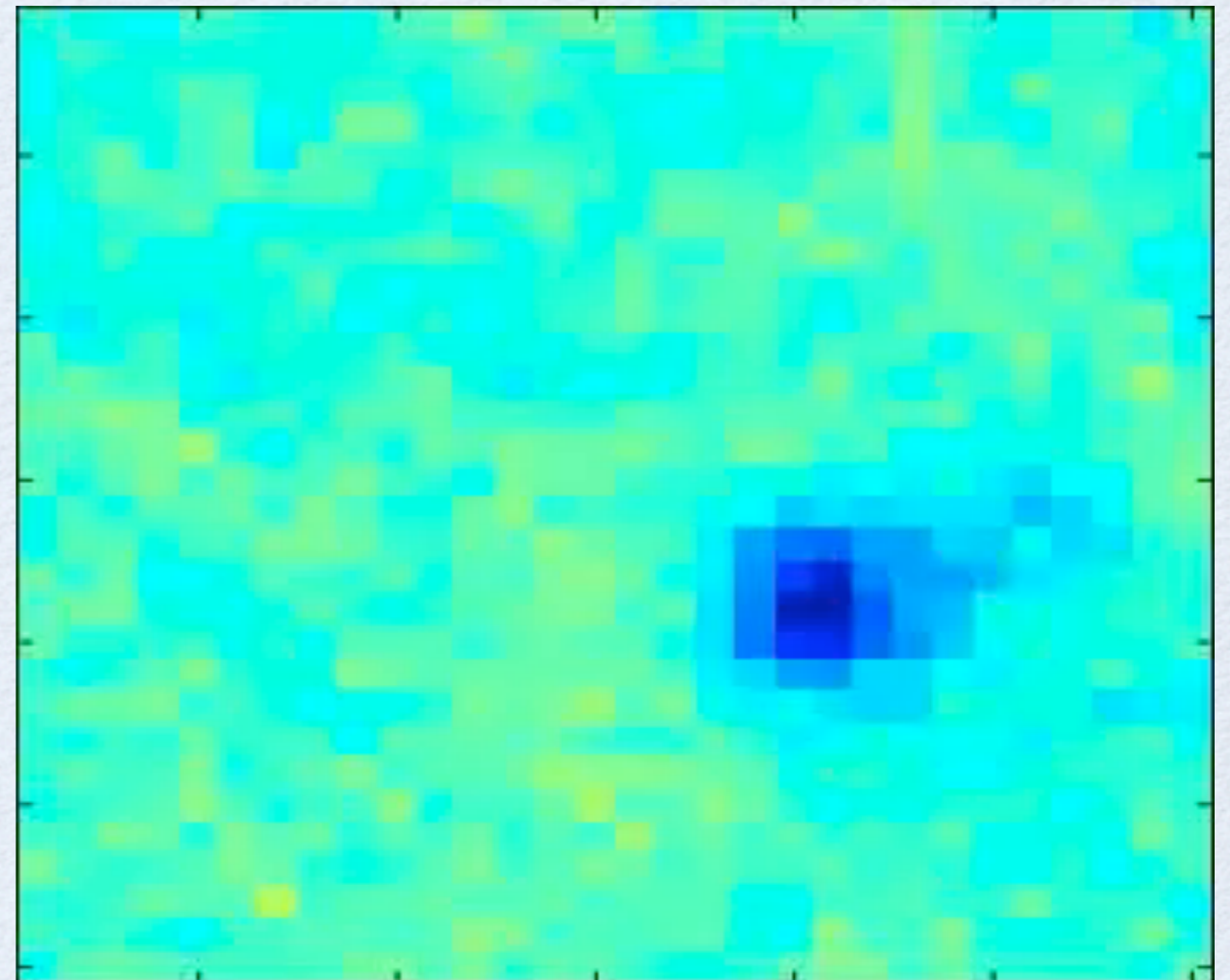
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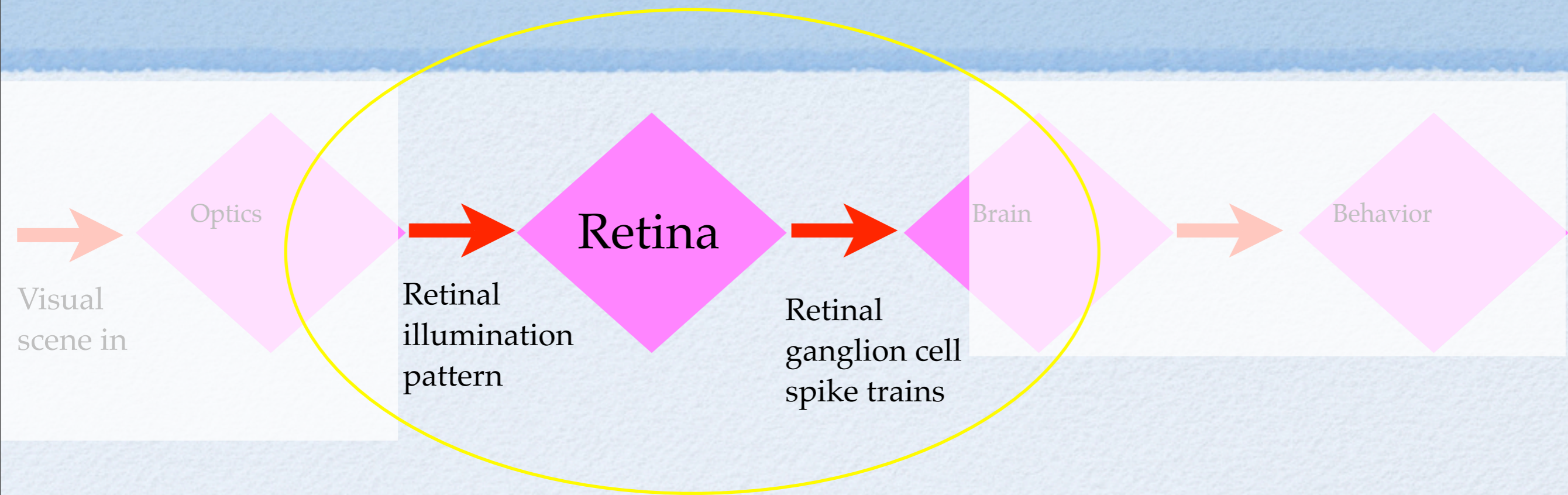
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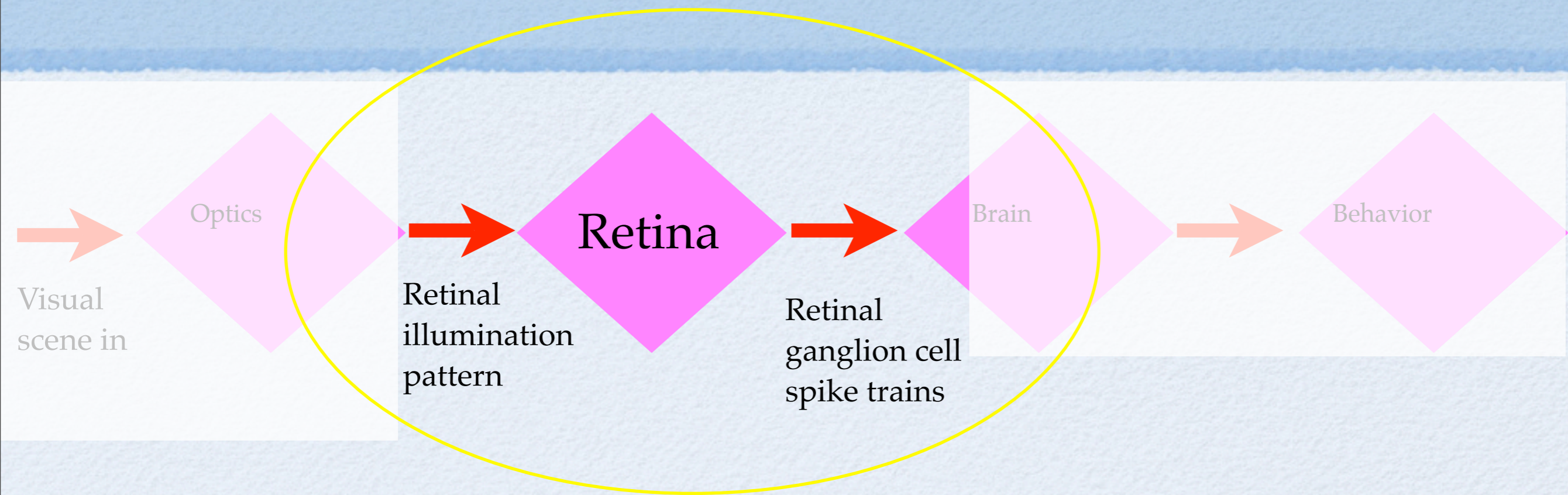


Takehome Part III

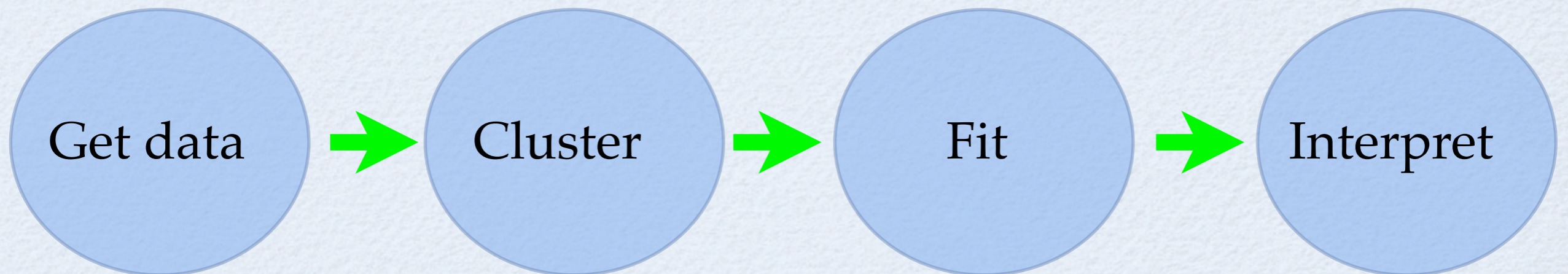


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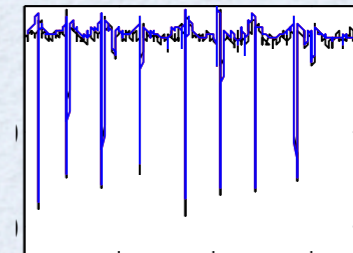
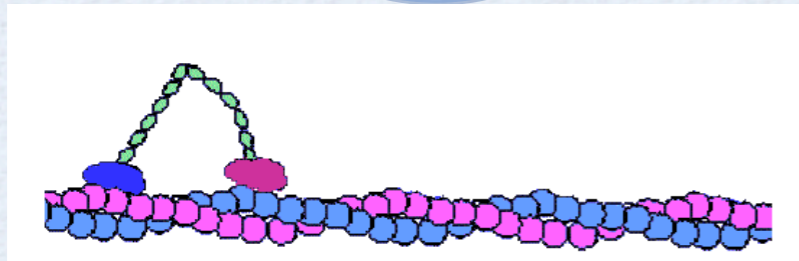
Full circle

OK, I was a scatterbrain and gave you three talks. But wait -- if I can fill in the spaces

Medical tests

Changepoint Analysis

Multi-Electrode Array



Full circle

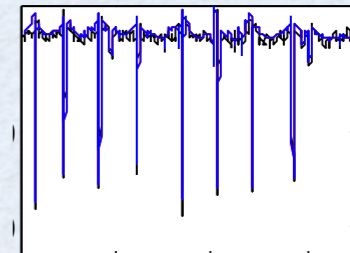
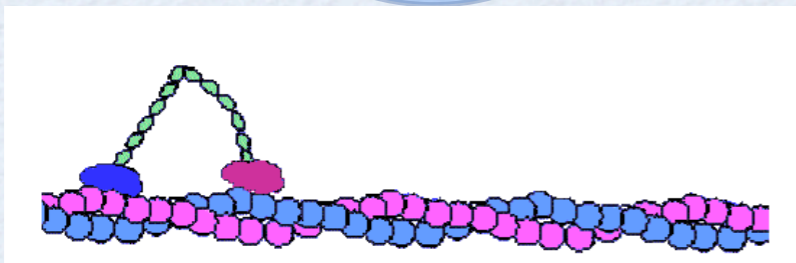
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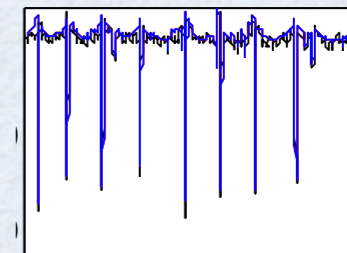
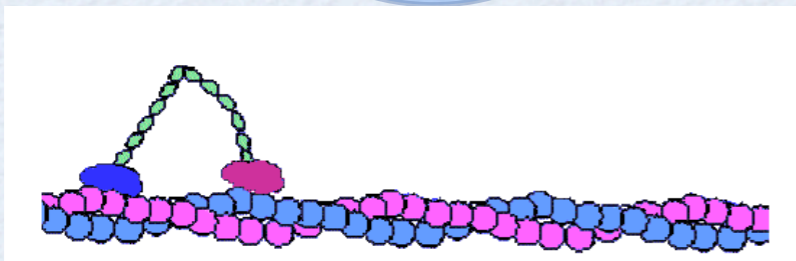
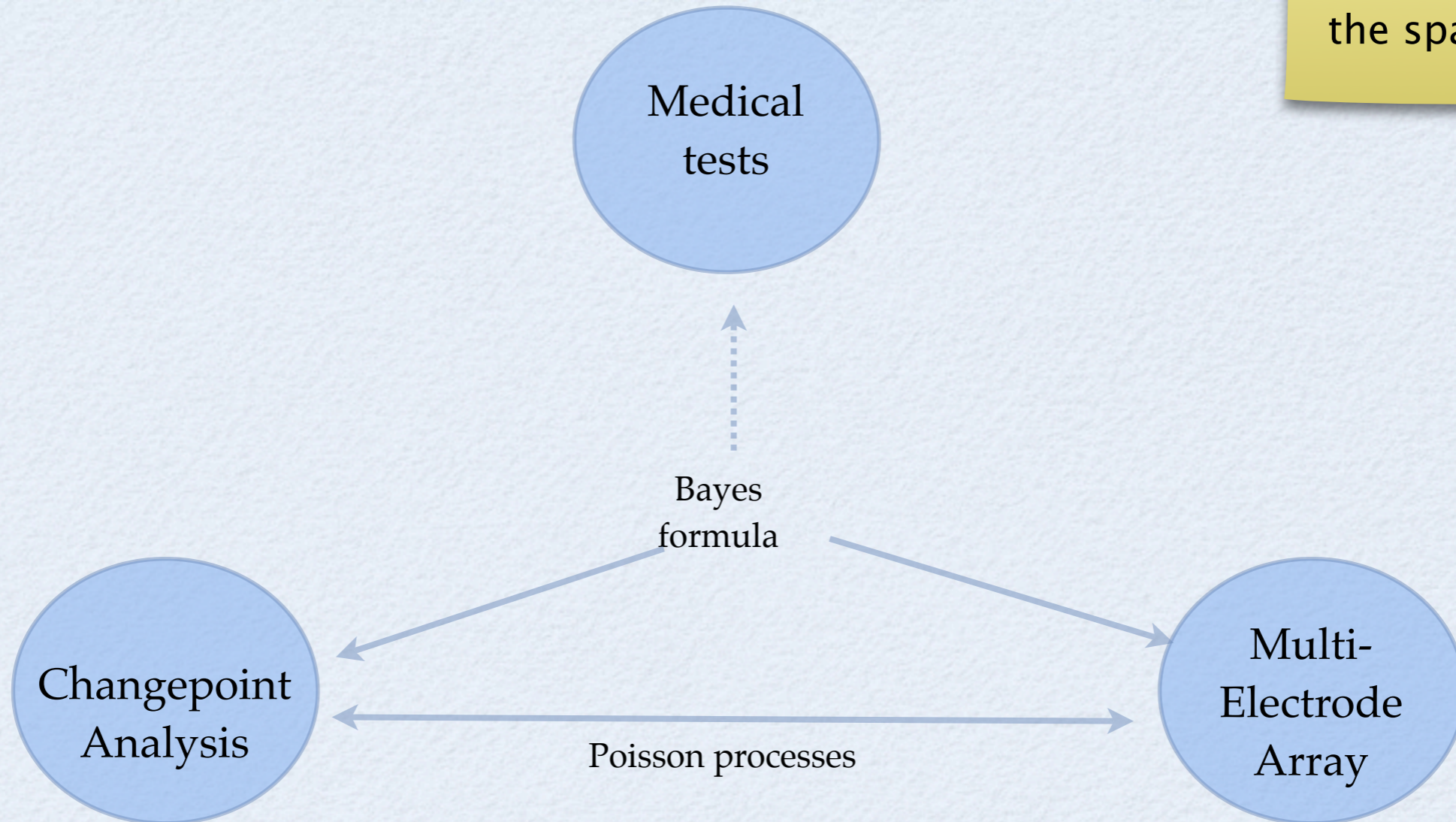
Poisson processes

Multi-Electrode Array



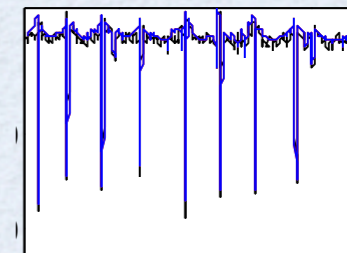
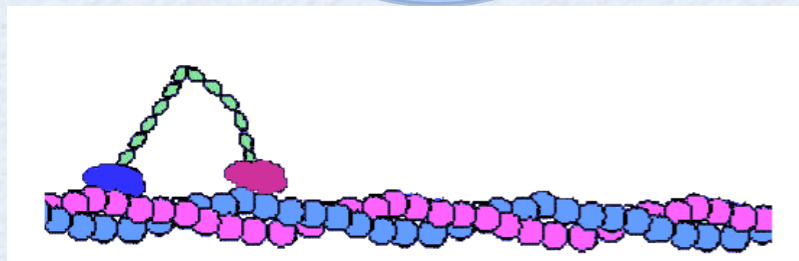
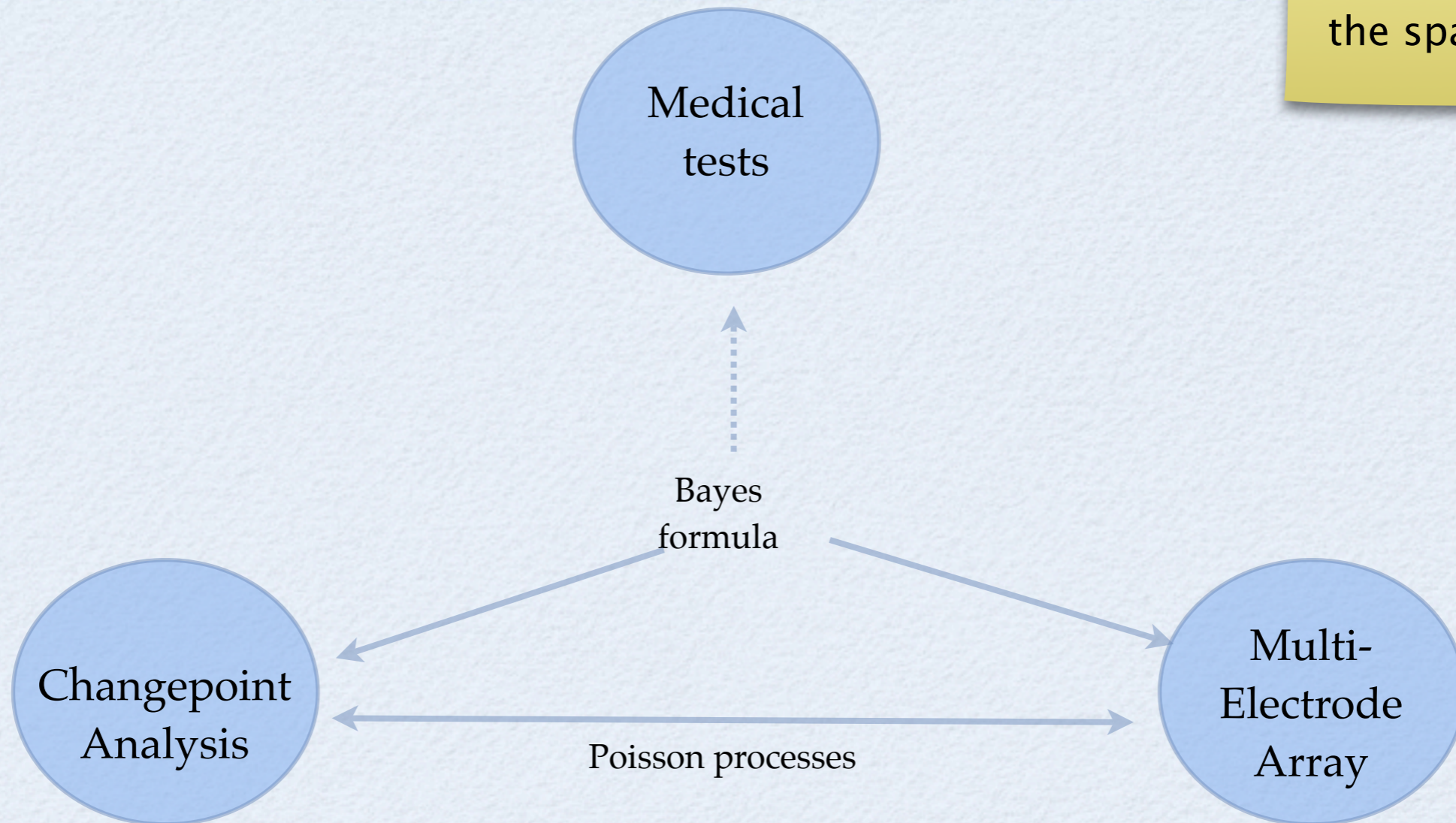
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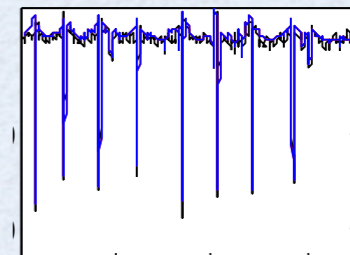
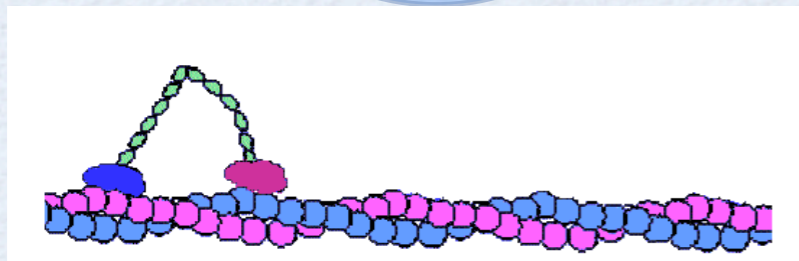
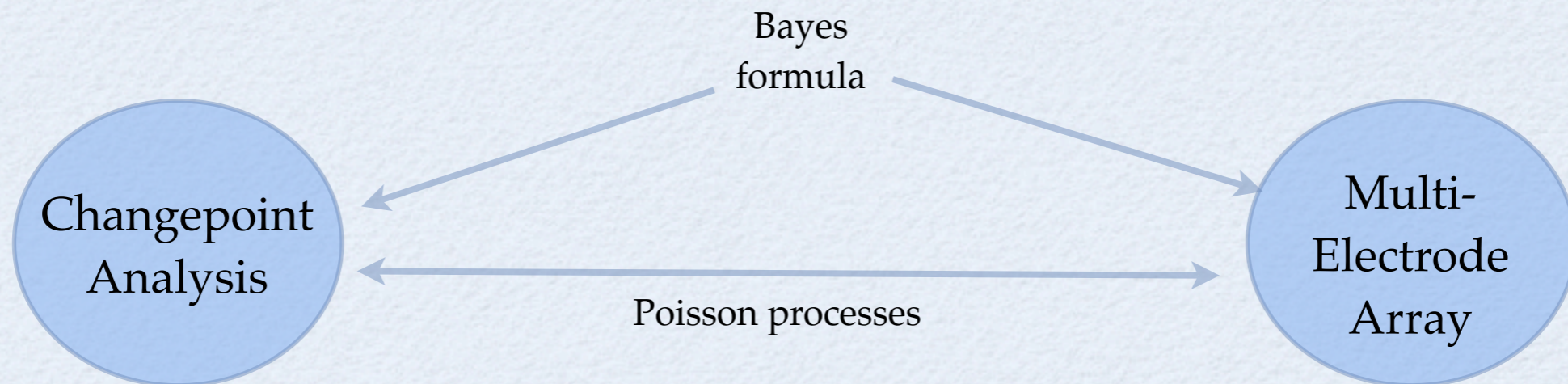
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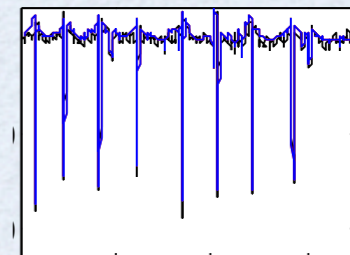
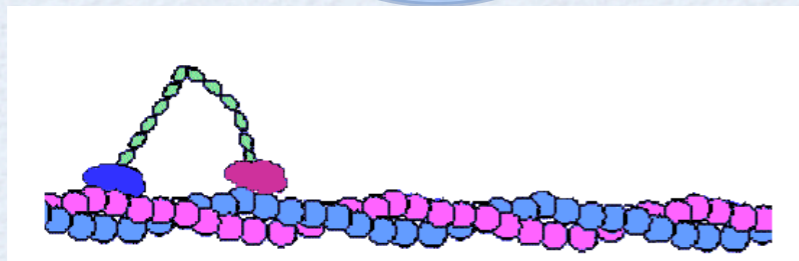
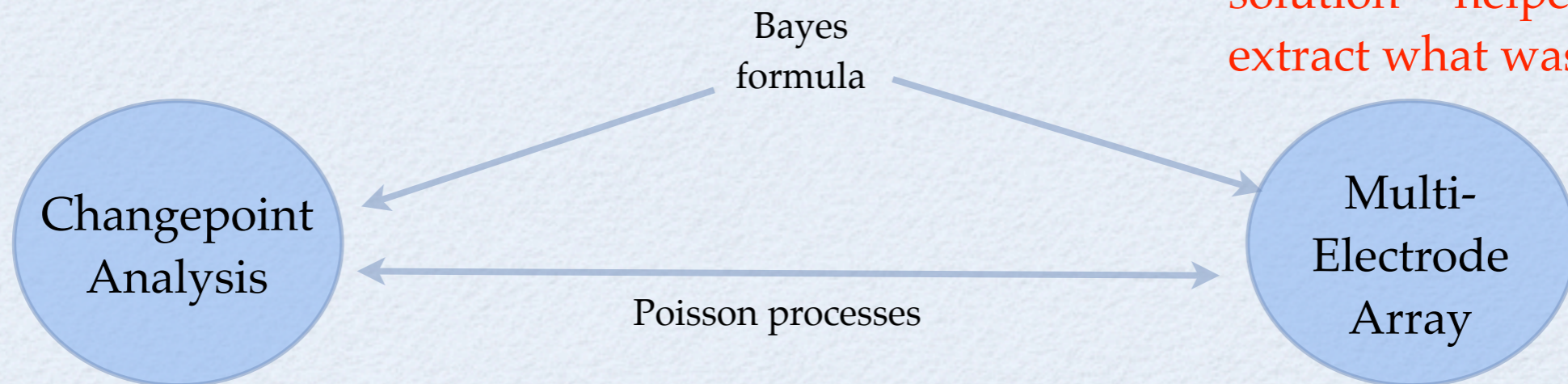
Theory can cut across apparently different kinds of experiment, offering useful methods to one domain from another without having to reinvent everything.

Wait, there's more



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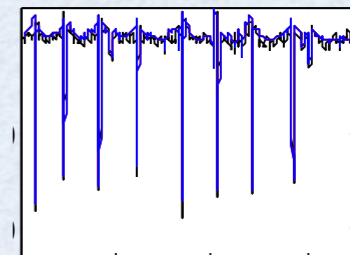
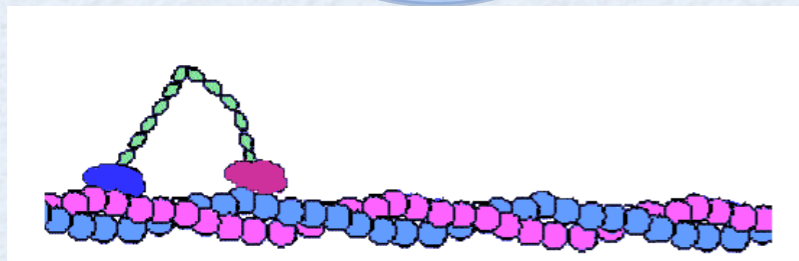
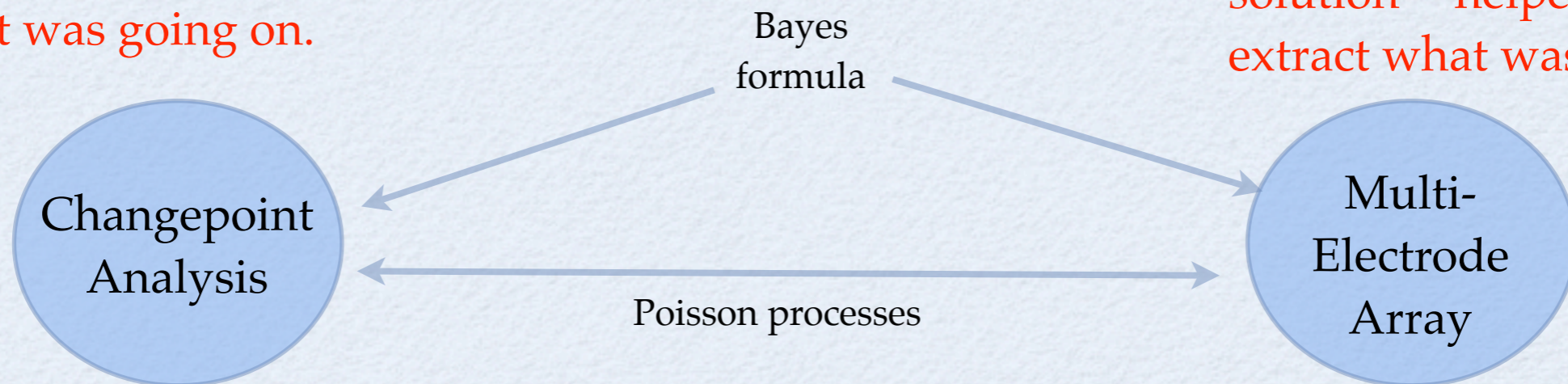
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Wait, there's more

A physical model -- photon theory -- helped us to extract what was going on.

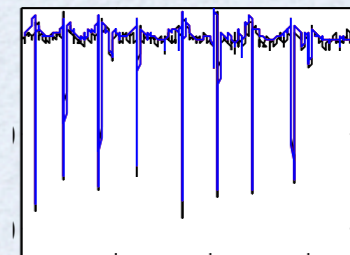
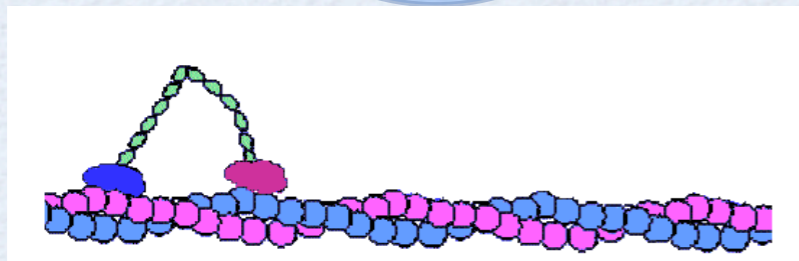
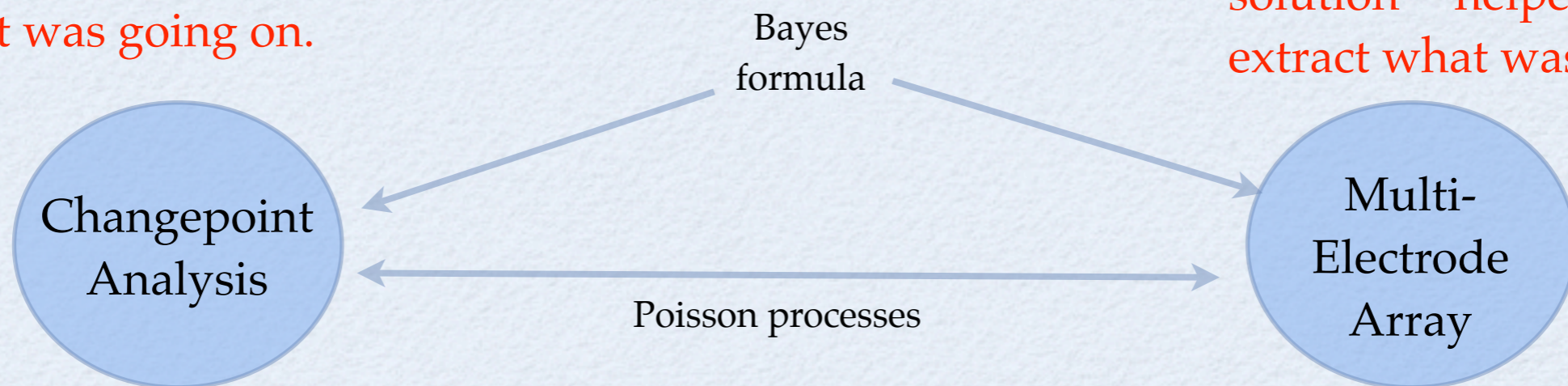
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There is something weirdly -- *unreasonably* -- effective about approaching biological systems with a **physical model**. I don't understand why. I don't *need* to understand why.

Go long

along the lines of, "clean up the bibliography"

Go long

Often, when we want to justify theory, we scratch our heads and say, “Well Hodgkin and Huxley was a big deal.”

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Indeed. But that sort of cherry-picking approach can leave the impression that theory is something that happens every 50 years or so. It's also too reverent.

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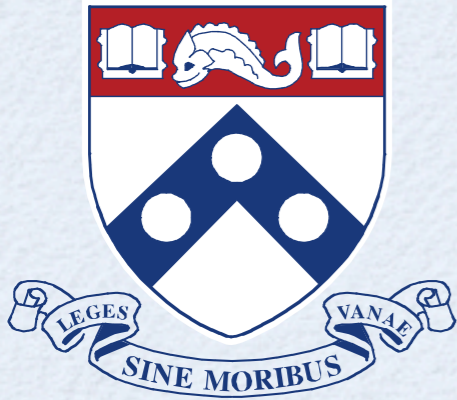
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We like to teach famous success stories in science, but we don’t always remember to present them as showcases of the utility of physical modeling.

Thanks

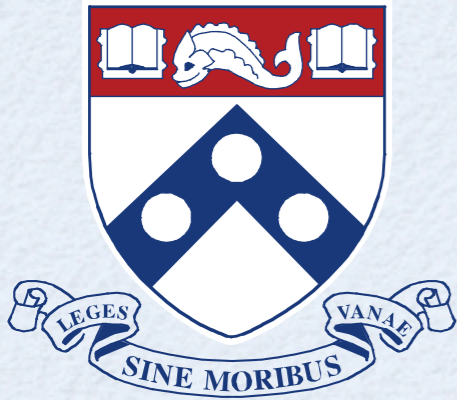


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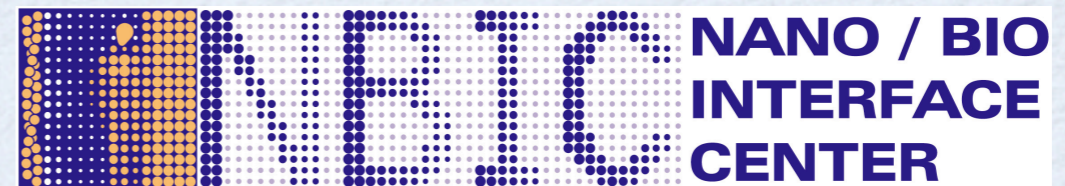
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NSF DMR,
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Training grant



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